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AFATL-TR-73-32

AN ANALYSIS OF METHODS FOR
EXTRACTING AERODYNAMIC COEFFICIENTS
FROM TEST DATA

UNIVERSITY OF FLORIDA

TECHNICAL REPORT AFATL-TR-73-32
FEBRUARY 1973

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AD 913508

An Analysis Of Methods For Extracting Aerodynamic Coefficients From Test Data

Donald C. Daniel

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FOREWORD

This analysis was conducted by the University of Florida, Gainesville, Florida, under Contract F08635-71-C-0080 with the Air Force Armament Laboratory, Eglin Air Force Base, Florida. The program monitor for the Armament Laboratory was Dr. George B. Findley (DLGC). This effort was conducted during the period January 1971 to February 1973 and was partially supported by the Air Force Office of Scientific Research (AFOSR) under its project 9871.

The principal investigators for the University of Florida were Drs. T. E. Bullock and M. H. Clarkson.

This report is based on a dissertation previously submitted by the author to the Graduate School of the University of Florida in partial fulfillment of the requirements for the degree Doctor of Philosophy.

This technical report has been reviewed and is approved.


RANDALL L. FETTY, Colonel, USAF
Chief, Product Assurance Division

ABSTRACT

An analysis of numerical methods for extracting aerodynamic coefficients from dynamic test data has been conducted. The emphasis of the analysis is on the effects that random measurement errors in the data and random disturbances in the system have on the accuracy with which the coefficients for linear and nonlinear systems can be determined. Both deterministic and stochastic methods for extracting the coefficients and determining their uncertainties are considered.

The deterministic technique considered, due to Chapman and Kirk, provides excellent estimates of both linear and nonlinear static pitching moment coefficients for the range of measurement errors and system noise considered. Somewhat less accurate estimates of linear damping coefficients are obtained. Nonlinear pitch damping coefficients extracted using this deterministic technique are affected considerably by both measurement errors and system noise. The estimated standard deviations of the extracted coefficients obtained using standard techniques are generally adequate when the data being analyzed contain only measurement errors.

The stochastic approach considered demonstrates the feasibility of using an extended Kalman filter, with a parameter augmented state vector, for determining the values of the aerodynamic coefficients and their uncertainties from noisy dynamic test data.

The specific filter used generally reaches near steady-state conditions in its estimates of the parameters in less than one second. Variations in the initial parameter variances or in the estimates of the noise statistics essentially affect only the determination of the nonlinear damping parameter.

Parameter estimates obtained from the extended filter compare favorably with previously obtained results using deterministic techniques. Estimates of the parameter uncertainties provided by the filter are generally superior to those obtained with deterministic techniques particularly when system noise has corrupted the data.

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NOMENCLATURE

<u>Symbol</u>	<u>Definition</u>
A	Reference area
d	Reference length
I	Vehicle moment of inertia about an axis through the center of gravity and normal to its pitch plane
q	Dynamic pressure, $1/2 \rho v^2$
V	Freestream velocity
$C_{m\alpha 0}$	Static pitching moment coefficient derivative, rad^{-1}
$C_{m\alpha 2}$	Static pitching moment coefficient derivative, rad^{-3}
$C_{m\alpha 4}$	Static pitching moment coefficient derivative, rad^{-5}
$C_{mq 0}$	Pitch damping coefficient
$C_{mq 2}$	Pitch damping coefficient, rad^{-2}
$C_{m\delta\alpha}$	Static pitching moment coefficient at $\alpha=0$
α	Pitch angle
ρ	Freestream density
B^{-1}	The inverse of the matrix B
B^T	The transpose of the matrix B
$E[g(X)]$	Expected value of $g(X)$ $E[g(X)] = \int_{-\infty}^{\infty} g(\xi) F_X(\xi) d\xi$ where $F_X(\xi)$ is the probability density function of the random variable X
δ_{ij}	Kronecker delta
$\delta(t-t')$	Dirac delta function

SECTION I

INTRODUCTION

1. Subject Matter

This report presents an analysis of methods for determining aerodynamic coefficients of flight vehicles for dynamic test data. This determination is accomplished by finding numerical values of the aerodynamic coefficients appearing in the equations of motion such that the solutions to these equations are adequate representations of the given test data. When this determination has been made, the coefficients are said to have been extracted from the dynamic data.

The emphasis of the analysis is placed on the effect that random measurement error in the data and random disturbances in the system have on the accuracy with which the coefficients for linear and nonlinear systems can be determined. Both deterministic and stochastic system models for extracting the coefficients and determining their uncertainties are considered.

2. Historical Background

The determination of aerodynamic coefficients from dynamic data is generally agreed to have begun with the work of Fowler, Gallop, Lock, and Richmond¹ in the first quarter of this century. Their basic technique was concerned with determining moment and damping characteristics of artillery shells by firing these projectiles through spaced cards and reconstructing the pitch and yaw angle time histories by observing the obliqueness of the holes that resulted when the shells passed through the cards. Their technique of data measurement is still in use at some ballistic ranges to this date.²

Nielsen and Synge³ later clarified the linear theory of Fowler et al in their work during, and immediately following, World War II.

Coefficient extraction techniques that are currently in use at many ballistic and wind tunnel facilities evolved from the work of Murphy^{4,5} and Nicolaidis,⁶ both of whom have considered various combinations of degrees of freedom for a variety of flight vehicles. Their aerodynamic coefficient extraction techniques feature a least squares fit to dynamic data of the exact solutions of linear equations of motion, or approximate closed form solutions of slightly nonlinear equations of motion using the quasi-linearization technique of Kryloff and Bogoliuboff.⁷ Ex-

tensions of the work of Murphy and Nicolaides, particularly for more complex nonlinear systems, have been made by Eikenberry and Ingram.⁹

The requirement for closed form solutions of the equations of motion has recently been eliminated by the formulation of least squares techniques which fit numerical solutions of the equations of motion to dynamic data. The minimization of the least squares criterion function involves differential corrections, as do the techniques of Murphy and Nicolaides; the required partial derivatives are determined by numerically integrating parametric differential equations which are derived from the equations of motion. Contributions to this numerical coefficient extraction technique have been made independently by Chapman and Kirk,¹⁰ Knafler,¹¹ Goodman^{12,13} and Meissinger,¹⁴ although the method is usually referred to as the Chapman-Kirk technique. The computational requirements of this technique are sometimes extensive,¹⁵ but this is usually outweighed by the fact that it can be used to analyze highly nonlinear aerodynamic forces and moments.¹⁶

All of the coefficient extraction techniques that have been discussed to this point are deterministic in nature in that the modeling of the equations of motion does not account for random disturbances in the system. Most angular or translational motion data obtained from ballistic ranges or wind tunnel dynamic tests have, nevertheless, been affected by random system noise and random measurement errors. The effects of noise of these types on the accuracy of coefficients extracted from dynamic data using deterministic techniques are of current interest.¹⁷ With the exception of Ingram's⁹ partial analysis of noise effects on Eikenberry's "Wobble" program,⁸ however, little has been reported concerning these effects. One of the primary goals of the research reported here is to show some effects of random measurement errors and system noise on the accuracy of aerodynamic coefficients extracted from dynamic data using the most general of the various deterministic techniques, that of Chapman and Kirk.

In recent years increasing attention has been given to parameter and state variable estimation through the use of stochastic modeling of the physical system of interest. Most of this work has been done by optimal control specialists and is a direct result of pioneering contributions in linear filtering theory by Kalman^{18,19} and Bucy.¹⁹ The Kalman filter provides estimates of the states of noisy, linear dynamic systems as well as estimates of the state variable uncertainties. Bryson and Ho,²⁰ among others, have extended the Kalman filter theory to include estimates of the states for nonlinear systems.

Mehra²¹ has recently proposed a maximum likelihood technique for the determination of aerodynamic coefficients from dynamic data using the Kalman filter for linear systems and the extended Kalman filter when the system is nonlinear. Results verifying his proposals are, as yet, unavailable.

An additional maximum likelihood technique has been developed by Grove et al²² and has been used with modest success by Suit.²³

A stochastic approach to the coefficient extraction problem using an extended Kalman filter with parameter augmented state vector is developed in this report. Results obtained with this filter are also presented.

3. Scope of the Report

The remainder of this report is devoted to an analysis of the effects of random system noise and measurement errors on coefficients extracted from dynamic data using the method of Chapman and Kirk, as well as the development and evaluation of an extended Kalman filter for solving essentially the same problem. The equation of motion used in the analysis is one describing the one-degree-of-freedom pitching motion of a rigid body. Pitching moments that are both linear and nonlinear functions of angle-of-attack are considered.

The theory of the Chapman-Kirk technique is recounted briefly in Section II, followed by the analysis of noise effects on this method in Section III.

Developments leading to the formulation of the extended Kalman filter to be used are given in Section IV. Results obtained with the extended filter are presented and discussed in Section V.

A summary is given in Section VI.

Computer program listings and definitions of program variables are given in the two appendices.

SECTION II

THE COEFFICIENT EXTRACTION TECHNIQUE OF CHAPMAN AND KIRK

1. Development of the Extraction Algorithm

A brief reconstruction of the Chapman-Kirk algorithm for extracting aerodynamic coefficients from dynamic data is given in this section. The standard technique for estimating uncertainties in coefficients determined from a least squares fit of a given function to experimental data is also presented.

The extraction technique of Chapman and Kirk has two very basic requirements: (1) the differential equation of motion for the body of interest must be given and (2) a set of experimental data based on the observed motion of the body must be available.

As an example, consider the nonlinear equation of pitching motion for a rigid body

$$\ddot{\alpha} + (C_4 + C_6\alpha^2)\dot{\alpha} + (C_3 + C_5\alpha^2)\alpha = 0 \quad (1)$$

subject to the initial conditions

$$\alpha(0) = C_1$$

$$\dot{\alpha}(0) = C_2$$

where $\dot{\alpha}$ indicates the derivative of α with respect to time. Suppose, also, that the time history of the pitch angle as recorded during some experiment, $\alpha_e(t_i)$, $i=1,2,\dots,m$, is also available.

The technique of Chapman and Kirk is used to determine the values of the C_j ($j=1,2,\dots,6$) in Equation (1) which result in the solution to this equation of motion being a best fit to the test data in a least squares sense. Thus, it is necessary to minimize the least squares criterion function

$$S = \sum_{i=1}^m \left[\alpha_e(t_i) - \alpha_c(t_i) \right]^2 \quad (2)$$

where $\alpha_c(t_i)$ is obtained from the solution to Equation (1).

Now it is well known that in order to determine parameters directly by a least squares fit of a given function to test data, the parameters must appear linearly in the function. This requirement is met in the problem at hand by expanding α_c in a truncated Taylor series about the numerical solution resulting from some initial estimates of the parameters of interest, C_{j0} , i.e.,

$$\alpha_c(t_i) = \alpha_{co}(t_i) + \sum_{j=1}^6 \left(\frac{\partial \alpha}{\partial C_j} \right)_i \Delta C_j, \quad (3)$$

where $\left(\frac{\partial \alpha}{\partial C_j} \right)$ are evaluated for $C_j = C_{j0}$ and $\Delta C_j = C_j - C_{j0}$.

Substituting (3) into (2) yields

$$S = \sum_{i=1}^m \left\{ \alpha_e(t_i) - \alpha_{co}(t_i) - \sum_{j=1}^6 \left[\frac{\partial \alpha}{\partial C_j} \right]_i \Delta C_j \right\}^2. \quad (4)$$

Now, assuming that the $\left[\frac{\partial \alpha}{\partial C_j} \right]$ are known, Equation (4) is a

function of the ΔC_j 's only. Therefore, to determine the values of these coefficients that will minimize this equation, it is necessary to take the partial derivative of Equation (4) with respect to each ΔC_j , set each of the resulting equations to zero and solve for the ΔC_j 's. Carrying out these operations yields the following matrix equation

$$[A] [\Delta C] = [B] \quad (5)$$

where A is, in this case, a 6 X 6 matrix with elements given by

$$A_{jk} = \sum_{i=1}^m \left[\frac{\partial \alpha}{\partial C_j} \right]_i \left[\frac{\partial \alpha}{\partial C_k} \right]_i, \quad (6)$$

ΔC is a 6 X 1 column matrix, or vector, and B is a 6 X 1 column matrix with elements given by

$$B_j = \sum_{i=1}^m \left[\alpha_e(t_i) - \alpha_{co}(t_i) \right] \left[\frac{\partial \alpha}{\partial C_j} \right]_i \quad (7)$$

The solution to Equation (5) is

$$[\Delta C] = [A]^{-1} [B] \quad (8)$$

The solutions for the ΔC_j 's obtained from Equation (8) are exactly correct only if α_c is a linear function of the C_j 's as assumed by Equation (3). This condition of linearity is seldom the case, and the process must be repeated with new initial guesses,

$$C_{j1} = C_{j0} + \Delta C_{j0} \quad (9)$$

until the change in the criterion function [Equation (2)] from one iteration to the next is sufficiently small.

The algorithm just presented requires that time histories of the influence coefficients, $\frac{\partial \alpha}{\partial C_j}$, be available.

These time histories are determined by numerically integrating parametric differential equations which are derived by differentiating the equation of motion with respect to each of the parameters of interest. As an example, the parametric differential equation for C_1 , the initial condition of α is

$$\frac{\partial}{\partial C_1} [\ddot{x} + (C_4 + C_6 \alpha^2) \dot{\alpha} + (C_3 + C_5 \alpha^2) \alpha] = 0$$

$$\frac{\partial \ddot{x}}{\partial C_1} + \frac{\partial C_4}{\partial C_1} \dot{\alpha} + \frac{\partial C_6}{\partial C_1} \alpha^2 \dot{\alpha} + 2C_6 \alpha \dot{\alpha} \frac{\partial \alpha}{\partial C_1} + C_6 \alpha^2 \frac{\partial \dot{\alpha}}{\partial C_1} + \frac{\partial C_3}{\partial C_1} \alpha$$

$$+ C_3 \frac{\partial \alpha}{\partial C_1} + \frac{\partial C_5}{\partial C_1} \alpha^3 + 3C_5 \alpha^2 \frac{\partial \alpha}{\partial C_1} = 0$$

Assuming that the parameters are independent of each other and that the order of differentiation can be reversed, the final form of the desired equation is

$$\frac{d^2}{dt^2} \left[\frac{\partial \alpha}{\partial C_1} \right] + C_6 \alpha^2 \frac{d}{dt} \left[\frac{\partial \alpha}{\partial C_1} \right] + (C_3 + 3C_5 \alpha^2 + 2C_6 \alpha \dot{\alpha}) \left[\frac{\partial \alpha}{\partial C_1} \right] = 0 \quad (10)$$

subject to the initial conditions

$$\frac{\partial \alpha(0)}{\partial C_1} = \frac{\partial C_1}{\partial C_1} = 1$$

$$\frac{\partial \dot{\alpha}(0)}{\partial C_1} = \frac{\partial C_2}{\partial C_1} = 0 \quad (11)$$

The complete set of parametric differential equations for the given equation of motion [Equation (1)] is given in Section III.

In summary, the process for extracting numerical coefficients from test data given the system model [Equation (1)] and criterion function [Equation (2)] is

1. Estimate the numerical values of the C_j 's.
2. Integrate Equation (1) to obtain $\alpha_{CO}(t_i)$.
3. Determine $\left[\frac{\partial \alpha}{\partial C_j} \right]_i$.
4. Solve Equation (8) for the ΔC_j 's.
5. Repeat the process with $C_{j1} = C_{j0} + \Delta C_{j0}$ until the change in Equation (2) is sufficiently small.

2. Estimation of Extracted Parameter Uncertainties

The estimation of the uncertainties, or standard deviations, of parameters that have been determined by the least squares fitting of a given function to test data is a well-known result available in a variety of references (see, for example, References 24 or 25).

The least square parameter uncertainties are estimated in this report by

$$\sigma_j = \sqrt{A_{jj}^{-1}} \cdot \sqrt{\frac{S}{m-K}} \quad (12)$$

where A_{jj}^{-1} is the j th diagonal element of the inverse Grammian matrix [Equations (6) and (8)], S is the sum of the squares of the residuals as given by Equation (2), m is the total number of data points, and K is the total number of parameters being determined by the fit.

SECTION III

ANALYSIS OF THE CHAPMAN-KIRK COEFFICIENT EXTRACTION TECHNIQUE

1. Coefficient Extraction Computer Program

This section provides a detailed description of a one-degree-of-freedom coefficient extraction computer program based on the previously described iterative process of Chapman and Kirk. This program considers the pitching motion of a symmetric missile about a fixed point. It is used to determine values of static pitching moment coefficient derivatives, pitch damping coefficients, and a trim term so that the solution to the appropriate differential equation of motion is a best fit to test data in a least squares sense.

The program was developed to be used as an economical tool in the evaluation of the sensitivity to noise and convergency sensitivity of the Chapman-Kirk technique when operating in its least complex mode.

a. Computational Equations

(1) Equation of Motion

The complete equation of motion that is used in the program is

$$\ddot{\alpha} + (C_4 + C_6 \alpha^2) \dot{\alpha} + (C_3 + C_5 \alpha^2 + C_7 \alpha^4) \alpha + C_8 = 0 \quad (13)$$

with initial conditions

$$\alpha(0) = C_1, \quad \dot{\alpha}(0) = C_2 \quad (14)$$

where

$$C_3 = - \frac{(C_{m\alpha 0}) q A d}{I} \quad C_4 = - \frac{(C_{mq 0}) q A d^2}{2VI}$$

$$C_5 = - \frac{(C_{m\alpha 2}) q A \dot{\alpha}}{I}$$

$$C_6 = - \frac{(C_{mq 2}) q A \dot{\alpha}^2}{2VI}$$

$$C_7 = - \frac{(C_{m\alpha 4}) q A \dot{\alpha}}{I}$$

$$C_8 = - \frac{(C_{m\delta\alpha}) q A \dot{\alpha}}{I}$$

(2) Parametric Differential Equations

The eight parametric differential equations for the given equation of motion [Equation (13)] and initial conditions [Equation (14)] are of the form

$$\ddot{P}_j + A\dot{P}_j + BP_j = F_j, \quad j = 1, 2, \dots, 8 \quad (15)$$

where

$$P_j = \frac{\partial \alpha}{\partial C_j}, \quad \dot{P}_j = \frac{d}{dt} \frac{\partial \alpha}{\partial C_j}, \quad \ddot{P}_j = \frac{d^2}{dt^2} \frac{\partial \alpha}{\partial C_j}$$

$$A = (C_4 + C_6 \alpha^2)$$

and

$$B = (C_3 + 3C_5 \alpha^2 + 2C_6 \alpha \dot{\alpha} + 5C_7 \alpha^4).$$

The values of the initial conditions $P_j(0)$ and $\dot{P}_j(0)$ as well as the functional forms of the nonhomogeneous term F_j are given in Table I for $j=1, 2, \dots, 8$.

TABLE I. INITIAL CONDITIONS AND NONHOMOGENEOUS
TERMS FOR PARAMETRIC DIFFERENTIAL EQUATIONS

j	$P_j(0)$	$\dot{P}_j(0)$	F_j
1	1	0	0
2	0	1	0
3	0	0	$-\alpha$
4	0	0	$-\dot{\alpha}$
5	0	0	$-\alpha^3$
6	0	0	$-\dot{\alpha}\alpha^2$
7	0	0	$-\alpha^5$
8	0	0	-1

b. Program Description

The program is written in Fortran IV for use primarily on an IBM 360/65. The paragraphs that follow describe the functions of the main program and its four subroutines, the required input data, and the program options. A listing of the complete program is given in Appendix I.

(1) Main Program Functions

The functions of the main program in their approximate order of occurrence are as follows:

- (a) Reads and writes input data.
- (b) Computes initial parameter values, [These are the C_j 's that appear in Equation (13).]
- (c) Calls the numerical integration subroutine ADDUM.
- (d) Writes current parameter values.

- (e) Computes the matrix elements required for incrementing the parameter values.
- (f) Calls the matrix inversion subroutine MINV.
- (g) Computes the sum of the squares of the residuals between the calculated and experimental data points.
- (h) Computes the root-mean-square residual and root-mean-square error (Reference 24) of the current fit to data.
- (i) Computes the estimated standard deviations (Reference 24) of the current parameter values.
- (j) Tests the difference between the current and previous values of the root-mean-square error to determine if the iteration process has converged.
- (k) Computes the incremental changes for the parameters (if convergence has not occurred).
- (l) Computes the updated parameter values (if convergence has not occurred).
- (m) Returns to (c) (if convergence has not occurred).
- (n) Computes the values of the extracted coefficients from the current parameter values (after convergence is achieved).
- (o) Computes the estimated standard deviations of the extracted aerodynamic coefficients (after convergence is achieved).
- (p) Writes extracted aerodynamic coefficients, estimated standard deviations, and the pitch angle output that represents the final fit to the experimental data (after convergence is achieved).

(2) Subroutine Functions

The names and functions of each of the four subroutines that are used in the coefficient extraction program are given below.

ADDUM.-- Subroutine ADDUM integrates the equation of motion and parametric differential equations using a fourth order Runge-Kutta method for starting and a fourth order Adams-Bashforth predictor-corrector method for running. It calls subroutines XDOT and OUT. This subroutine is described in detail in Reference 26.

XDOT.-- Subroutine XDOT computes current values of first derivatives that are required by ADDUM.

OUT.-- Subroutine OUT stores the results of the numerically integrated equation of motion [Equation (13)] and parametric differential equations [Equation (15)].

MINV.-- Subroutine MINV inverts a $K \times K$ matrix using a standard Gauss-Jordan technique and is described in detail in Reference 27.

(3) Required Input Data

The program reads six categories of input data. These categories and the specific data in each are delineated in Appendix I. The format and units of the entries on a specific data card can be determined from the program listing and nomenclature list provided in this appendix.

(4) Program Options

This program has options for extracting various combinations of aerodynamic coefficients from the given test data, in addition to the option of extracting no coefficients and merely integrating the equation of motion. These options are controlled by the numerical value of the number of first order equations to be integrated, N , which is read by the program on the first data card. The various options are given in Table II.

TABLE II. PROGRAM OPTIONS

N	Coefficients to be Extracted
2	No coefficients are extracted. The program integrates the equation of motion for the given input data and prints the results.
8	$\alpha_o, \dot{\alpha}_o, C_{m\alpha o}$
10	$\alpha_o, \dot{\alpha}_o, C_{m\alpha o}, C_{mqo}$
12	$\alpha_o, \dot{\alpha}_o, C_{m\alpha o}, C_{mqo}, C_{m\alpha 2}$
14	$\alpha_o, \dot{\alpha}_o, C_{m\alpha o}, C_{mqo}, C_{m\alpha 2}, C_{mq2}$
16	$\alpha_o, \dot{\alpha}_o, C_{m\alpha o}, C_{mqo}, C_{m\alpha 2}, C_{mq2}, C_{m\alpha 4}$
18	$\alpha_o, \dot{\alpha}_o, C_{m\alpha o}, C_{mqo}, C_{m\alpha 2}, C_{mq2}, C_{m\alpha 4}, C_{m\delta\alpha}$

2. Analysis of the Sensitivity to Noise of the Chapman-Kirk Technique

The evaluation of the sensitivity to noise of the Chapman-Kirk coefficient extraction technique has been approached from several directions. Coefficients have been extracted from numerous sets of computer program generated data containing only measurement errors as well as data containing system noise and measurement errors.

In addition to the use of artificially generated data, some very preliminary work has been done with noisy wind tunnel data obtained from actual experimentation. The design of the experiment used to produce the dynamic data is described by Turner in Reference 28 along with some preliminary results obtained with the one-degree-of-freedom program described earlier in this section of the report.

Since the wind tunnel experimentation is still in a developmental stage, however, the results presented in the

remainder of this section are those obtained from artificially generated data.

a. Generation of Noisy Dynamic Data

The computer program UFNOISE described in detail in Reference 29 was used to generate the dynamic data from which the aerodynamic coefficients were extracted. This program simulates the pitching motion of symmetric missile oscillating in a wind stream about a fixed point, for any given initial pitch angle displacement and initial pitch angle rate, by numerically integrating the equation of motion. The program has two options for simulating system noise; it considers the magnitude and direction of the freestream velocity vector as normally distributed random variables, with programmer set mean values and standard deviations, to determine the system noise perturbation accelerations or it simply selects random perturbation accelerations which have zero mean, are normally distributed, and have a programmer set standard deviation. Both methods are essentially equivalent. New values of the system noise perturbations are randomly selected at each numerical integration step. The program also has the option of simulating random measurement errors of the pitch angle by superimposing normally distributed random noise on the output of the numerical integration.

The basic equation of motion used to generate the noisy dynamic data was a slightly simplified form of Equation (13):

$$\ddot{\alpha} + (C_4 + C_6\alpha^2)\dot{\alpha} + (C_3 + C_5\alpha^2)\alpha = w(t) \quad (16)$$

subject to

$$\alpha(0) = C_1, \quad \dot{\alpha}(0) = C_2 \quad (17)$$

The true values of the physical and aerodynamic constants used in generating the data are given in Table III. The various noise level standard deviations and mean values are discussed in the following paragraphs to which they are pertinent.

TABLE III. TRUE VALUES OF CONSTANTS
USED IN GENERATING DATA

Constant	Value
$d(\text{Ft})$	0.333
$A(\text{Ft}^2)$	0.0873
$I(\text{slug ft}^2)$	0.1080
$V_m(\text{ft/sec})$	500.0
$q(\text{lb/ft}^2)$	297.0
$C_{m\alpha_0}(\text{rad}^{-1})$	-2.00
$C_{m\alpha_2}(\text{rad}^{-3})$	-24.5
C_{mq_0}	-60.0
$C_{mq_2}(\text{rad}^{-2})$	-163.0
$\alpha_0(\text{rad})$	0.5235
$\alpha_0(\text{rad/sec})$	0.0

b. Effects of Random Measurement Errors

Aerodynamic coefficients were extracted from a total of nine sets of data containing only measurement errors. Each data set was made up of 201 discrete points which result from integrating the equation of motion [Equation (16)] in increments of 0.005 second for a total of 1.000 second.

The desired mean value of the measurement errors for each of the nine data sets was 0.0 radian. The desired standard deviations of the random errors were 0.00146 radian, 0.00582 radian, and 0.01745 radian, with three different sets of data being generated at each noise level. These standard deviations are typical of measurement uncertainty for data of this type (Reference 9).

The above standard deviations correspond to percent noise levels of 1.1, 4.4, and 13.2, respectively, where percent noise level is defined in Reference 30 as

$$\text{Percent Noise} = \frac{3\sigma}{\text{Approximate average peak amplitude}} \quad (18)$$

with σ being the standard deviation of interest. The approximate average peak amplitude can be determined in a variety of ways. The value used here is the mean positive peak amplitude for the three cycles that result when the nominal equation of motion is integrated for 1.000 second.

The results of this portion of the analysis are given in Table IV and in Figures 1 through 10.

Table IV is primarily a summary of the percent error in the extracted aerodynamic coefficients together with their normalized estimated standard deviations for the various actual measurement errors, σ_m . The percent error in the extracted coefficients is defined by

$$\text{Percent Error } (C_j) = \frac{(\hat{C}_j - C_j) \cdot 100}{C_j} \quad (19)$$

where \hat{C}_j is the coefficient estimate determined with the program in the fitting of the data and C_j is the true value of the coefficient of interest. The normalized estimated standard deviations are the estimated standard deviations calculated from Equation (12) divided by the true value of the parameter of interest,

$$\text{Percent } \sigma_{jn} = \frac{\sigma_j}{C_j} \cdot 100 \quad (20)$$

Table IV also contains the RMS residual for each fit to the noisy data and the percent noise levels based on the values of σ_m .

TABLE IV. SUMMARY OF RESULTS FOR EXTRACTING AERODYNAMIC COEFFICIENTS
FROM DYNAMIC DATA CONTAINING RANDOM MEASUREMENT ERRORS

Data Set	σ_m (Rad)	Percent Noise	Percent Error in Extracted Coefficients and Normalized Estimated Standard Deviations				RMS Residual (Rad)
			$C_3(\sigma_3)$	$C_4(\sigma_4)$	$C_5(\sigma_5)$	$C_6(\sigma_6)$	
1	0.00158	1.2	0.07(0.39)	-0.48(0.93)	0.09(0.30)	2.24(8.38)	0.00155
2	0.00142	1.1	0.21(0.34)	0.06(0.82)	-0.41(0.26)	-1.85(7.32)	0.00136
3	0.00144	1.1	0.25(0.36)	1.34(0.86)	0.19(0.28)	-10.40(7.70)	0.00143
4	0.00587	4.4	-0.92(1.48)	0.79(3.48)	0.70(1.13)	-4.42(31.26)	0.00582
5	0.00607	4.5	2.44(1.45)	2.48(3.50)	-1.94(1.11)	-11.40(31.41)	0.00589
6	0.00560	4.3	-0.28(1.41)	0.62(3.35)	0.05(1.08)	-12.05(30.01)	0.00556
7	0.01842	14.0	-1.49(4.63)	24.94(10.89)	0.92(3.53)	-226.94(97.51)	0.01810
8	0.01672	12.5	2.90(4.02)	-2.00(9.67)	-4.18(3.00)	5.42(85.06)	0.01622
9	0.01837	14.7	0.48(4.70)	11.44(11.22)	-1.42(3.57)	-84.11(99.54)	0.01906

An example of one of the program fits to a noisy data set is shown on Figure 1. Figures 2 through 5 show the variation of the percent error in the individual extracted coefficients with measurement error. The variation of RMS residual with measurement error is shown on Figure 6. The variations of the normalized estimated standard deviations of the extracted coefficients with measurement error are shown in Figures 7 through 10.

An analysis of Table IV and Figures 2 through 10 reveals the following:

- a. The extracted static pitching moment coefficient derivatives show little or no error for the entire range of measurement errors considered. (See Figures 2 and 3.)
- b. The extracted pitch damping coefficients show some error for the lower noise values ($\sigma_m \leq 0.005$ radian) and deviate significantly for $\sigma_m \approx 0.018$ radian. (See Figures 4 and 5.)
- c. The variation of RMS residual with measurement noise is essentially linear with a slope of unity (Figure 6). This indicates that the RMS residual, as calculated by the coefficient extraction program, is a good estimate of the amount of measurement noise in the data when this is the only type of noise present.
- d. All variations of the normalized estimated standard deviations of the extracted coefficients, σ_{jn} , increase linearly with measurement error (Figures 7 through 10). Approximate values of the slopes of these variations are given in Table V. These normalized uncertainty ratios give the relative (to each other) uncertainty which can be expected when extracting coefficients from data containing measurement noise using the given coefficient extraction program.
- e. The true value of a given coefficient is contained within the interval defined by its extracted value $\pm 3\sigma_j$ for every coefficient extracted from the nine sets of data.

TABLE V. EXTRACTED COEFFICIENT UNCERTAINTY RATIOS

Coefficient (C_j)	Approximate Normalized Uncertainty Ratio ($\Delta\sigma_j/\Delta\sigma_m$)
C_3	2
C_5	2
C_4	5
C_6	50

c. Effects of Random Measurement Errors and System Noise

For this portion of the analysis, attempts were made to extract aerodynamic coefficients from twelve basic sets of dynamic data. Each data set was made up of 201 discrete points resulting from integrating the equation of motion in increments of 0.005 second for 4.00 seconds; the pitch angle was output every four integration steps so that the time between data points was 0.020 second. The time between data points used in this portion of the analysis is different from that previously used. This change was made so that the time between data points would correspond to that used by Turner²⁸ in actual experimentation.

Of the twelve basic data sets, nine contain system noise and measurement errors, whereas three contain only measurement errors. Coefficients were extracted from these latter three to determine if the above-mentioned change in the time increment between data points had any appreciable effect on the extracted coefficients or their uncertainties. The estimates of the uncertainties were improved due primarily to the fact that a longer data record was being analyzed.

The desired mean values of the freestream velocity vector direction fluctuations and measurement errors for all data sets were zero. The desired standard deviation of the measurement errors, σ_m , was 0.00582 radian ($3\sigma_m = 1.0$ degree) for all sets; the desired standard deviation of the velocity magnitude fluctuations, σ_v , was 5.0 ft/sec for all nine data sets containing system noise. The desired standard deviations of the velocity direction, σ_a , were 0.02910 radian ($3\sigma_a = 5$ degrees), 0.05820 radian ($3\sigma_a = 10$ degrees), and 0.11640 radian ($3\sigma_a = 20$ degrees), with three different sets of data being generated at each noise level.

The random velocity fluctuations act as a forcing function for the equation of motion and cause oscillations even if the vehicle has no initial displacement. The resultant maximum amplitudes of these forced oscillations are usually within a certain magnitude or noise band width. The widths of the noise bands for the data used in this analysis are generally equivalent to the $3\sigma_a$ values, and these have been used to calculate the percent system noise values. The approximate average peak amplitude used in the percent noise calculations is the mean positive peak amplitude for the 10 cycles that result when the nominal equation of motion is integrated for approximately four seconds.

The results of this portion of the analysis are given in Table VI and in Figures 11 through 19.

Table VI is a summary of the percent error in the extracted aerodynamic coefficients and their normalized estimated standard deviations along with the various noise level standard deviations, percent system noise levels, and the RMS residual of each fit to a noisy data set. The percent system noise levels are similar to those encountered experimentally by Turner in Reference 28.

TABLE VI. SUMMARY OF RESULTS FOR EXTRACTING AERODYNAMIC COEFFICIENTS FROM DYNAMIC DATA CONTAINING
RANDOM MEASUREMENT ERRORS AND SYSTEM NOISE

Data Set	σ_m (Rad)	σ_v (ft./sec)	σ_a (Rad)	Percent System Noise	Percent Error in Extracted Coefficients and Normalized Estimated Standard Deviations				RMS Residual (ft./sec)
					$C_3(\sigma_3)$	$C_4(\sigma_4)$	$C_5(\sigma_5)$	$C_6(\sigma_6)$	
10	0.00527	0.00	0.00	0.00	-0.27(0.25)	-0.84(0.93)	0.68(0.65)	5.72(12.92)	0.00520
11	0.00589	0.00	0.00	0.00	-0.34(0.28)	0.29(1.04)	1.51(0.73)	-3.60(14.43)	0.00574
12	0.00574	0.00	0.00	0.00	0.42(0.27)	-0.36(1.01)	-0.54(0.69)	4.40(13.88)	0.00563
13	0.00611	5.05	0.02910	41.4	0.55(0.64)	8.25(2.75)	9.55(1.88)	329.59(43.58)	0.01238
14	0.00597	5.18	0.03038	43.1	4.19(0.63)	5.84(2.41)	1.35(1.63)	-44.80(34.40)	0.01244
15	0.00527	4.88	0.02828	40.2	2.87(0.55)	20.73(2.10)	-2.80(1.24)	-247.20(26.39)	0.01026
16	0.00603	5.01	0.005516	78.4	2.71(1.24)	-42.80(3.95)	6.54(3.64)	605.77(66.79)	0.02806
17	0.00562	4.90	0.06120	86.8	14.44(1.32)	-0.90(5.35)	5.44(3.87)	197.10(22.51)	0.02627
18	0.00600	4.96	0.05662	80.3	1.41(0.70)	-35.67(2.33)	14.59(2.30)	586.34(42.19)	0.01488
19	0.00572	5.08	0.11262	152.9			DIVERGED		
20	0.00547	4.85	0.11706	158.9			DIVERGED		
21	0.00571	4.96	0.12058	163.7			DIVERGED		

Figure 11 is an example of one of the program fits to a noisy data set. Figures 12 through 15 show the variation of percent error for the extracted coefficients with system noise. The variations of the normalized estimated standard deviations of the extracted coefficients with system noise level are shown on Figures 16 through 19.

An analysis of Table VI and Figures 12 through 19 reveals the following:

- a. The errors in the linear static pitching moment coefficient derivative parameter, C_3 , are generally less than 5 percent for the entire range of system noise considered. The errors in the corresponding nonlinear term, C_5 , are generally less than 10 percent. (See Figures 12 and 13.)
- b. The pitch damping parameters show significant error for all nonzero system noise levels. The extracted values of the nonlinear term, C_6 , are sometimes in error by several orders of magnitude at the higher noise levels. (See Figures 14 and 15.)
- c. The estimated standard deviations of the extracted coefficients generally increase as the system noise increases. The estimated standard deviations of coefficients extracted from dynamic data containing both system noise and measurement errors are generally too small and do not reflect the true uncertainty of the extracted coefficients as was true in the previous case when the data contained only measurement errors. (See Table VI and Figures 16 through 19.)
- d. All attempts at extracting the complete set of coefficients from data with a system noise band of approximately 20 degrees failed (see Table VI). The failures resulted when a value of C_6 was eventually calculated by the iterative process which caused the solution to the equation of motion to diverge.
- e. The divergence problem can sometimes be circumvented by not attempting to extract C_{mq2} from extremely noisy data. The resulting coefficients that are extracted, C_{ma0} , C_{ma2} , and C_{mq0} , have accuracies comparable to those extracted from noisy data with a ± 10 degrees system noise band.

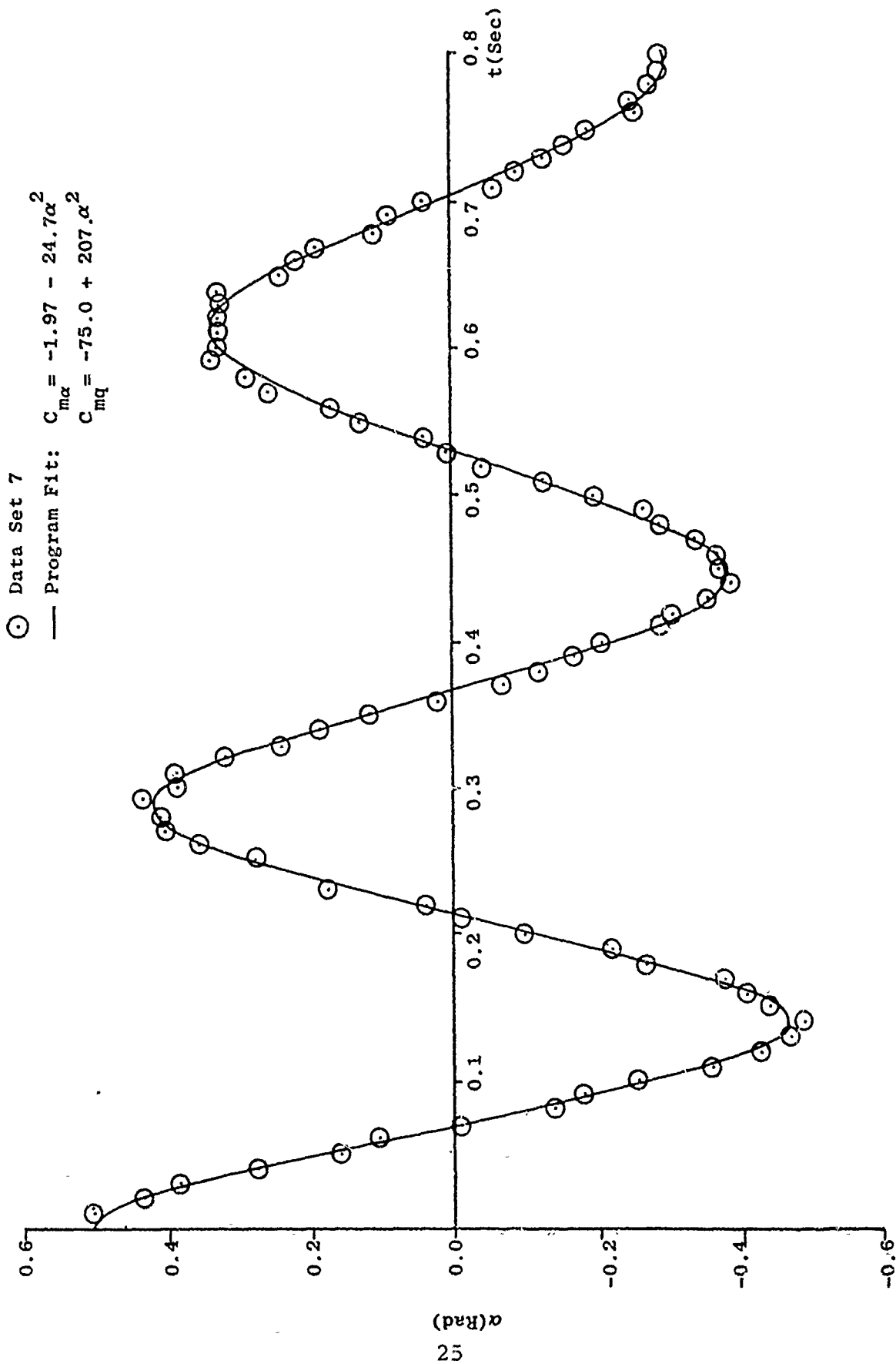


Figure 1. Example of Program Fit to Data Containing Measurement Errors

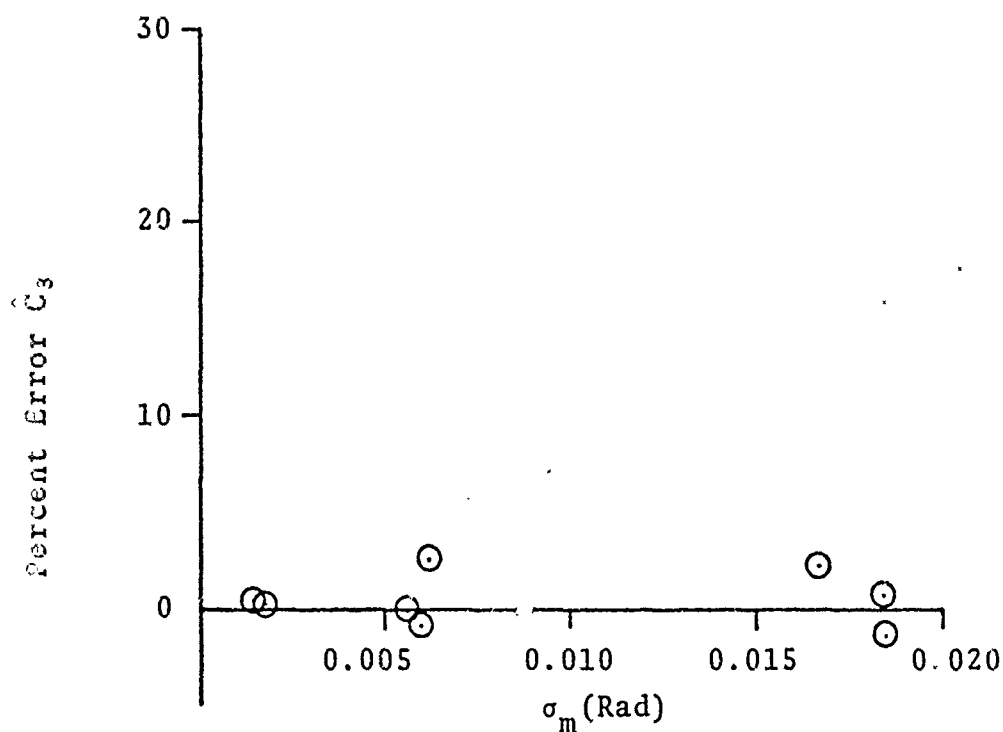


Figure 2. Variation of Percent Error in Extracted C_3 with Measurement Error

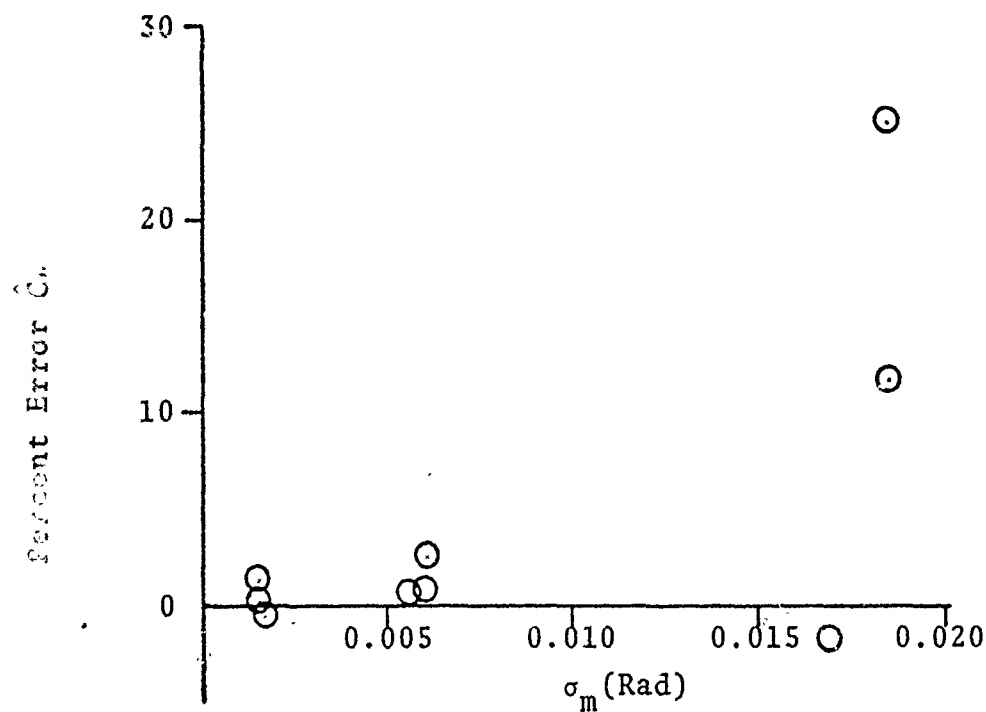


Figure 3. Variation of Percent Error in Extracted C_4 with Measurement Error

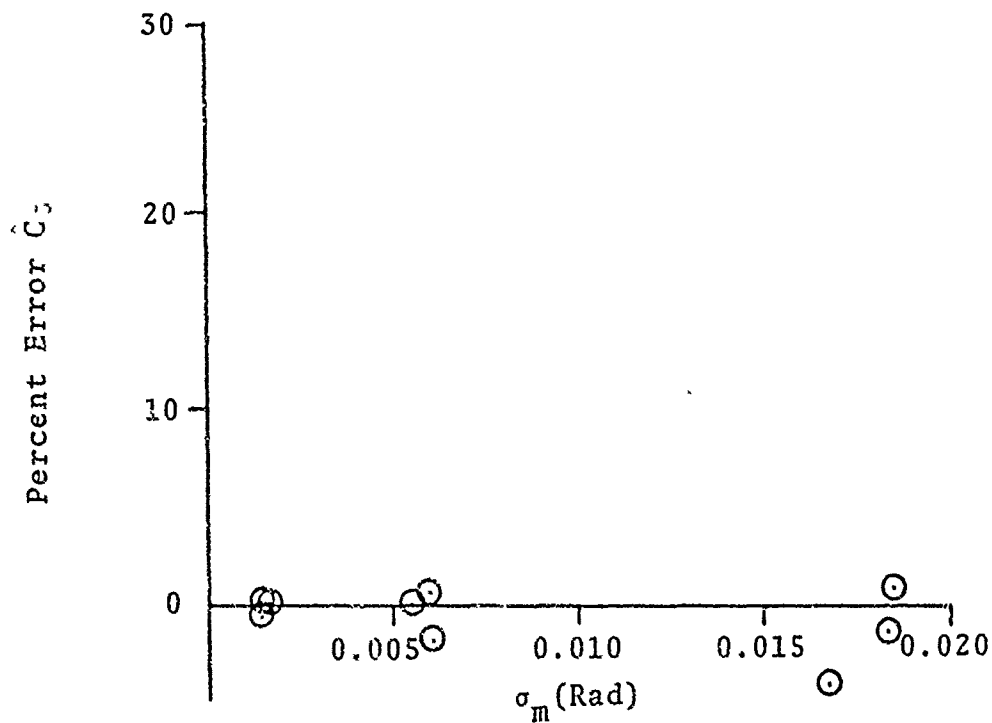


Figure 4. Variation of Percent Error in Extracted C_5 with measurement Error

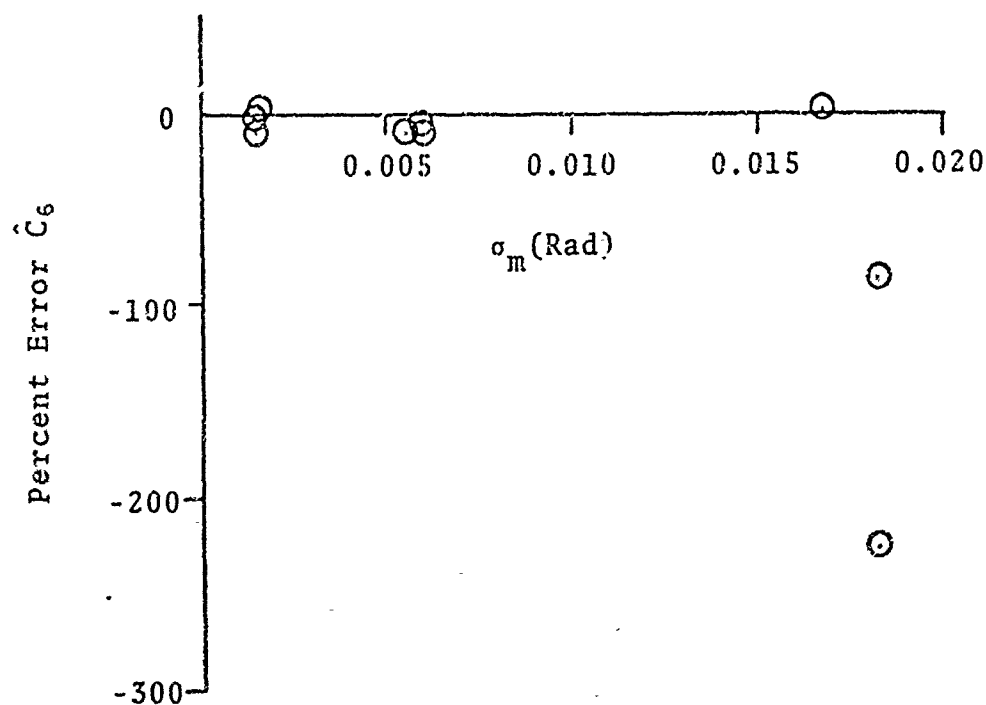


Figure 5. Variation of Percent Error in Extracted C_6 with Measurement Error

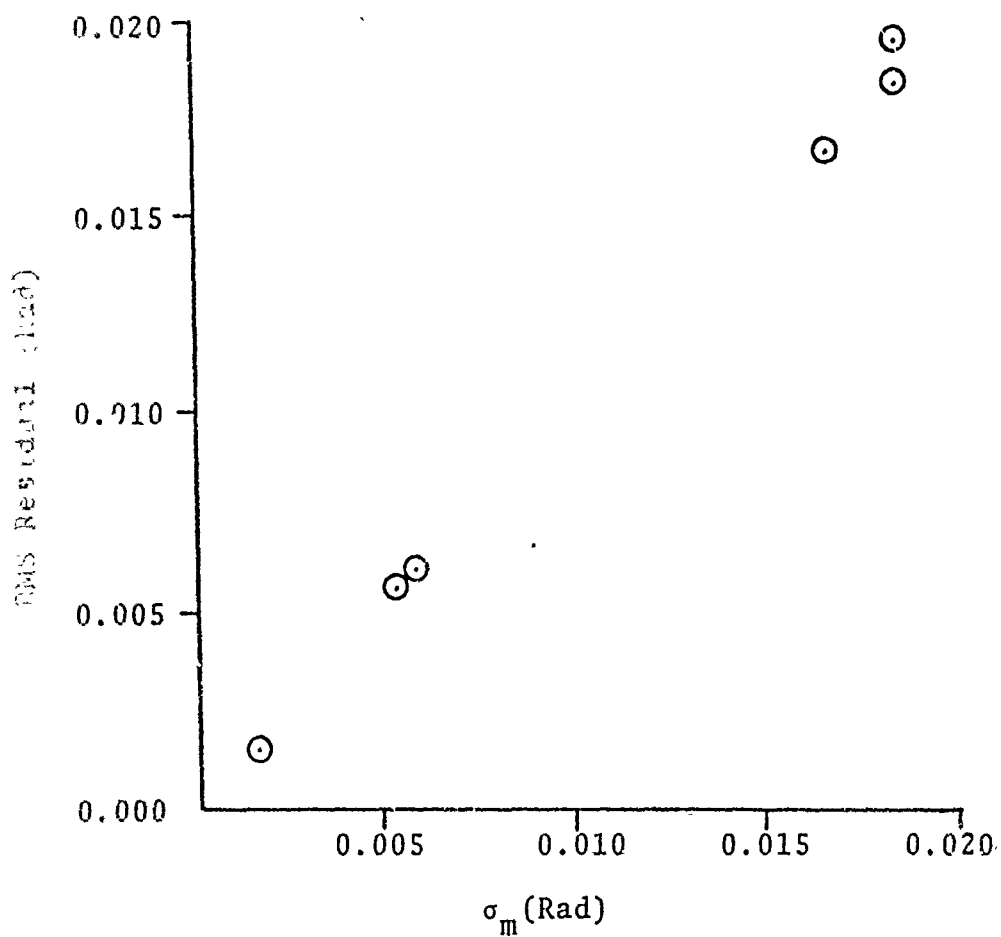


Figure 6. Variation of RMS Residual with Measurement Error

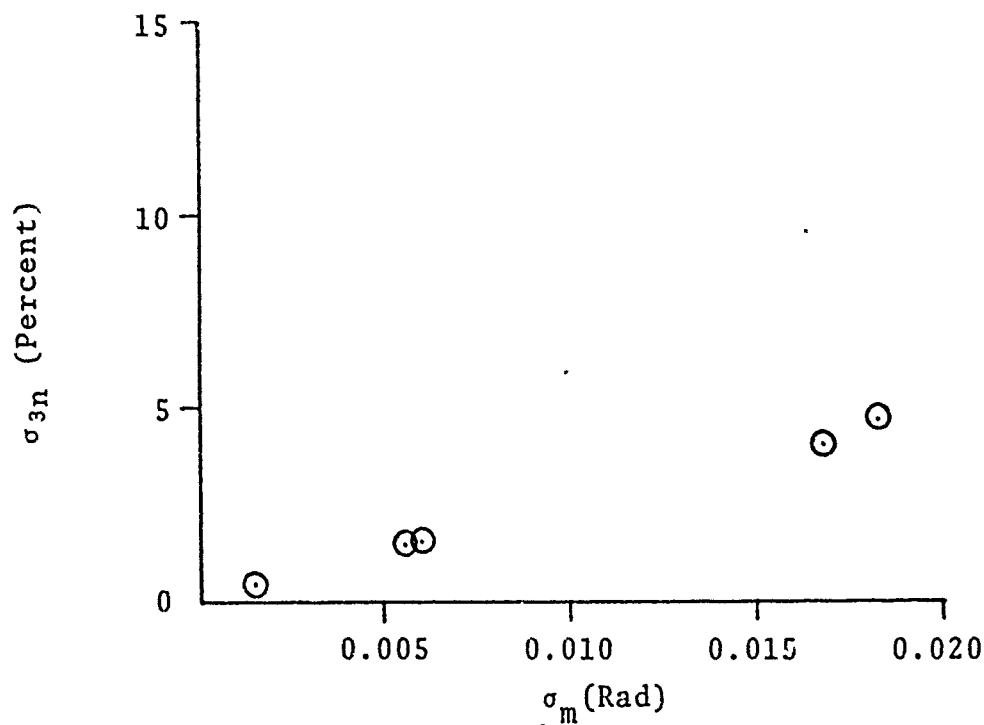


Figure 7. Variation of Normalized Estimated Standard Deviation of \hat{C}_3 with Measurement Error

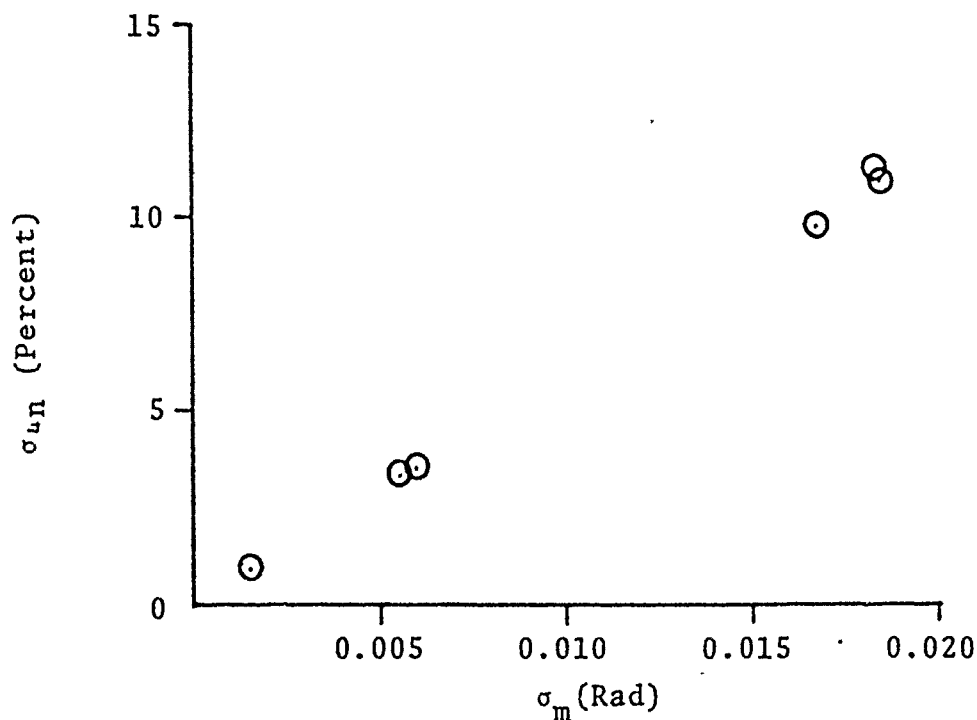


Figure 8. Variation of Normalized Estimated Standard Deviation of \hat{C}_4 with Measurement Error

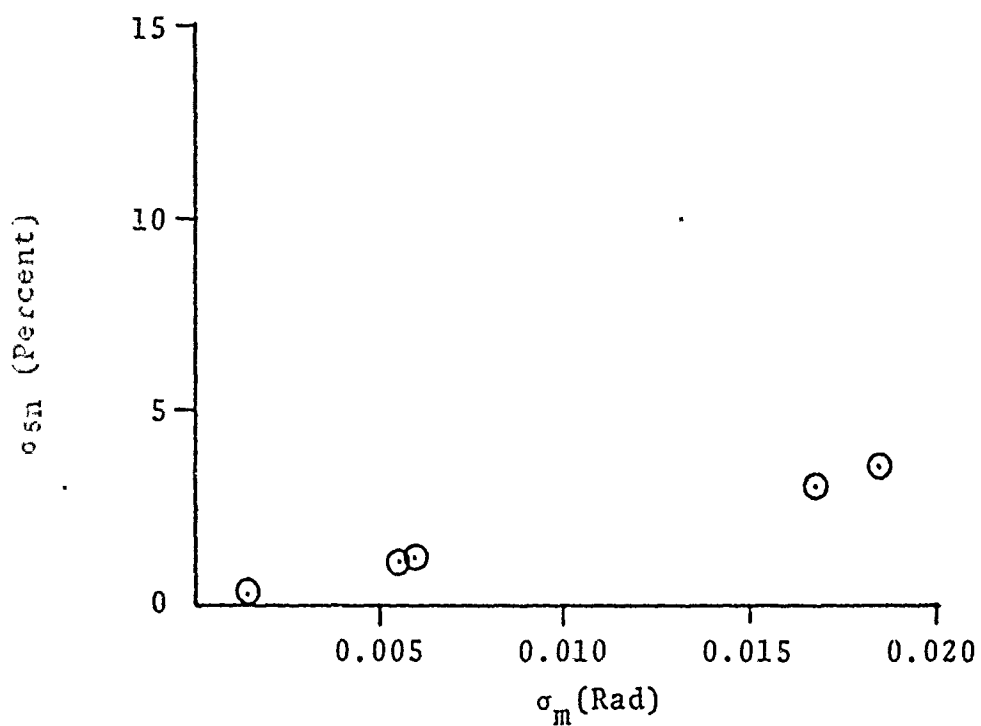


Figure 9. Variation of Normalized Estimated Standard Deviation of \hat{C}_5 with Measurement Error

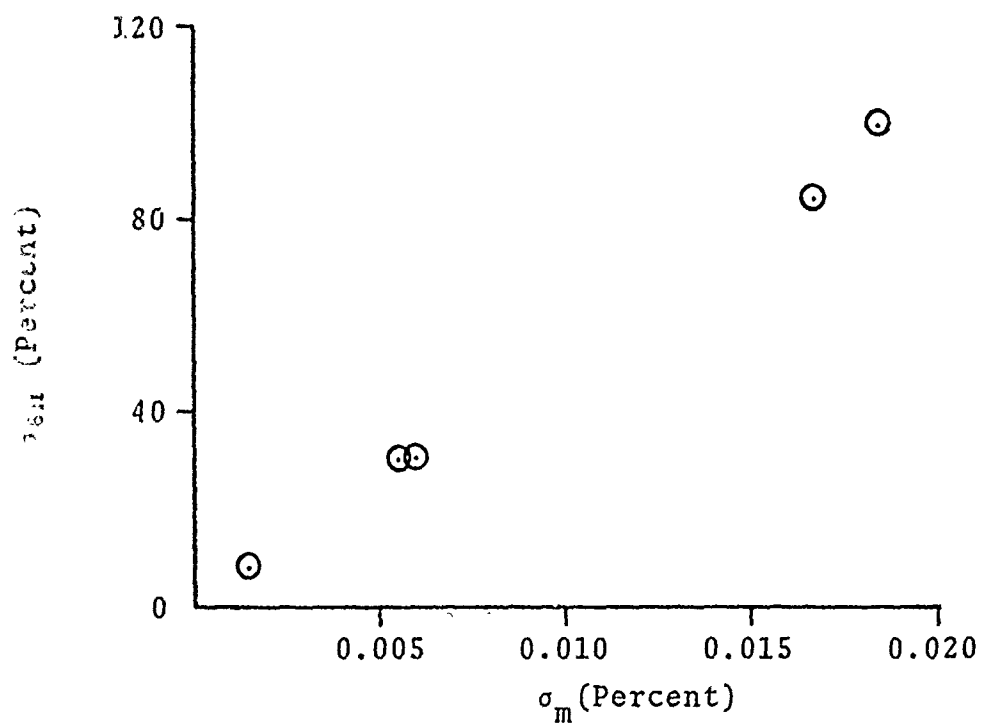


Figure 10. Variation of Normalized Estimated Standard Deviation of \hat{C}_6 with Measurement Error

○ Data set 17
 — Program Fit:
 $C_{m\alpha} = -2.29 - 25.8\alpha^2$
 $C_{mq} = -59.5 - 485\alpha^2$

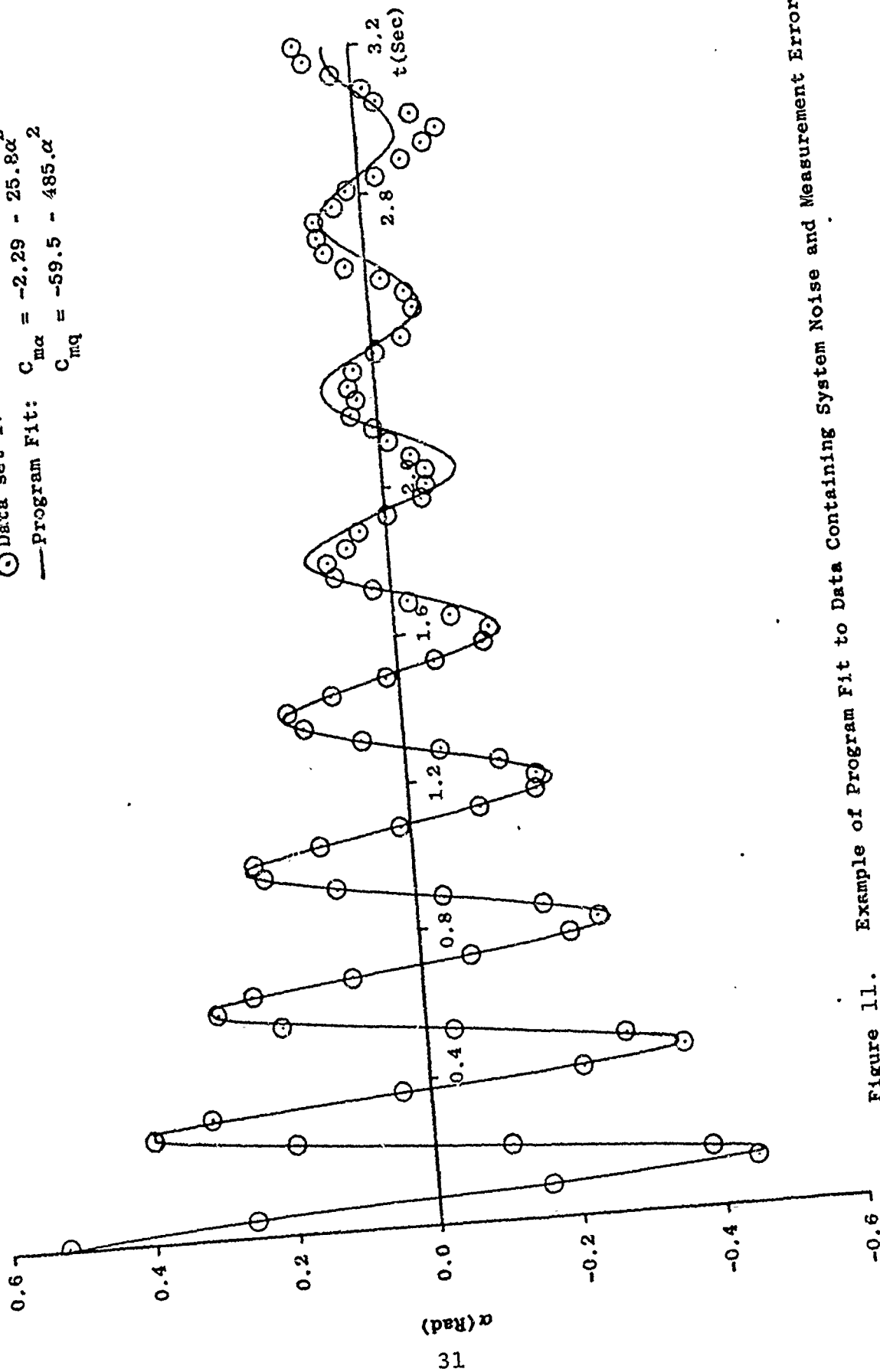


Figure 11. Example of Program Fit to Data Containing System Noise and Measurement Errors

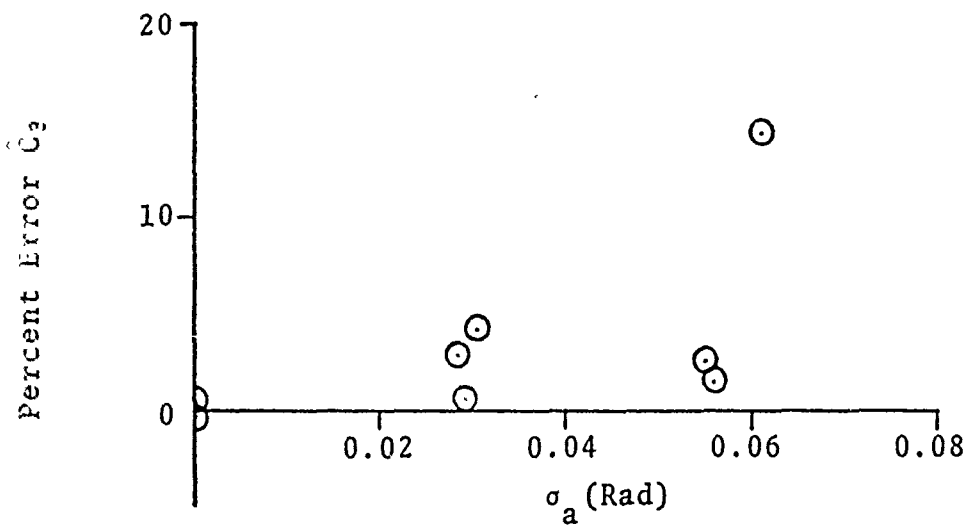


Figure 12. Variation of Percent Error in Extracted C_3 with System Noise

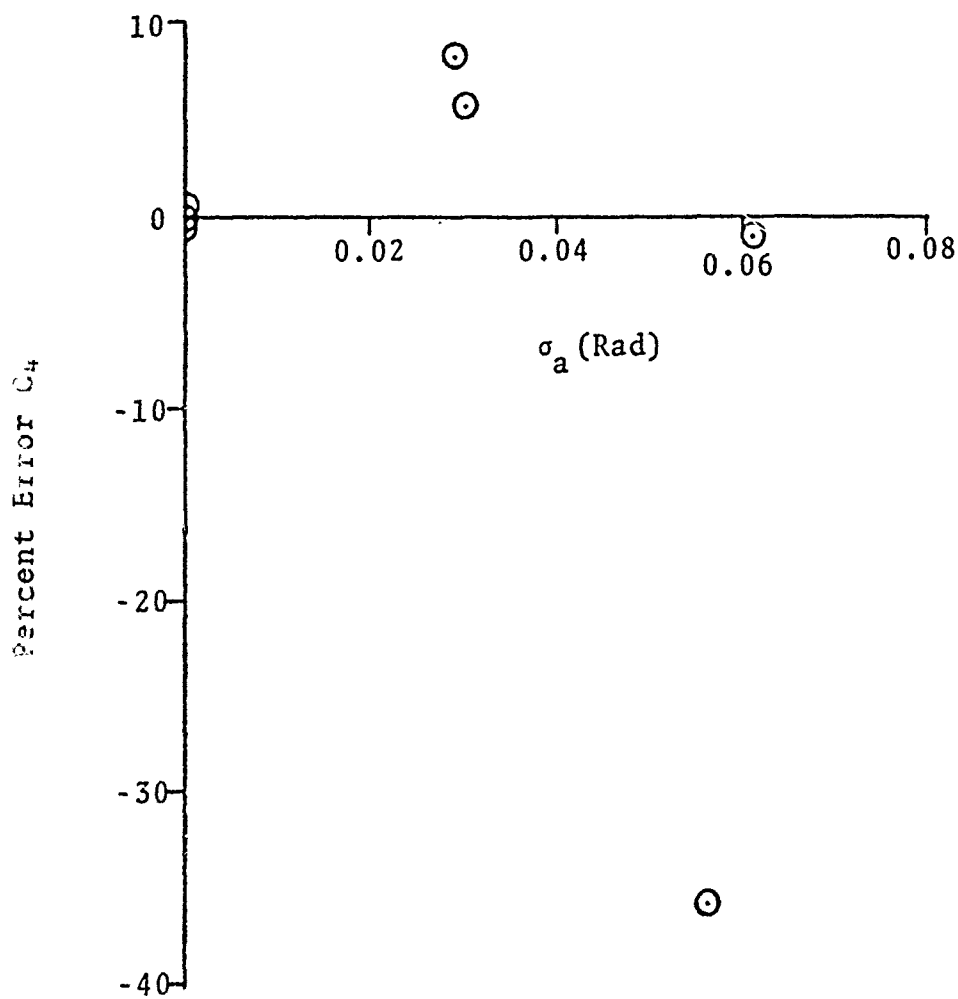


Figure 13. Variation of Percent Error in Extracted C_4 with System Noise

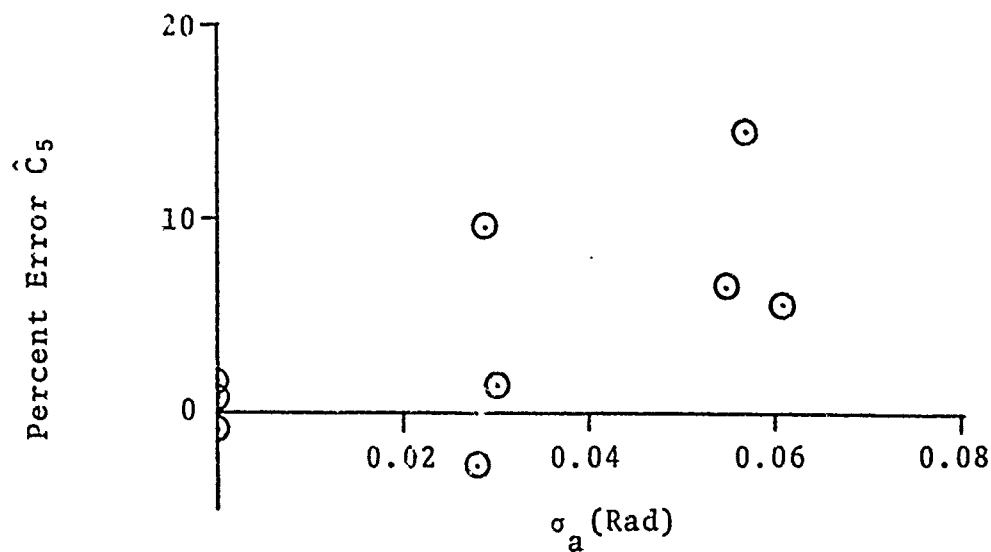


Figure 14. Variation of Percent Error in Extracted C_5 with System Noise

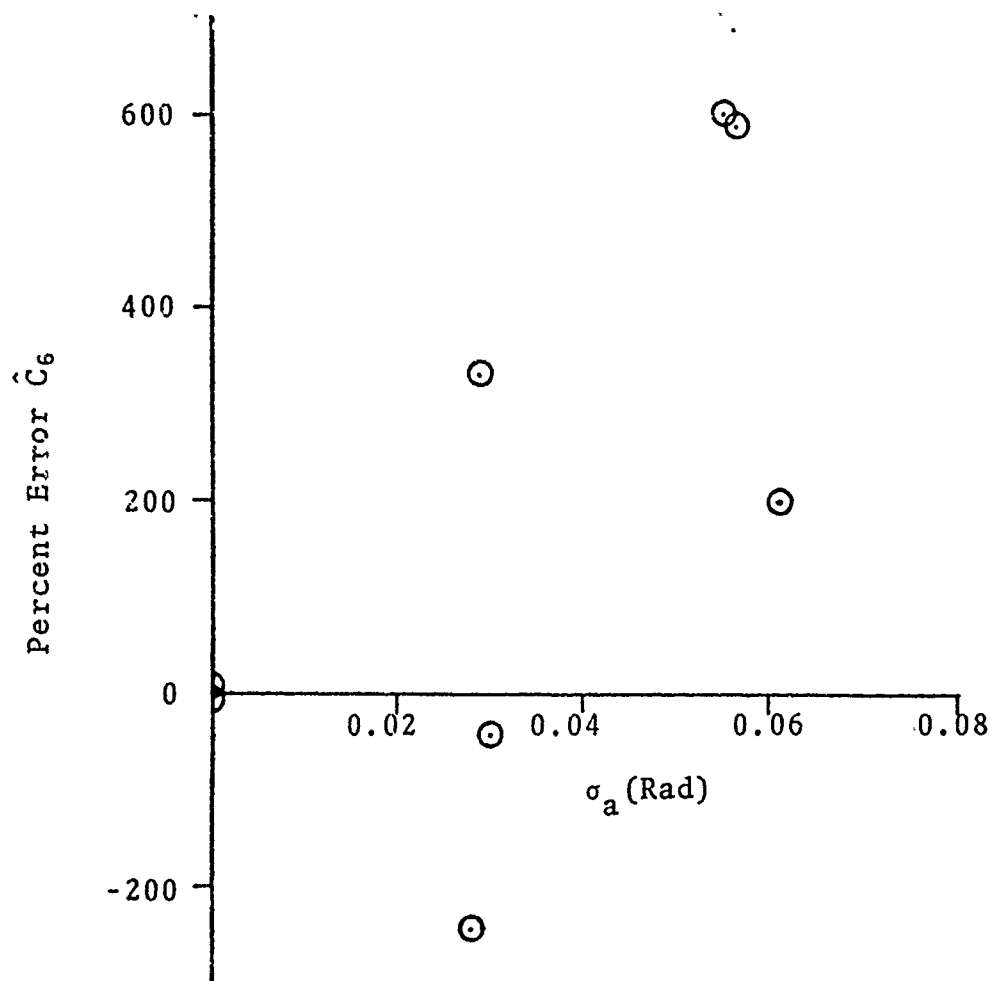


Figure 15. Variation of Percent Error in Extracted C_6 with System Noise

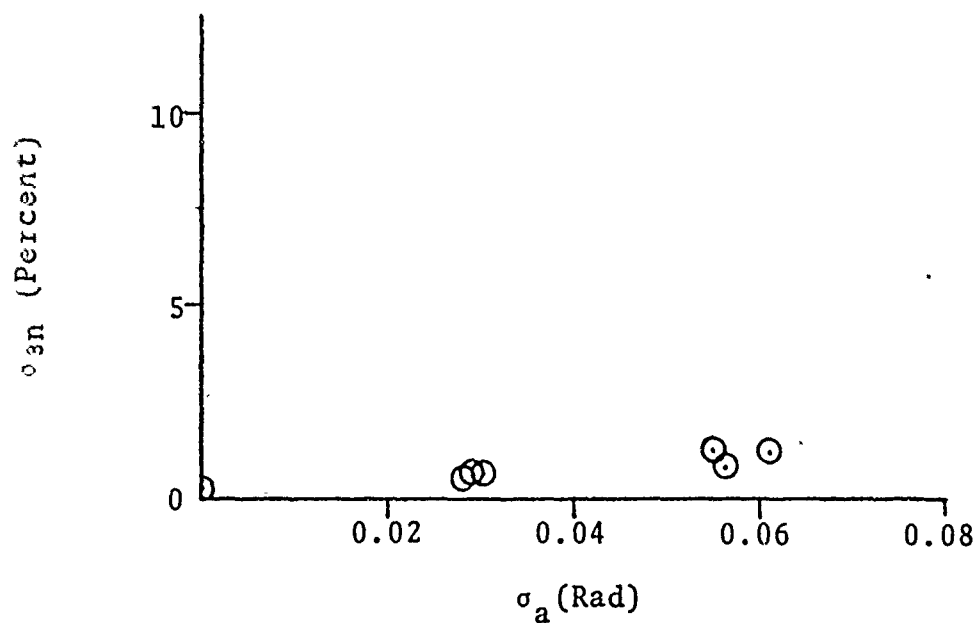


Figure 16. Variation of Normalized Estimated Standard Deviation of \hat{C}_3 with System Noise

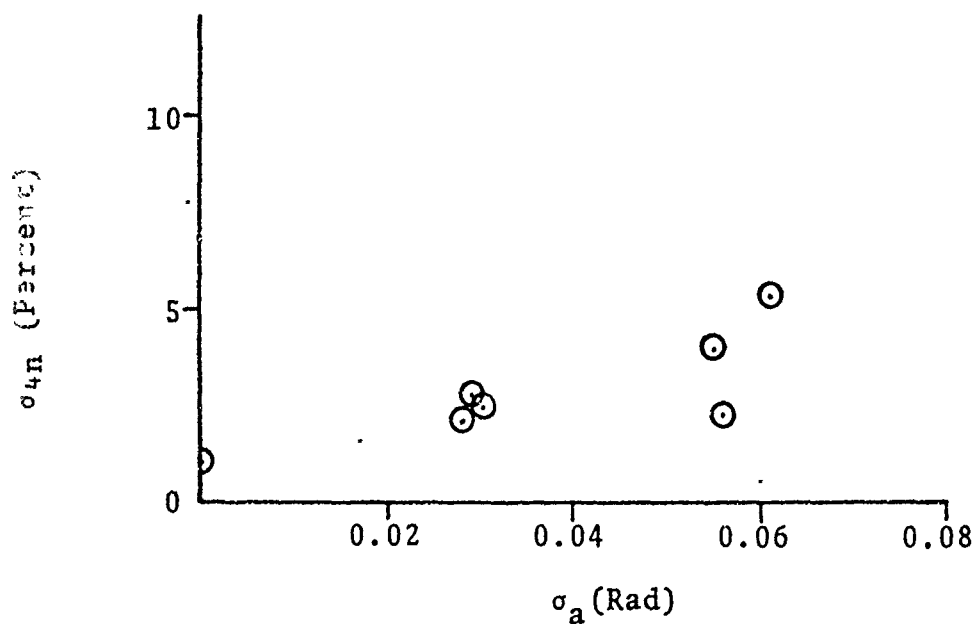


Figure 17. Variation of Normalized Estimated Standard Deviation of \hat{C}_4 with System Noise

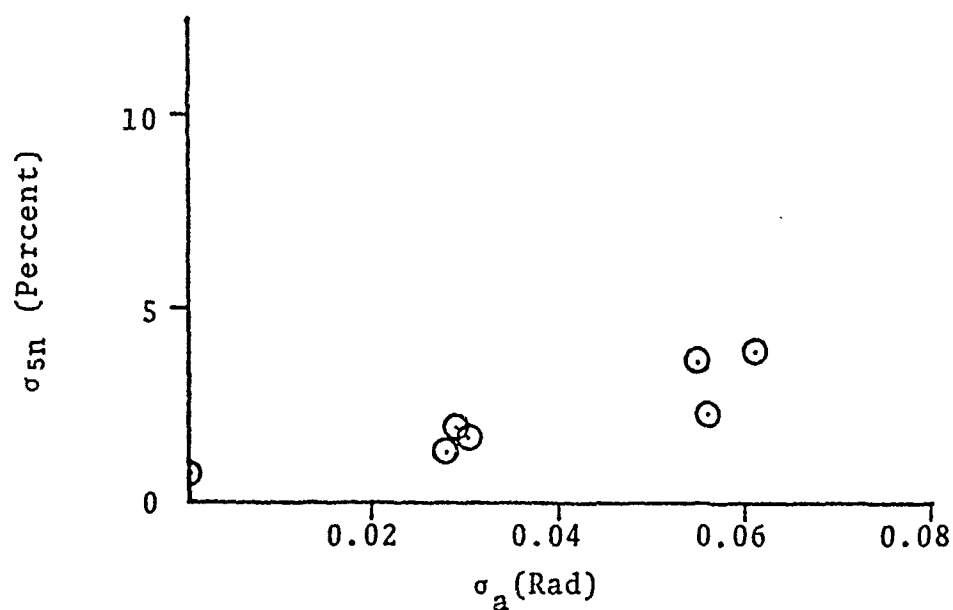


Figure 18. Variation of Normalized Estimated Standard Deviation of \hat{C}_5 with System Noise

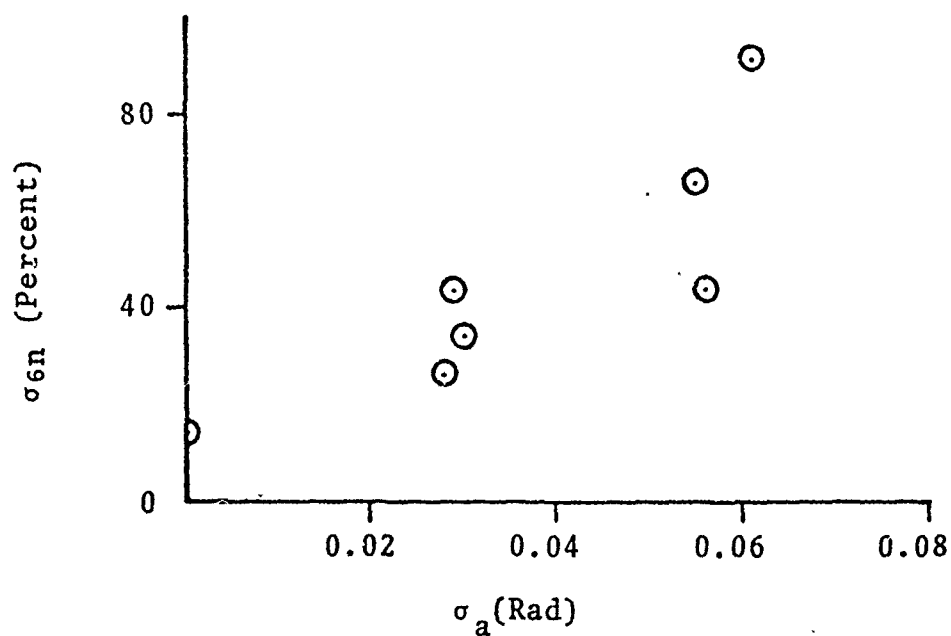


Figure 19. Variation of Normalized Estimated Standard Deviation of \hat{C}_6 with System Noise

SECTION IV

DEVELOPMENT OF THE EXTENDED KALMAN FILTER FOR ESTIMATING PARAMETERS AND THEIR UNCERTAINTIES

1. Introduction

This chapter presents an abridged derivation of the Kalman filter for discrete and continuous linear systems, followed by a general statement of the extended Kalman filter for continuous nonlinear systems. The extended filter is then used as a base for the development of a specific algorithm for estimating states and parameters of the second order equation of pitching motion for a missile being forced by random disturbances.

2. Development of the Kalman Filter

The original derivation of the Kalman filter was presented by Kalman¹⁸ in 1960 for multistage systems making discrete linear transitions from one stage to another. Kalman and Bucy¹⁹ gave the analogous development for continuous linear systems approximately one year after the first work was published. The purpose of the linear filter is to provide estimates of the state of a system by making use of measurements of all, or some, of the state vector components of the system. The system is assumed to be operating in the presence of random disturbances, the statistical properties (i.e., mean and variance) of which are known. The measurements of the state vector components of the system are also assumed to have random errors of known statistics.

In addition to the original derivations of Kalman and Bucy, several other methods offering various degrees of insight but leading to the same results are available. A brief development taken primarily from Bryson and Ho²⁰ is presented here. Other derivations or developments of the filter equations are given by Jazwinsky,³¹ Sorenson,³² and Barham.³³

The treatment given here starts with a static system and is extended for a single-stage linear transition which leads directly to linear multistage process. Finally, by making use of a limiting process, the desired form of the equations for a continuous linear dynamic system is given.

a. Static System

The problem at hand is to estimate the n -component state vector X of a static system using the p -component measurement vector, z , containing random errors, v , which are independent of the state. The measurement vector can be represented as

$$z = HX + v \quad , \quad (21)$$

where H is a known $p \times n$ matrix. Conditions on the measurement errors are

$$E(v) = 0 \quad (22)$$

$$E(vv^T) = R \quad , \quad (23)$$

where R is a known matrix of dimension p . It is assumed that a prior estimate of the state, designated as \bar{X} , is available and also that the covariance of the prior estimate, M, is known. Thus

$$M = E[(X - \bar{X})(X - \bar{X})^T] \quad , \quad (24)$$

where M is of dimension n .

The desired estimate of X, taking into account the measurement, z, is the weighted-least-squares estimate, \hat{X} , which minimizes

$$J = [(X - \bar{X})^T M^{-1} (X - \bar{X}) + (z - HX)^T R^{-1} (z - HX)] \quad (25)$$

Differentiating Equation (25) with respect to X, setting the resulting equation to zero, and solving for X yields

$$\hat{X} = \bar{X} + PH^T R^{-1} (z - H\bar{X}) \quad , \quad (26)$$

where

$$P^{-1} = M^{-1} + H^T R^{-1} H \quad . \quad (27)$$

The quantity P is the covariance matrix of the error in the state estimate after measurement, X , i.e.,

$$P = E[(X - \hat{X})(X - \hat{X})^T] \quad . \quad (28)$$

b. Single-Stage Transitions

It is now desired to estimate the state of a system that makes a discrete transition from state 0 to state 1 according to the linear relation

$$X_1 = \phi_0 X_0 + \Gamma_0 w_0 \quad , \quad (29)$$

where ϕ_0 is a known transition matrix of dimension n and Γ_0 is a known $n \times r$ matrix. The mean and variance of the random forcing vector are assumed to be known and are given by

$$E(w_0) = \bar{w}_0 \quad , \quad E[(w_0 - \bar{w}_0)(w_0 - \bar{w}_0)^T] = Q_0 \quad . \quad (30)$$

The statistical properties of the random state vector are assumed to be known initially as

$$E(X_0) = \hat{X}_0 \quad , \quad E[(\hat{X}_0 - X_0)(\hat{X}_0 - X_0)^T] = P_0 \quad . \quad (31)$$

It is also assumed that X_0 and w_0 are independent. From this information X_1 is a random vector whose mean value and covariance are

$$\bar{X}_1 = \phi_0 \hat{X}_0 + \Gamma_0 \bar{w}_0 \quad (32)$$

$$M_1 = \phi_0 P_0 \phi_0^T + \Gamma_0 Q_0 \Gamma_0^T \quad . \quad (33)$$

Suppose now that measurements of the state are made after the transition to state 1; then from Equations (26) and (27) the best estimate of X_1 is

$$\hat{X}_1 = \bar{X}_1 + P_1 H_1^T R^{-1} (z - H_1 \bar{X}_1) \quad , \quad (34)$$

where

$$\begin{aligned} P_1 &= (M_1^{-1} + H_1^T R^{-1} H_1)^{-1} \\ &= M_1 - M_1 H_1^T (H_1 M_1 H_1^T + R_1^{-1})^{-1} H_1 M_1 \end{aligned} \quad (35)$$

c. Linear Multistage Processes

The developments of the two preceding subsections can be extended for linear, stochastic, multistage processes. Given the following difference equation system model

$$X_{i+1} = \phi_i X_i + \Gamma_i w_i \quad (36)$$

where

$$E(X_0) = \bar{x}_0 \quad (37)$$

$$E(w_i) = \bar{w}_i \quad (38)$$

$$E[(X_0 - \bar{x}_0)(X_0 - \bar{x}_0)^T] = M_0 \quad (39)$$

$$E[(w_i - \bar{w}_i)(w_j - \bar{w}_j)] = Q_i \delta_{ij} \quad (40)$$

$$E[(w_i - \bar{w}_i)(X_0 - \bar{x}_0)^T] = 0 \quad (41)$$

and measurements of the state

$$z_i = H_i X_i + v_i \quad (42)$$

where

$$E(v_i) = 0 \quad (43)$$

$$E(v_i v_j^T) = R_i \delta_{ij} \quad (44)$$

$$E[(w_i - \bar{w}_i) v_j^T] = 0, \quad E[(X_0 - \bar{X}_0) v_i^T] = 0, \quad (45)$$

the estimate of the state is

$$X_i = \bar{X}_i + K_i (z_i - H_i \bar{X}_i) \quad (46)$$

where

$$\bar{X}_{i+1} = \phi_i \bar{X}_i + \Gamma_i \bar{w}_i \quad (47)$$

$$K_i = P_i H_i^T R_i^{-1} \quad (48)$$

$$\begin{aligned} P_i &= (M_i^{-1} + H_i^T R_i^{-1} H_i)^{-1} \\ &= M_i - M_i H_i^T (H_i M_i H_i^T + R_i^{-1})^{-1} H_i M_i \end{aligned} \quad (49)$$

$$M_{i+1} = \phi_i P_i \phi_i^T + \Gamma_i Q_i \Gamma_i^T \quad (50)$$

Equations (46), (47) and (48) are the discrete Kalman filter with the state variable covariances given by Equations (49) and (50). As can be seen from the above equations, the Kalman filter is essentially the same as the system model [Equation (36)]. The differences are (1) the actual system noise, which is random from one stage to the next in Equation (36), is replaced by its mean or expected value, and (2) there is a correction term based on the difference between the actual measurement of the state and its predicted value. The difference term is multiplied by a gain, K_i , which is essentially the ratio of the uncertainty in the state to the uncertainty in the measurement. If the covariances of the measurement errors are large, the gain will be relatively small and the corresponding difference term will have little effect on the estimate of the state; if, however, the system noise is relatively large or the measurement errors are very small, the gain will be large and differences between the actual and predicted measurements of the state at a given stage have increased significance in the state variable estimates.

d. The Continuous Kalman Filter

By applying a limiting process to the discrete filter with the time between stages tending to zero, the linear system model becomes

$$\dot{X} = F(t)X + G(t)w(t) \quad , \quad (51)$$

and the continuous Kalman filter is given by

$$\dot{\hat{X}} = F\hat{X} + G\bar{w} + PH^TR^{-1}(z - H\hat{X}) \quad (52)$$

$$\dot{P} = FP + PF^T + GG^T - PH^TR^{-1}HP \quad (53)$$

3. The Extended Kalman Filter

The extension of the Kalman filter for estimating the states of nonlinear systems in the presence of noise has been given by Eryson and Ho²⁰ and Jazwinski,³¹ among others, as

$$\dot{\hat{X}} = f(\hat{X}, t) + G(t)\bar{w}(t) + P \left[\frac{\partial h}{\partial X} \right]^T R^{-1} [z(t) - h(\hat{X}, t)] \quad (54)$$

$$\dot{P} = \frac{\partial f}{\partial X} P + P \left[\frac{\partial f}{\partial X} \right]^T + GQG^T - P \left[\frac{\partial h}{\partial X} \right]^T R^{-1} \left[\frac{\partial h}{\partial X} \right] P \quad (55)$$

for the nonlinear system

$$\dot{X} = f(X, t) + G(t)w(t) \quad (56)$$

with measurements

$$z(t) = h(X, t) + v(t) \quad (57)$$

where

$$E[w(t)] = \bar{w}(t) \quad (58)$$

$$E[w(t) - \bar{w}(t)][w(t') - \bar{w}(t')] = Q(t) \delta(t - t') \quad (59)$$

$$E[X(t_0)] = \bar{X}_0 \quad (60)$$

$$E[X(t_0) - \bar{X}_0][X(t_0) - \bar{X}_0]^T = P_0 \quad (61)$$

$$E[X(t_0) - \bar{X}_0][w(t) - \bar{w}(t)]^T = 0 \quad (62)$$

$$E[v(t)] = 0 \quad (63)$$

$$E[v(t) v^T(t')] = R(t) \delta(t - t') \quad (64)$$

$$E[X(t_0) - \bar{X}_0][v(t)]^T = 0 \quad (65)$$

$$E[w(t) - \bar{w}(t)][v(t')]^T = 0 \quad (66)$$

4. The Parameter Estimation Algorithm

In this subsection a specific parameter and state estimation algorithm using the extended Kalman filter is developed for the basic equation of pitching motion previously analyzed with the Chapman-Kirk technique.

The nonlinear equation of motion, or system model, is thus

$$\ddot{\alpha} + (C_4 + C_6 \alpha^2) \dot{\alpha} + (C_3 + C_5 \alpha^2) \alpha = w(t) \quad (67)$$

where

$$E[w(t)] = 0, \quad E[w(t)w(t')] = q\delta(t-t') \quad (68)$$

The measurements of the state of the system are assumed to be given by

$$z(t) = \alpha(t) + v(t) \quad (69)$$

where

$$E[v(t)] = 0, \quad E[v(t)v^T(t')] = r\delta(t-t') \quad (70)$$

Now to reduce Equation (67) to the required system of first order differential equations, let

$$\alpha = X_1$$

$$\dot{\alpha} = X_2 \quad (71)$$

At this point, to estimate the aerodynamic parameters appearing in the equation of motion in addition to the state variables, the state vector is augmented by setting

$$C_3 = X_3$$

$$C_4 = X_4$$

$$C_5 = X_5$$

$$C_6 = X_6 \quad (72)$$

with the constraints

$$\dot{X}_3 = 0$$

$$\dot{X}_4 = 0$$

$$\dot{X}_5 = 0$$

$$\dot{X}_6 = 0 \quad (73)$$

Making use of Equations (71), (72), and (73), Equation (67) now reduces to the following nonlinear system of first order equations:

$$\dot{X}_1 = X_2$$

$$\dot{X}_2 = -(X_4 + X_6 X_1^2) X_2 - (X_3 + X_5 X_1^2) X_1 + w(t)$$

$$\dot{X}_3 = 0$$

$$\dot{X}_4 = 0$$

$$\dot{X}_5 = 0$$

$$\dot{X}_6 = 0 \quad , \quad (74)$$

with linear measurements

$$z = X_1 + v \quad . \quad (75)$$

Comparing Equations (74) and (75) with Equations (56) and (57), the following matrices may be identified:

$$G(t) = \begin{bmatrix} 0 \\ 1 \\ 0 \\ 0 \\ 0 \\ 0 \\ 0 \end{bmatrix} \quad H(t) = \begin{bmatrix} X_1 \\ 0 \\ 0 \\ 0 \\ 0 \\ 0 \\ 0 \end{bmatrix}$$

$$f(X, t) = \begin{bmatrix} X_2 \\ -(X_4 + X_6 X_1^2) X_2 - (X_3 + X_5 X_1^2) X_1 \\ 0 \\ 0 \\ 0 \\ 0 \end{bmatrix} \quad . \quad (76)$$

Now, since the mean value of the system noise is assumed to be zero [Equation (68)] and because the measurements of the state are related linearly to the state [Equation (75)], the extended Kalman filter, previously given by Equations (54) and (55), can be simplified somewhat to

$$\dot{\hat{X}} = (\hat{X}, t) + PH^T \begin{bmatrix} 1 \\ r \end{bmatrix} [z - H\hat{X}] \quad (77)$$

$$P = \frac{\partial f}{\partial X} P + P \left[\frac{\partial f}{\partial X} \right]^T + GqG^T - PH^T \begin{bmatrix} 1 \\ r \end{bmatrix} HP \quad , \quad (78)$$

where

$$H = [1 \ 0 \ 0 \ 0 \ 0 \ 0]$$

Applying Equations (76) and (77) yields

$$\begin{bmatrix} \dot{\hat{X}}_1 \\ \dot{\hat{X}}_2 \\ \dot{\hat{X}}_3 \\ \dot{\hat{X}}_4 \\ \dot{X}_5 \\ \dot{\hat{X}}_6 \end{bmatrix} = \begin{bmatrix} X_2 \\ -(X_4 + X_6 X_1^2) X_2 - (X_3 + X_5 X_1^2) X_1 \\ 0 \\ 0 \\ 0 \\ 0 \end{bmatrix} + \begin{bmatrix} P_{11} & P_{12} & P_{13} & P_{14} & P_{15} & P_{16} \\ P_{12} & P_{22} & P_{23} & P_{24} & P_{25} & P_{26} \\ P_{13} & P_{23} & P_{33} & P_{34} & P_{35} & P_{36} \\ P_{14} & P_{24} & P_{34} & P_{44} & P_{45} & P_{46} \\ P_{15} & P_{25} & P_{35} & P_{45} & P_{55} & P_{56} \\ P_{16} & P_{26} & P_{36} & P_{46} & P_{56} & P_{66} \end{bmatrix} \begin{bmatrix} 1 \\ 0 \\ 0 \\ 0 \\ 0 \\ 0 \end{bmatrix} + \frac{1}{r} (z - [1 \ 0 \ 0 \ 0 \ 0 \ 0]) \begin{bmatrix} \hat{X}_1 \\ \hat{X}_2 \\ \hat{X}_3 \\ \hat{X}_4 \\ \hat{X}_5 \\ \hat{X}_6 \end{bmatrix} .$$

Carrying out the prescribed multiplications and equating like components of the resulting two column matrices yields the desired filter equations,

$$\dot{\hat{X}}_1 = \hat{X}_2 + \frac{P_{11}}{r} (z - \hat{X}_1)$$

$$\dot{\hat{X}}_2 = -(\hat{X}_4 + \hat{X}_6 \hat{X}_1^2) \hat{X}_2 - (\hat{X}_3 + \hat{X}_5 \hat{X}_1^2) \hat{X}_1 + \frac{P_{12}}{r} (z - \hat{X}_1)$$

$$\dot{\hat{X}}_3 = \frac{P_{13}}{r} (z - \hat{X}_1)$$

$$\dot{\hat{X}}_4 = \frac{P_{14}}{r} (z - \hat{X}_1)$$

$$\dot{\hat{X}}_5 = \frac{P_{15}}{r} (z - \hat{X}_1)$$

$$\dot{\hat{X}}_6 = \frac{P_{16}}{r} (z - \hat{X}_1) \quad (79)$$

The necessary covariance equations are obtained from the matrix Ricatti equation [Equation (78)] making use of Equations (76) and the fact that

$$\left[\frac{\partial f}{\partial \hat{X}} \right]_{\hat{X}=\hat{X}} = \begin{bmatrix} 0 & 1 & 0 & 0 & 0 & 0 \\ -2\hat{X}_1\hat{X}_2\hat{X}_6 - \hat{X}_3 - 3\hat{X}_1^2\hat{X}_5 & -(\hat{X}_4 + \hat{X}_6\hat{X}_1^2) & -\hat{X}_1 & -\hat{X}_2 & -\hat{X}_1^3 & -\hat{X}_1^2\hat{X}_2 \\ 0 & 0 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 & 0 \end{bmatrix}$$

The results of substituting Equations (76) and (80) into (78), carrying out the rather tedious matrix multiplications, and equating like terms are

$$\dot{P}_{11} = 2P_{12} - \frac{1}{r}P_{11}^2$$

$$\dot{P}_{12} = P_{22} - AP_{11} - BP_{12} - \hat{X}_1P_{13} - \hat{X}_2P_{14} - CP_{15} - DP_{16} - \frac{1}{r}P_{11}P_{12}$$

$$\dot{P}_{13} = P_{23} - \frac{1}{r}P_{11}P_{13}$$

$$\dot{P}_{14} = P_{24} - \frac{1}{r}P_{11}P_{14}$$

$$\dot{P}_{15} = P_{25} - \frac{1}{r}P_{11}P_{15}$$

$$\dot{P}_{16} = P_{26} - \frac{1}{r}P_{11}P_{16}$$

$$\dot{P}_{22} = q - 2(AP_{12} + BP_{22} + \hat{X}_1P_{23} + \hat{X}_2P_{24} + CP_{25} + DP_{26}) - \frac{1}{r}P_{12}^2$$

$$\dot{P}_{23} = -AP_{13} - BP_{23} - \hat{X}_1P_{33} - \hat{X}_2P_{34} - CP_{35} - DP_{36} - \frac{1}{r}P_{12}P_{13}$$

$$\dot{P}_{24} = -AP_{14} - BP_{24} - \hat{X}_1P_{34} - \hat{X}_2P_{44} - CP_{45} - DP_{46} - \frac{1}{r}P_{12}P_{14}$$

$$\dot{P}_{25} = -AP_{15} - BP_{25} - \hat{X}_1P_{35} - \hat{X}_2P_{45} - CP_{55} - DP_{56} - \frac{1}{r}P_{12}P_{15}$$

$$\dot{P}_{26} = -AP_{16} - BP_{26} - \hat{X}_1P_{36} - \hat{X}_2P_{46} - CP_{56} - DP_{66} - \frac{1}{r}P_{12}P_{16}$$

$$\dot{P}_{33} = -\frac{1}{r}P_{13}^2$$

$$\dot{P}_{34} = -\frac{1}{r}P_{13}P_{14}$$

$$\dot{P}_{35} = -\frac{1}{r}P_{13}P_{15}$$

$$\dot{P}_{36} = -\frac{1}{r}P_{13}P_{16}$$

$$\dot{P}_{44} = -\frac{1}{r}P_{14}^2$$

$$\dot{P}_{45} = -\frac{1}{r}P_{14}P_{15}$$

$$\dot{P}_{46} = -\frac{1}{r}P_{14}P_{16}$$

$$\dot{P}_{55} = -\frac{1}{r}P_{15}^2$$

$$\dot{P}_{56} = -\frac{1}{r}P_{15}P_{16}$$

$$\dot{P}_{66} = -\frac{1}{r}P_{16}^2$$

where

$$A = 2\hat{X}_1\hat{X}_2\hat{X}_6 + \hat{X}_3 + 3\hat{X}_1^2\hat{X}_5$$

$$B = \hat{X}_4 + \hat{X}_6\hat{X}_1^2$$

$$C = \hat{X}_1^3$$

$$D = \hat{X}_1^2\hat{X}_2$$

(82)

The desired estimates of the states and parameters are obtained by numerically integrating the filter equations [Equations (79)] and their covariances [Equations (81)]. Initial estimates of the states and their covariances as well as the variances of the system noise and measurement errors are assumed to be available. In the event that the data consist of discrete measurements, which is usually the case, the constant time between data points should be equal to the numerical integration step size. This is necessary because the state estimates are updated at each integration increment and, in doing this, the filter requires knowledge of the difference between measured and estimated state values.

If the filter successfully adjusts the true states and the imbedded parameters so that X_1 is an adequate match to the data, z , the time derivatives of the filter equations for the parameters become small [see Equations (79)] and the parameters reach steady-state or near steady-state conditions.

SECTION V

USE OF THE EXTENDED KALMAN FILTER FOR ESTIMATING PARAMETERS AND THEIR UNCERTAINTIES

1. Introduction

This section presents an analysis using the extended Kalman filter with parameter augmented state vector to determine aerodynamic coefficients and their uncertainties from noisy dynamic data. A description of the digital computer program used in the analysis is given first, followed by results obtained with this program from a variety of noisy data sets.

2. The Extended Kalman Filter Program

The extended Kalman filter program is written in Fortran IV for use primarily on an IBM model 360/65 digital computer. The basic function of the program is to integrate numerically the 27 first order differential equations which are given in the previous chapter and which comprise the extended Kalman filter with parameter augmented state vector. The specific functions of the main program and its subroutines, as well as the required input data, are described in the following paragraphs. A listing of the complete program is given in Appendix II.

a. Main Program Functions

The main program reads and writes the input data and calls the numerical integration subroutine, ADDUM.

b. Subroutine Functions

The names and functions of each of the three subroutines that are used in the extended Kalman filter program are given below.

(1) ADDUM

Subroutine ADDUM integrates the 27 filter and covariance equations using a fourth order Runge-Kutta method for starting, followed by a fourth order Adams-Bashforth predictor-corrector method for running. It is essentially identical to the subroutine of the same name used in the Chapman-Kirk coefficient extraction program and is described in detail in Reference 26. Subroutines XDOT and OUT are called from ADDUM.

(2) XDOT

Subroutine XDOT computes values of the time derivatives for the differential equations being integrated by ADDUM.

(3) OUT

This subroutine writes the output of the numerical integration.

c. Required Input Data

The program reads four categories of input data. These categories and the specific data in each are delineated in Appendix II. The format and units of the entries on a specific data card can be determined from the program listing and nomenclature list, which are also given in Appendix II.

3. Analysis of the Extended Kalman Filter

The use of the extended Kalman filter for estimating the parameters of interest is analyzed for linear and nonlinear systems with data containing only measurement errors as well as data containing both system noise and measurement errors. The analysis begins with a linear system and data containing only measurement errors and progresses through increasing stages of difficulty up to nonlinear systems and data which have been corrupted by both system noise and measurement errors. The previously discussed computer program and algorithm require no modifications to consider the linear case; by initially setting the parameters that are the numerical coefficients of the nonlinear terms in the equation of motion (and their variances) equal to zero, the extended filter for a nonlinear system reduces to the one required for a linear system.

All data used in the analysis were generated by the previously mentioned computer program, UFNOISE (Reference 29). The true values of the constants used in the data generation are those previously given in Table III with the exception of the initial value of the pitch angle for the linear system. This initial displacement is a more realistic 0.1745 radian for the linear cases considered.

a. Linear Systems with Measurement Errors Only

After the extended Kalman filter computer program had been constructed and checked, it became apparent that the technique would probably be best understood by considering relatively simple cases at first

and then progressing to more difficult ones as confidence in the technique was gained. To this end, the first success with the program was realized when analyzing data containing only measurement errors for a linear system.

The pitch angle data used in this part of the analysis were generated by integrating the equation of motion

$$\ddot{\alpha} + C_4 \dot{\alpha} + C_3 \alpha = 0 \quad (83)$$

and superimposing random Gaussian errors on the output. The standard deviation of the errors is 0.00588 radian, which corresponds to 3σ measurement errors of approximately 1.0 degree. The numerical integration step size used in generating the data was 0.005 second, and the pitch angle was output every integration increment. The equation was integrated for a total of 2.00 seconds.

(1) Basic Filter Performance

The results of using the extended Kalman filter to identify the correct values of the parameters of interest, X_3 and X_4 , are shown on Figures 20 and 21. These figures show time histories of the percent errors in the estimated values of the parameters that were computed by the extended filter program when analyzing the noisy data described previously. The percent error is defined by

$$\text{PERCENT ERROR } (\hat{X}_j) = \frac{(\hat{X}_j - X_j) \cdot 100}{X_j} \quad (84)$$

The input for the measurement error variance was the one previously computed by the simulation program based on the errors that were actually put into the data. The initial values of the parameter variances were computed from knowledge of the true value of the parameters and the arbitrarily selected initial values of the parameters. These sample variances are defined by

$$P_{jj}(0) = [X_j(0) - \hat{X}_j(0)]^2 \quad (85)$$

This method of initializing the parameter variances was chosen merely to generate the required initial numerical values. In actual experimentation, the initial variance values would depend on the method of selecting the initial parameter estimates and prior knowledge as to how accurately these initial parameter estimates were in relation to the true values of the parameters.

The initial values of all covariances were chosen as zero. This implies that errors in the estimates of the individual state vector components are initially uncorrelated. The initial values of the variances for the pitch angle and pitch angle rate were also chosen as zero. For wind tunnel dynamic experimentation where the model is initially held rigid at some given displacement to the wind stream, this seems to be a valid assumption.

As can be seen from Figures 20 and 21, the filter does an excellent job of identifying the parameters of interest. The X_3 parameter, which corresponds to the $C_{m\dot{\alpha}}$ term and which is initially in error by 25 percent, is identified with approximately zero percent error in less than 0.3 second. The correct identification of the damping parameter takes slightly more time but the results are of equal quality.

(2) Effects of Variations in the Initial Parameter Variances

This subsection presents some effects on the near steady-state estimates of the parameters of interest and their near steady-state variances for various values of the initial parameter variances. The parameter variances were initialized to values that were 25 percent higher than those used in the previous sub section and also to values that were 25 percent lower than the referenced values. The effects of these initializations are shown in Figures 22 through 25 and in Table VII for the linear system of interest.

Table VII is a summary of the percent errors in the parameter variances and the normalized parameter uncertainties for the above-mentioned variations in the initial parameter variances.

The same results are depicted graphically in Figures 22 and 23. The numerical results in Table VII are based on the near steady-state parameter and variance values that result after the filter has integrated for 1.5 seconds. These correspond to the last point

TABLE VII. EFFECTS OF VARIATIONS IN THE INITIAL
PARAMETER VARIANCES ON NEAR STEADY-
STATE PARAMETERS AND THEIR VARIANCES
(LINEAR SYSTEM, MEASUREMENT ERRORS ONLY)

Variation in Initial Parameter Variance* (Percent)	Errors in Near Steady-State Parameters (Percent)		Normalized Near Steady-State Variances (Percent)	
	\hat{X}_3	\hat{X}_4	$\sqrt{P_{33}}_n$	$\sqrt{P_{44}}_n$
-25	0.17	0.55	0.12	0.91
0	0.19	0.46	0.11	0.91
25	0.22	0.38	0.11	0.91
*relative to sample variance				

shown on the supporting figures. The normalized parameter uncertainties are defined by

$$\text{PERCENT } (\sqrt{P_{jj}})_n = \frac{\sqrt{P_{jj}}}{x_j} \cdot 100 \quad (86)$$

Time histories of the parameter variances, P_{33} and P_{44} , are shown in Figures 24 and 25. All of the information presented in these figures indicates that initial values of parameter variances in the range investigated have no effect on final parameter estimates or their variances when using the extended Kalman filter to analyze data for a linear system containing only measurement errors.

b. Linear Systems with Measurement Errors and System Noise

Data for this portion of the analysis were generated with the computer program UFNOISE by numerically integrating the following equation of motion

$$\ddot{\alpha} + C_4 \dot{\alpha} + C_3 \alpha = w(t) \quad (87)$$

and superimposing random Gaussian errors on the output. The forcing function, $w(t)$, is also random in nature and Gaussian with zero mean. The standard deviation of the system noise was 4.80 rad/sec², which is of sufficient magnitude to drive the oscillations in a noise band of approximately 0.0872 radian (5 degrees) for zero initial displacement. The standard deviation of the measurement errors was 0.00556 radian. Also, as before, the equation of motion was integrated in increments of 0.005 second for a total of 2.00 seconds.

(1) Basic Filter Performance

The basic performance of the filter is demonstrated in Figures 26 and 27 which show time histories of the percent error in X_3 and X_4 . As can be seen from these figures, the initial estimates of X_3 and X_4 are 25 percent in error and these estimates are

corrected to within approximately 1 percent and 3 percent of their respective true values. The noise variances used are those given by the simulation which generated the pitch angle data. The initial parameter variances are calculated from Equation (85).

(2) Effects of Variations in the Initial Parameter Variances

The sensitivity of the filter to variations in the initial parameter variances is given in Table VIII and in Figures 28 through 32. Table VIII is a summary of the percent error in the parameters of interest and their normalized uncertainties for variations in the initial parameter variances of -25 percent, 0.0 percent and 25 percent relative to the sample variances.

Figures 28 and 29 show the variation of the near steady-state errors in the parameters and their normalized uncertainties as functions of the percent variations in the initial parameter variances. Figure 30 shows the two time histories of the error in the estimate of X_4 that result for two different initial values of the parameter variances. One of the initial variances is the sample variance based on the initial and true values of the parameter [Equation (85)], and the other is 25 percent less than this value.

The X_3 error time histories are essentially identical for both cases and are not presented. The variations of the parameter variances with time, for the three initial values considered, are shown in Figures 31 and 32.

From the information given in the figures mentioned above, it is obvious that neither X_3 nor P_{33} are affected by variations in the initial variances within the ± 25 percent range considered. The estimates of the damping parameter as well as its variance are affected, however, by these variations. The near steady-state damping parameter estimates vary almost linearly with initial variance values from a low of 0.47 percent to a high of 4.95 percent (see Figure 28). The variation in the normalized uncertainty is less pronounced, ranging from a low of 14.39 percent to a high of 15.81 percent (see Figure 39). Nevertheless, the near steady-state uncertainties are of sufficient magnitude so that the true values of the coefficients are within less

TABLE VIII. EFFECTS OF VARIATIONS IN THE INITIAL
PARAMETER VARIANCES ON NEAR STEADY-
STATE PARAMETERS AND THEIR VARIANCES
(LINEAR SYSTEM, MEASUREMENT ERRORS
AND SYSTEM NOISE)

Variation in Initial Parameter Variance* (Percent)	Errors in Near Steady-State Parameters (Percent)		Normalized Near Steady-State Variances (Percent)	
	\hat{X}_3	\hat{X}_4	$\sqrt{P_{33}}_n$	$\sqrt{P_{44}}_n$
-25	0.72	-0.47	2.42	14.39
0	0.88	2.66	2.42	15.18
25	0.98	4.95	2.42	15.81
*relative to sample variance				

than one standard deviation of their estimated value regardless of the initial parameter variance, when the standard deviation is taken as the square root of the near steady-state parameter variance. Attention is also called to the fact that since system noise is now present in the data, the near steady-state parameter variances are of greater magnitude than was the case when only measurement errors were present (see Figures 24, 25, 31, and 32).

(3) Filter Response to Large Errors in the Initial Parameter Estimates

The results of using the extended Kalman filter to identify the damping parameter X_4 , when the initial estimate of the parameter is in error by 100 percent are shown in Figure 33 for two initial parameter variance estimates. When the initial parameter variance is based on the initial estimate of X_4 , the near steady-state parameter estimate is approximately 14 percent high; for an initial parameter variance that is greater by approximately 25 percent, the resulting near steady-state value of X_4 is nearly 19 percent high. The corresponding parameter variances 1.5 seconds after the filter begins integrating are such that the estimated values of X_4 are within approximately one standard deviation of the true value.

(4) Comparison of Filter Performance to Chapman-Kirk Technique

The parameters of interest and their uncertainties were extracted from the given noisy data set using the previously described technique of Chapman and Kirk in order to be compared to the filter results. The percent error in the parameters and their normalized estimated standard deviations obtained with the Chapman-Kirk program are

X_3 : 12.5 percent (1.8 percent)

X_4 : 3.1 percent (0.3 percent)

The above results are similar to those presented in Table VI in that the estimated parameter uncertainties are not of sufficient magnitude to reflect the true error in the extracted parameters. Whereas the extended filter estimates the parameters and variances

so that the true value of the parameter is generally within one standard deviation of the estimate, it is shown that such is not the case when using the Chapman-Kirk technique. The errors in the above parameters are also slightly greater than those obtained with the extended filter, although this is probably of less consequence than the differences in the uncertainties computed by the two techniques.

c. Nonlinear System with Measurement Errors Only

The pattern used in the previous two sections of generating a data set, investigating the basic response of the filter in identifying the parameters of interest, and checking the sensitivity of the filter to variations in the initial parameter variances is essentially repeated here for the more complicated system model

$$\ddot{\alpha} + (C_4 + C_6\alpha^2)\dot{\alpha} + (C_3 + C_5\alpha^2)\alpha = 0 \quad . \quad (88)$$

In addition, some effects of errors in the measurement error variance are also included and discussed. Some comparisons of results achieved with the extended filter and the Chapman-Kirk technique are also made.

The standard deviation of the measurement errors in the generated data is 0.00569 radian. The mean value of the errors is zero.

(1) Basic Filter Performance:

The time histories of the four parameters of interest as computed by the extended filter while analyzing the data just described are shown in Figures 34, 35, 36 and 37. The initial errors in the parameters are -25 percent for X_3 , X_4 , and X_6 and -5 percent for X_5 . The measurement error variance is that obtained from the simulation. The initial values of the parameter variances are based on the difference between the true values of the parameters and the above-mentioned initial estimates.

The reason the initial estimate of X_5 is only 5 percent low instead of 25 percent, as is the case with the other parameters, is due to its magnitude and a weakness in the numerical scheme. The true value of X_5 is 1960 based on the data in Table III. A 25 percent initial error in this parameter results in an initial parameter sample variance of

$$P_{55}(0) = (1960 - (0.25)1960)^2 = 240,000$$

which is of sufficient magnitude to cause the filter to diverge for the numerical integration step size of 0.005 second. The divergence occurs because the large initial value of P_{55} results in large values of P_{25} and P_{15} which, in turn, cause the values of P_{55} that are fed into the numerical integration subroutine to be very large [Equations (81) and (82)]. The result of this cascading effect is a rapid decrease in P_{55} in large increments until it becomes negative (a physical impossibility since this term is the variance of a parameter). The structure of the covariance equations is such that a negative variance causes numerical divergence in the solutions to several of the equations.

To circumvent the problem, an initial error of -10 percent in X_5 , with a corresponding adjustment in $P_{55}(0)$, was tried. The result was again divergence. Finally, an initial estimate of -5 percent was found to be successful.

The results shown in Figures 34, 35, and 36 indicate that the filter does an excellent job in correctly identifying the linear and nonlinear static restoring moment parameters, X_3 and X_5 , as well as the linear pitch damping term, X_4 . The results for the nonlinear damping parameter, X_6 , are not as gratifying as can be seen by referring to Figure 37. The error in this term is still approximately -10 percent after 1.5 seconds. The reason for this problem is related to the magnitude of this parameter as compared to others in the equation of motion. The true value of X_3 and X_5 are 160 and 1960, respectively, although the magnitude of X_5 is effectively reduced by at least an order of magnitude when it is multiplied by α^3 . X_4 and X_6 , on the other hand, have true values of 1.60 and 4.35, respectively. After X_6 is multiplied by $\alpha^2 \dot{a}$, it is effectively the smallest parameter in the equation of motion by an order of magnitude and it is naturally more difficult to identify since it has the least effect on the trajectory.

(.) Effects of Variations in the Initial Parameter Variances

The effects of variations in the initial parameter variances are summarized in Table IX and in Figures 38 and 39. As can be seen from these results, the

TABLE IX. EFFECTS OF VARIATIONS IN THE INITIAL PARAMETER VARIANCES
ON NEAR STEADY-STATE PARAMETERS AND THEIR VARIANCES (NON-
LINEAR SYSTEM, MEASUREMENT ERRORS ONLY)

Variation in Initial Parameter Variance* (Percent)	Errors in Near Steady-State Parameters (Percent)						Normalized Near Steady-State Variances (Percent)		
	\hat{x}_3	\hat{x}_4	\hat{x}_5	\hat{x}_6	$\sqrt{P_{33}}$ n		$\sqrt{P_{44}}$ n	$\sqrt{P_{55}}$ n	$\sqrt{P_{66}}$ n
-50	-0.04	1.18	-0.10	-13.4	0.44	0.44	1.15	0.35	10.4
-25	-0.10	1.02	-0.01	-12.1	0.44	0.44	1.22	0.36	10.9
0	-0.13	0.94	0.04	-11.6	0.44	0.44	1.23	0.34	10.7
25	-0.15	0.93	0.06	-11.7	0.39	0.39	1.17	0.15	8.9
* relative to sample variance									

near steady-state values of the parameters and their variances are relatively insensitive over the range of initial parameter variances considered. Only the X_6 parameter estimate and its variance show a variation of more than 1 percent. As is obvious from the referenced table and figures, no data are available for initial parameter variances of 50 percent above the sample variances; the reason for this is, again, that the magnitude of the P_{55} variance caused the filter to diverge.

Attention is called to the fact that, for the range of initial variances considered, the near steady-state estimates of the X_3 , X_4 , and X_5 parameters are all within one standard deviation of their true values when the standard deviation is taken as the square root of the near steady-state variance for the parameter of interest. The X_6 parameter is always within two standard deviations of its true value.

(3) Effects of Errors in the Estimate of the Measurement Error Variance

Up to this point, exact values of the measurement error variance, r , have been used in all runs made with the extended filter. Exact knowledge of this quantity is available because the pitch-angle data being used with the extended filter are computer-generated with specified noise statistics. In actual test situations, true values of the noise parameters may not be known exactly and, in fact, methods for determining both measurement error variances and system noise variances from noisy dynamic data are of current research interest.³⁴

Some effects of errors in the estimates of the measurement error variance on the near steady-state parameter estimates and their uncertainties are given in Table X and in Figures 40 and 41. These results show that errors in the measurement noise variance of ± 25 percent have essentially no effect on estimates of the four parameters of interest. The uncertainties of the X_3 , X_4 , and X_5 parameters are changed by less than 1 percent over the range considered. The normalized X_6 uncertainty varies from approximately 5 percent to 12 percent.

TABLE X. EFFECTS OF ERRORS IN THE ESTIMATE OF THE MEASUREMENT ERROR VARIANCE ON NEAR STEADY-STATE PARAMETERS AND THEIR VARIANCES (NONLINEAR SYSTEM, MEASUREMENT ERRORS ONLY)

Error in Estimate of Measurement Error Variance, r (Percent)	Errors in Near Steady-State Parameters (Percent)				Normalized Near Steady-State Variances (Percent)			
	\hat{x}_3	\hat{x}_4	\hat{x}_5	\hat{x}_6	$\frac{\sqrt{P_{33}}}{n}$	$\frac{\sqrt{P_{44}}}{n}$	$\frac{\sqrt{P_{55}}}{n}$	$\frac{\sqrt{P_{66}}}{n}$
-25	-0.14	0.95	0.05	-12.2	0.27	0.91	0.0	4.93
0	-0.13	0.94	0.04	-11.6	0.44	1.23	0.34	10.73
25	-0.10	1.00	0.00	-11.9	0.50	1.37	0.40	12.19

(4) Filter Response to Larger Error in the Initial Estimate of X_5

The difficulty associated with numerically large initial parameter variance estimates has been discussed earlier. The sample variance of P_{55} which corresponds to a -5 percent error in the initial X_5 parameter estimate appears to be an approximate upper limit for the integration step size and physical constants used in this analysis. A run was made, however, with the previously described nonlinear data set containing only measurement errors where all initial parameter estimates were in error by -25 percent; the P_{55} variance, however, was set equal to the previously used sample variance which corresponds to a -5 percent initial estimate of X_5 . The parameter estimate time histories resulting from this initialization are shown in Figures 42 through 45. These results indicate that the near steady-state values of the parameters are very nearly equal to those presented when the sample variances based on the initial parameter estimates were used initially.

(5) Comparison of Filter Performance to Chapman-Kirk Technique

The linear and nonlinear static restoring moment and damping parameters were extracted from the noisy data set being considered in this section using the Chapman-Kirk program. The percent errors in the parameters and their normalized estimated uncertainties are given below.

X_3 : 0.36 percent (0.35 percent)

X_4 : 0.69 percent (0.43 percent)

X_5 : 0.41 percent (0.49 percent)

X_6 : 4.60 percent (11.27 percent)

Comparing the above results with those previously given in Table IX reveals that the estimates obtained using the extended filter are slightly more accurate for all but the X_4 parameter. Since the errors are, for the most part, less than 1 percent, the differences are essentially insignificant. The normalized uncertainties obtained with both techniques are equally good.

d. Nonlinear System with Measurement Errors and System Noise

In this section, as before, the basic filter performance, sensitivity of the filter to initial parameter variance variations and sensitivity of the filter to errors in the estimates of the noise variances are investigated. The data used in the analysis were generated by numerically integrating the following nonlinear equation with a random forcing function,

$$\ddot{\alpha} + (C_4 + C_6 \alpha^2) \dot{\alpha} + (C_3 + C_5 \alpha^2) \alpha = w(t) \quad . \quad (89)$$

Approximately the same standard deviations used previously are repeated once again; the standard deviation of the measurement errors is 0.00576 radian and the standard deviation of the system noise is 4.75 rad/sec².

(1) Basic Filter Performance

The time histories of the errors in the parameters of interest obtained with the extended filter are given in Figures 46 through 49. The response time required for the parameters to reach their near steady-state values is slightly greater than in the simpler cases that have been considered previously. The accuracy of the parameters is similar to that which was achieved for the nonlinear system with only measurement errors in the data with the exception of X_4 , which is approximately 5 percent higher.

(2) Effects of Variations in the Initial Parameter Variances

The effects of variations in the initial parameter variances are given in Table XI and are shown in Figures 50 and 51. Only the nonlinear damping parameter and its uncertainty vary by more than 1 percent for initial variance variations of up to 50 percent relative to the sample variances.

(3) Effects of Errors in the Estimates of the Noise Variances

The effects of errors in the estimates of r , the measurement error variance, and q , the system noise variance, are given in Table XII and are shown in

Figures 52 and 53. These figures show the variation of the near steady-state parameter errors and normalized parameter uncertainties as functions of the error in the estimate of the measurement error variance for various values of error in the system noise variance estimate. From these results it is obvious that only X_6 is affected by errors of ± 25 percent in the estimate of the system noise variance or measurement error variance.

(4) Comparison of Filter Performance to Chapman-Kirk Technique

The percent errors in the parameters and their normalized estimated standard deviations that are obtained using the Chapman-Kirk program with the data currently being analyzed are given below:

X_3 : 0.62 percent (0.40 percent)

X_4 : 2.63 percent (1.23 percent)

X_5 : 0.41 percent (0.61 percent)

X_6 : 46.3 percent (14.8 percent)

Comparing these results with the results in Table XI indicates that both techniques yield essentially the same accuracy in their determination of X_3 and X_5 . The filter does considerably better in estimating X_6 but is slightly worse in its estimate of X_4 .

The uncertainties computed using the filter are of sufficient magnitude so that the true value of the X_3 , X_5 , and X_6 parameters are within one standard deviation of the filter estimates; the true value of X_4 is within two standard deviations.

When using the Chapman-Kirk technique, only X_5 is within one estimated standard deviation of its true value; X_3 and X_4 are within two estimated standard deviations of their respective true values. The extracted value of X_6 is not within three estimated standard deviations of its true value.

TABLE XI. EFFECTS OF VARIATIONS IN THE INITIAL PARAMETER VARIANCES ON NEAR STEADY-STATE PARAMETERS AND THEIR VARIANCES (NON-LINEAR SYSTEM, MEASUREMENT ERRORS AND SYSTEM NOISE)

Variation in Initial Parameter Variance* (Percent)	Errors in Near Steady-State Parameters (Percent)				Normalized Near Steady-State Variances (Percent)			
	\hat{X}_3	\hat{X}_4	\hat{X}_5	\hat{X}_6	$\sqrt{P_{33}}$ n	$\sqrt{P_{44}}$ n	$\sqrt{P_{55}}$ n	$\sqrt{P_{66}}$ n
-50	-0.36	6.38	0.71	-16.9	1.74	4.27	1.16	16.7
-25	-0.57	6.51	0.87	-14.0	1.76	4.52	1.18	20.3
0	-0.59	6.40	0.92	-11.4	1.77	4.71	1.19	23.2
25	-0.58	6.17	0.92	-8.8	1.77	4.87	1.17	25.5
*relative to sample variance								

TABLE 4.1. EFFECTS OF ERRORS IN THE ESTIMATES OF NOISE VARIANCES ON NEAR STEADY-STATE PARAMETERS AND THEIR VARIANCES (NONLINEAR SYSTEM, MEASUREMENT ERRORS AND SYSTEM NOISE)

Errors in Estimates of Noise Variances (Percent)		Errors in Near Steady-State Parameters (Percent)					Normalized Near Steady-State Variances (Percent)			
q	r	\hat{X}_3	\hat{X}_4	\hat{X}_5	\hat{X}_6		$\sqrt{P_{33}}$ n	$\sqrt{P_{44}}$ n	$\sqrt{P_{55}}$ n	$\sqrt{P_{66}}$ n
-25	-25	-0.56	6.08	.90	-7.8		1.52	4.26	1.01	22.6
-25	0	-0.67	6.18	0.98	-8.4		1.57	4.29	1.07	22.9
0	-25	-0.50	6.38	0.88	-11.4		1.72	4.69	1.13	23.0
0	0	-0.59	6.40	0.92	-11.4		1.80	4.71	1.19	23.2
0	25	-0.62	6.43	0.95	-11.6		1.80	4.72	1.22	23.3
25	0	-0.52	6.51	0.87	-13.5		1.94	5.08	1.29	23.3
25	25	-0.54	6.51	0.89	-13.5		1.97	5.10	1.33	23.4

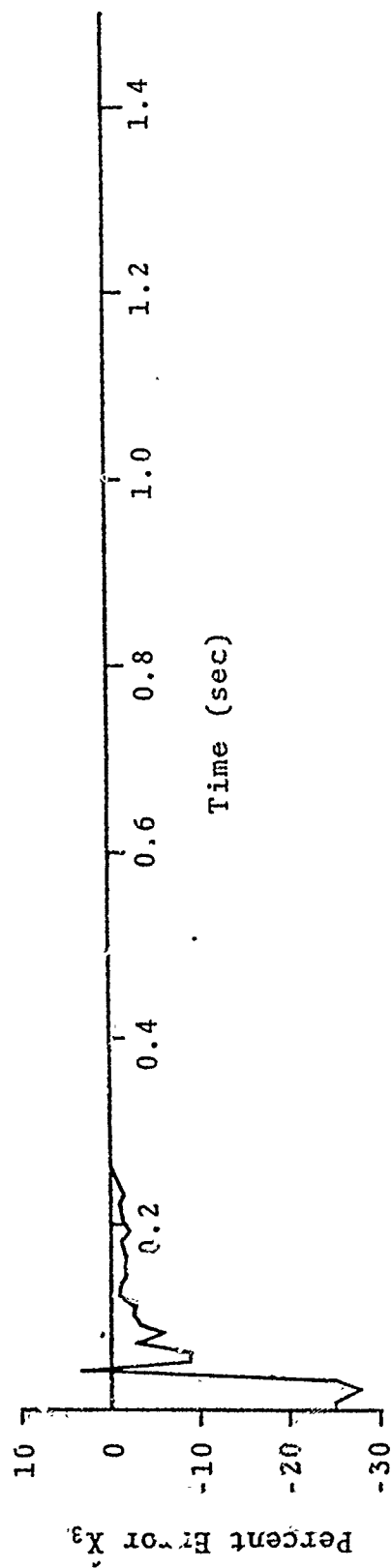


Figure 20. Variation of Percent Error in \hat{X}_3 with Time
(Linear System, Measurement Errors Only)

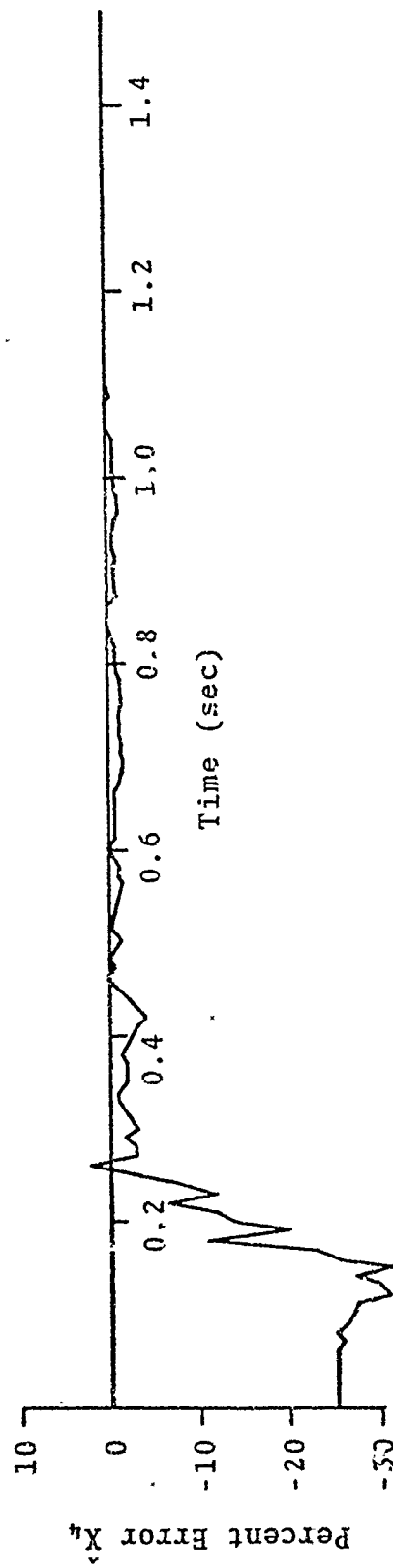


Figure 21. Variation of Percent Error in \hat{X}_4 with Time
(Linear System, Measurement Errors Only)

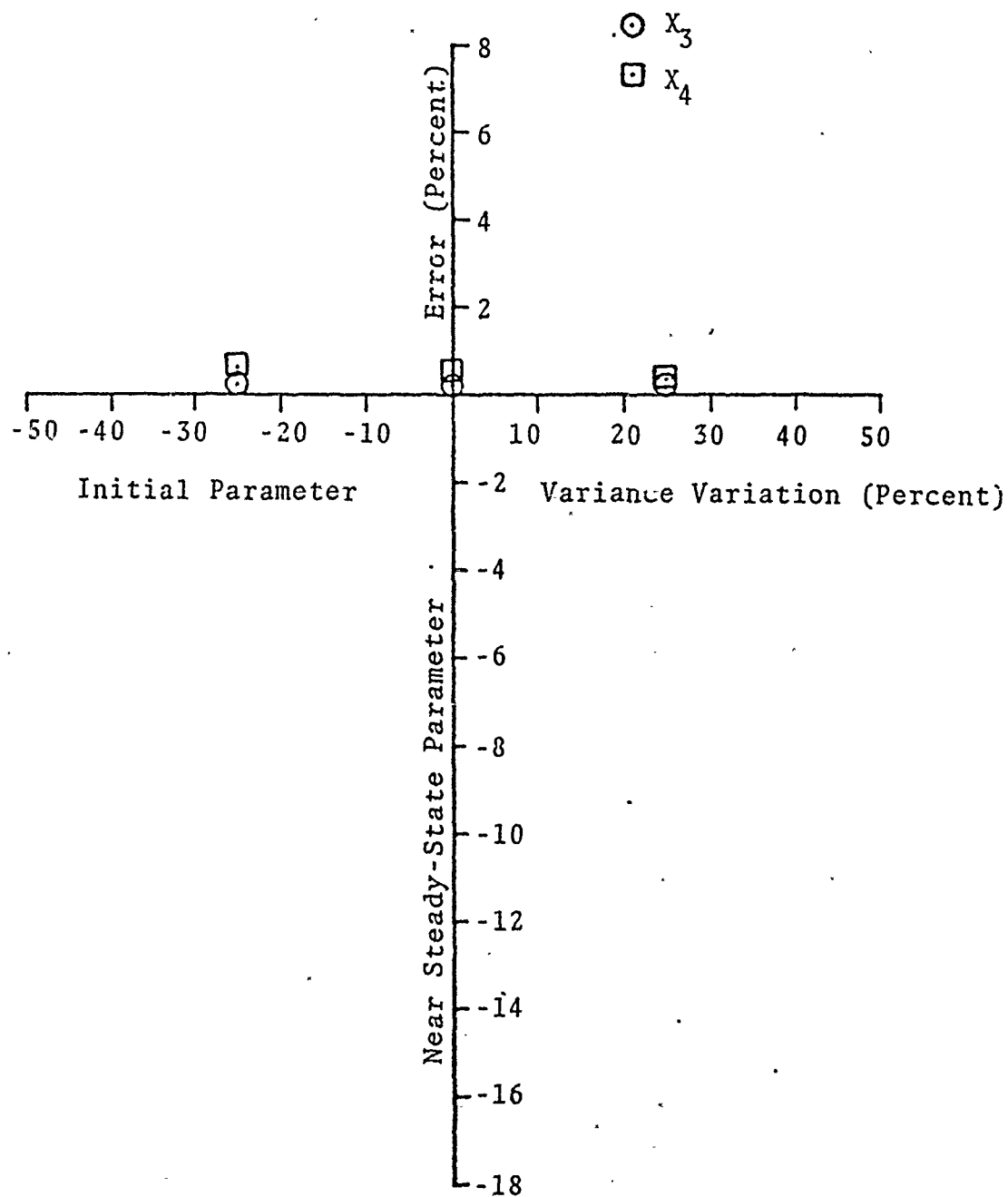


Figure 22. Variation of Near Steady-State Parameter Error with Initial Parameter Variance (Linear System, Measurement Errors Only)

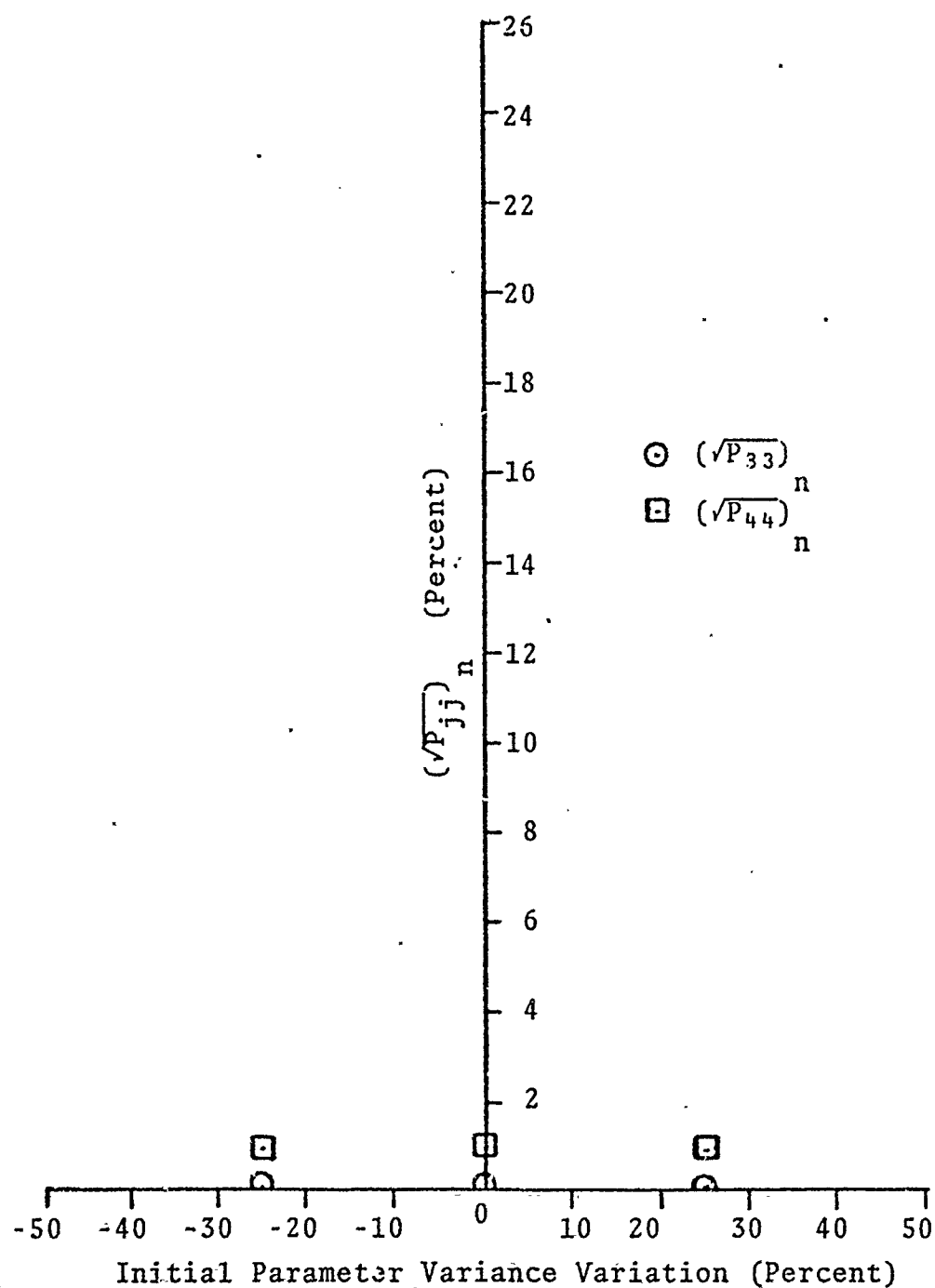


Figure 23. Variation of Normalized Near Steady-State Parameter Uncertainty with Initial Parameter Variance (Linear System Measurement Errors Only)

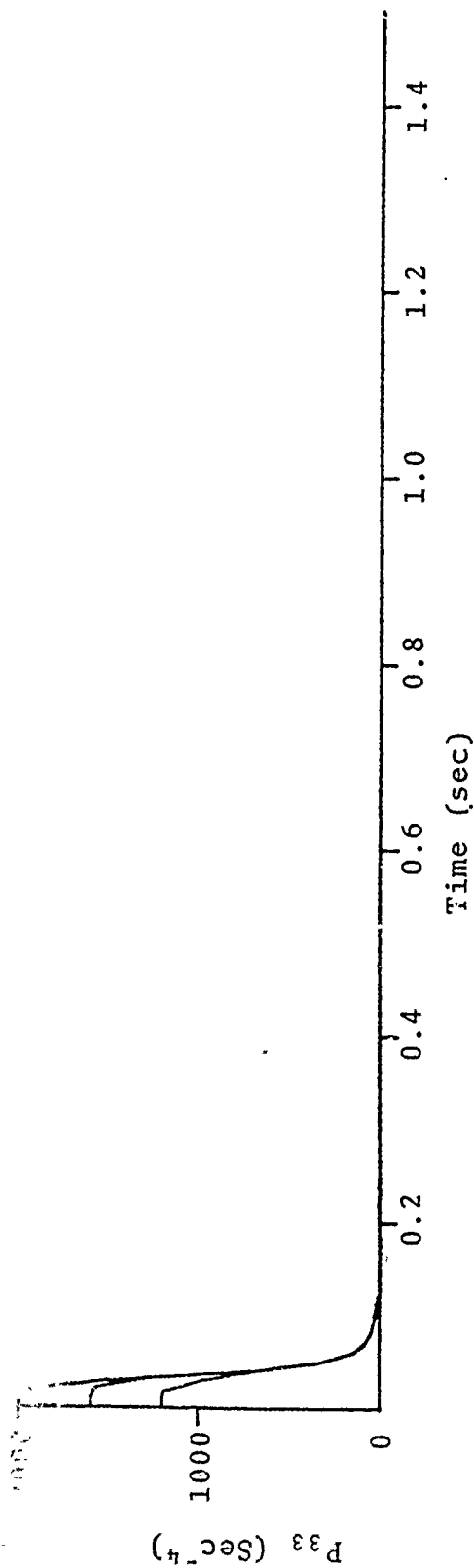


Figure 24. Variation of \hat{X}_3 Variance with Time for Three Initial Estimates (Linear System, Measurement Errors Only)

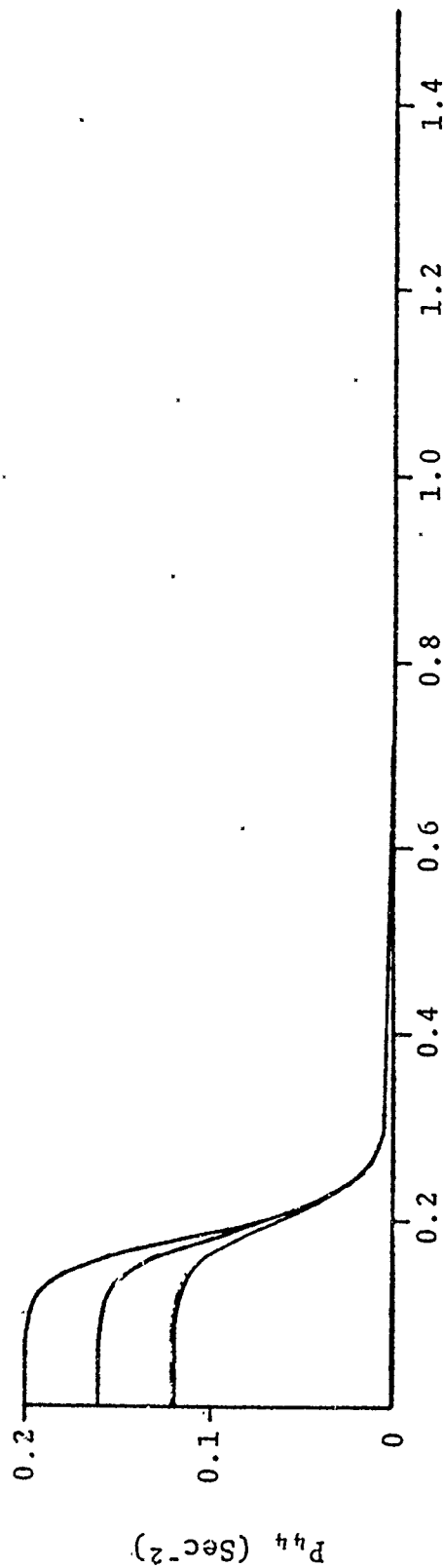


Figure 25. Variation of \hat{X}_4 Variance with Time for Three Initial Estimates (Linear System, Measurement Errors Only)

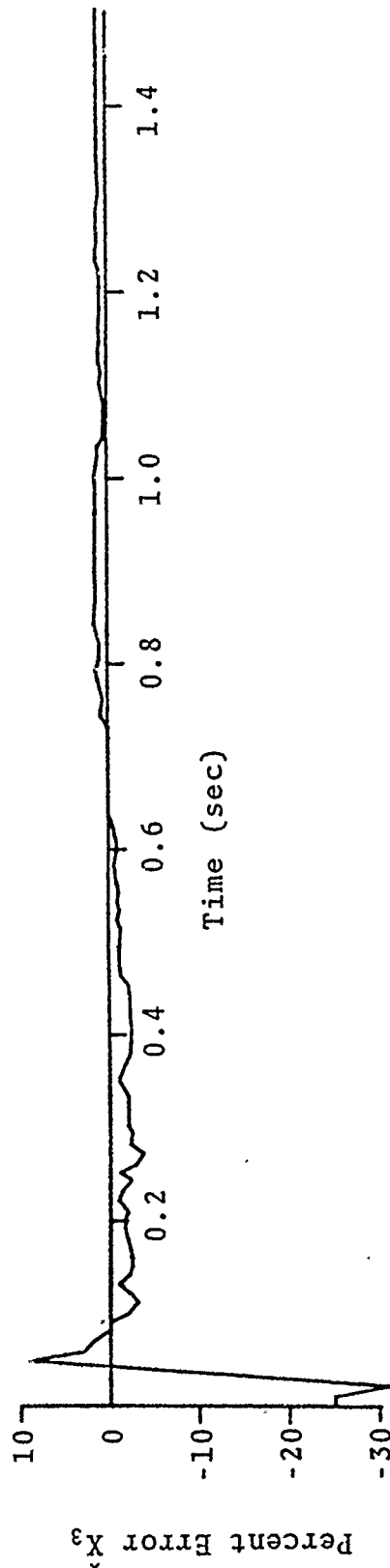


Figure 26. Variation of Percent Error in \hat{X}_3 with Time (Linear System, Measurement Errors and System Noise)

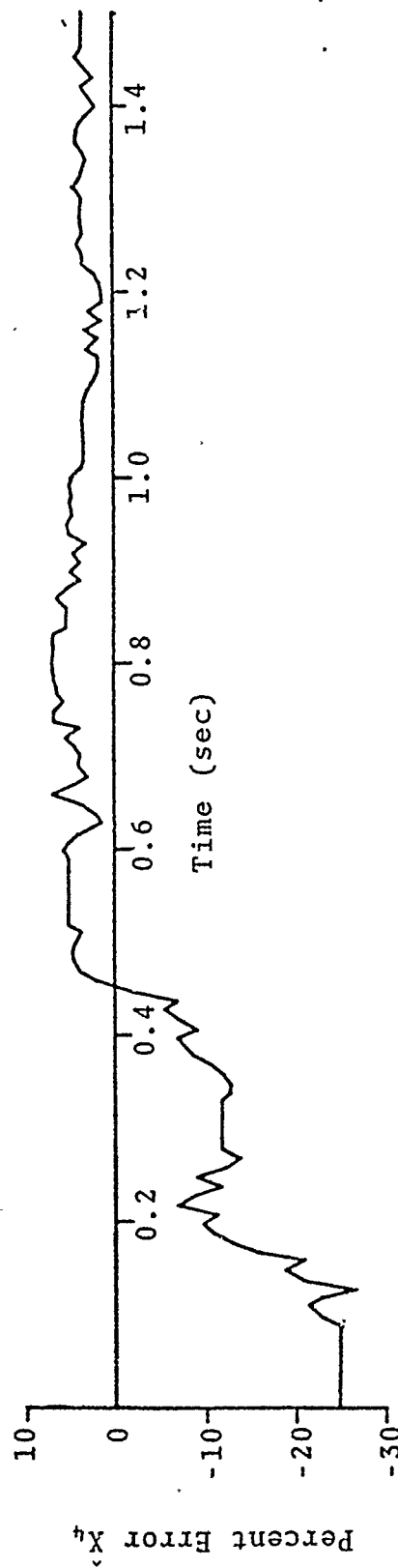


Figure 27. Variation of Percent Error in \hat{X}_4 with Time (Linear System, Measurement Errors and System Noise)

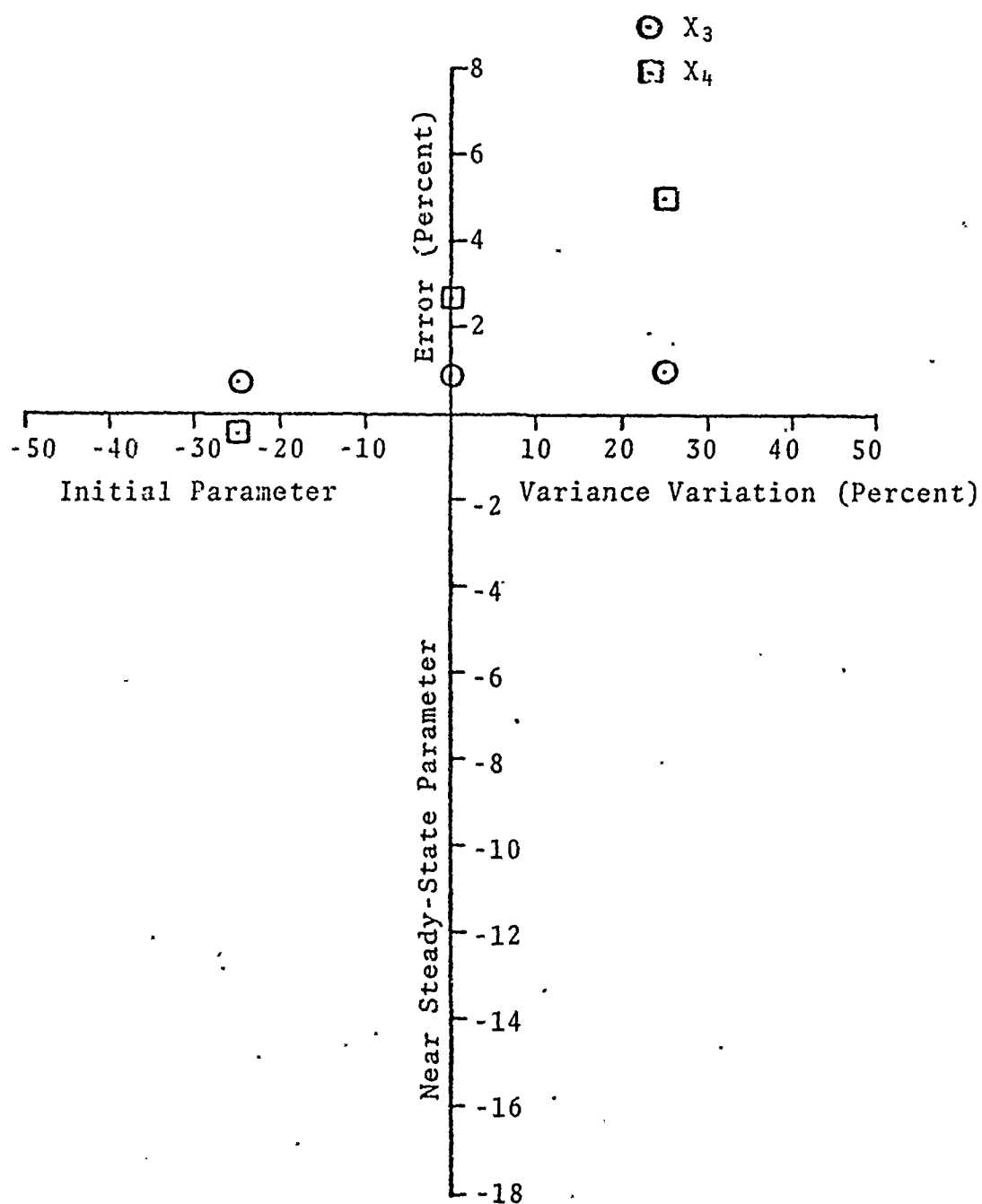


Figure 28. Variation of Near Steady-State Parameter Error with Initial Parameter Variance (Linear System, Measurement Errors and System Noise)

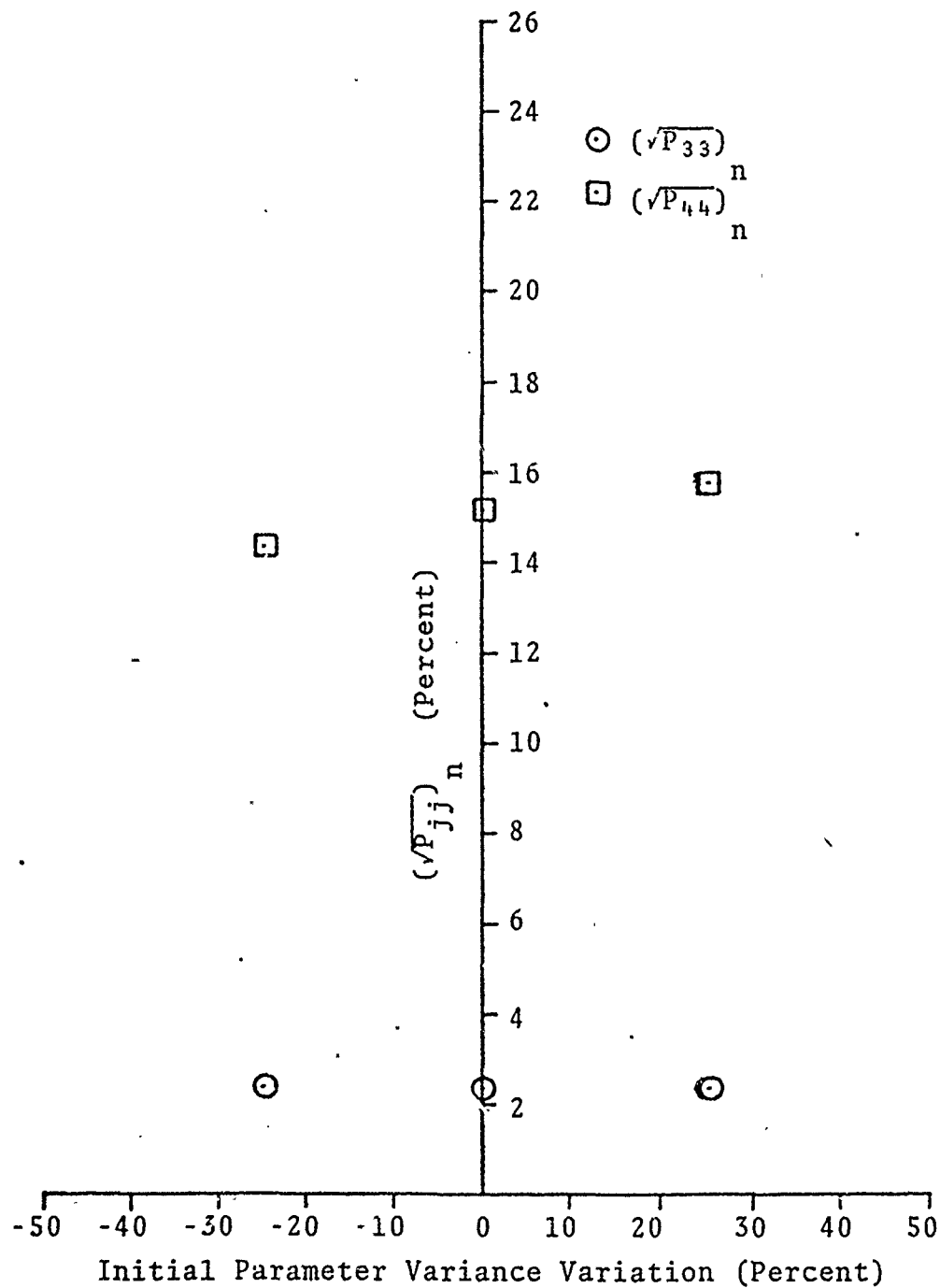


Figure 29. Variation of Normalized Near Steady-State Parameter Uncertainty with Initial Parameter Variance (Linear System, Measurement Errors and System Noise)

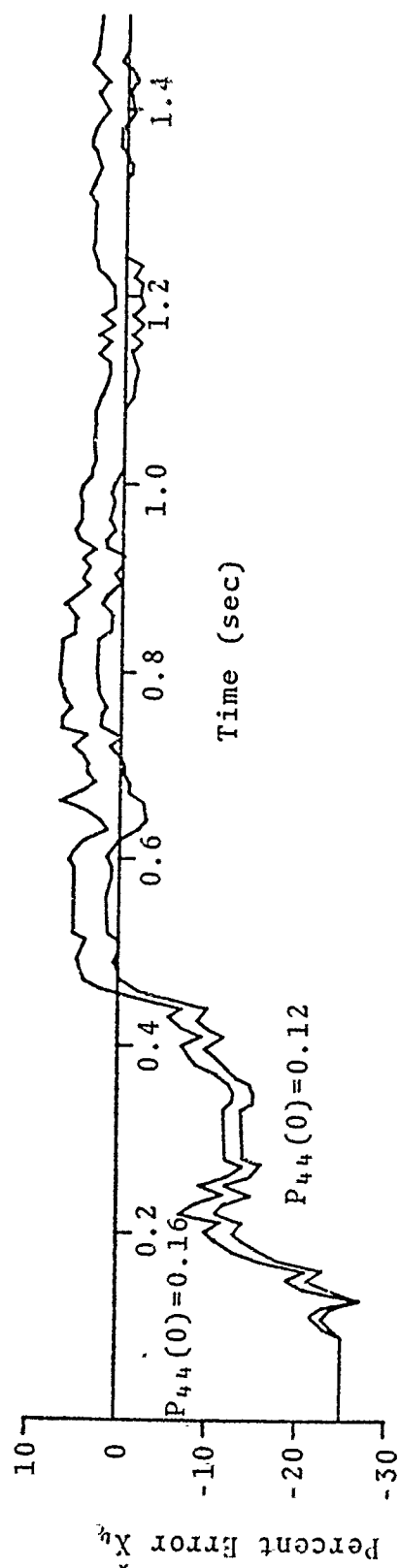


Figure 30. Variation of Percent Error in \hat{X}_4 with Time for Two Estimates of the Initial Parameter Variance (Linear System, Measurement Errors and System Noise)

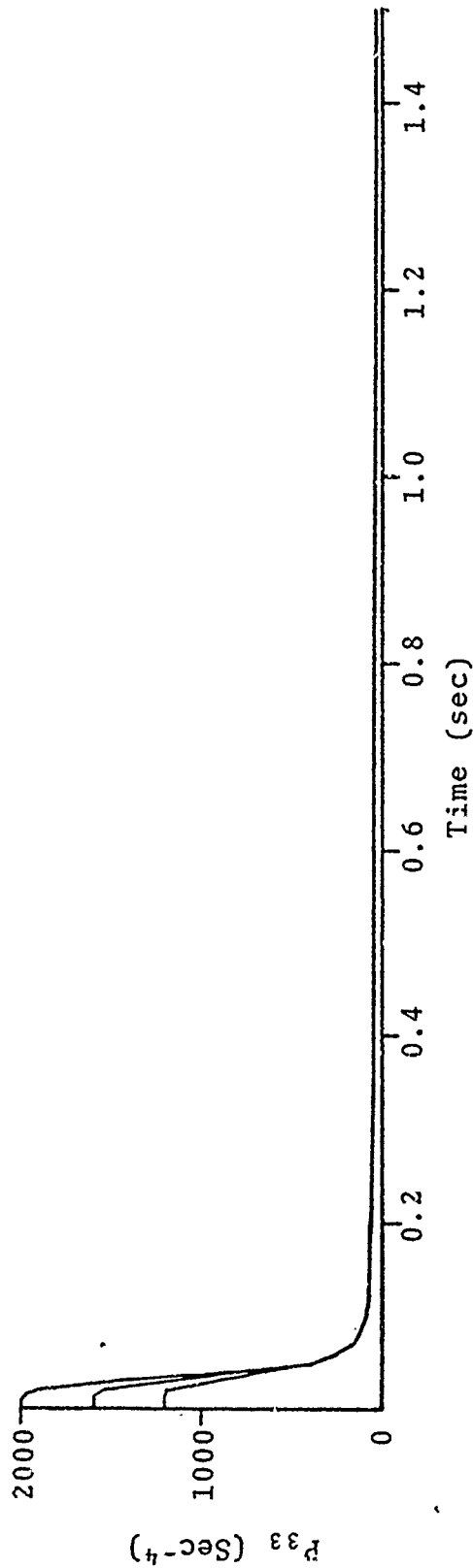


Figure 31. Variation of \hat{X}_3 Variance with Time for Three Initial Estimates
(Linear System, Measurement Errors and System Noise)

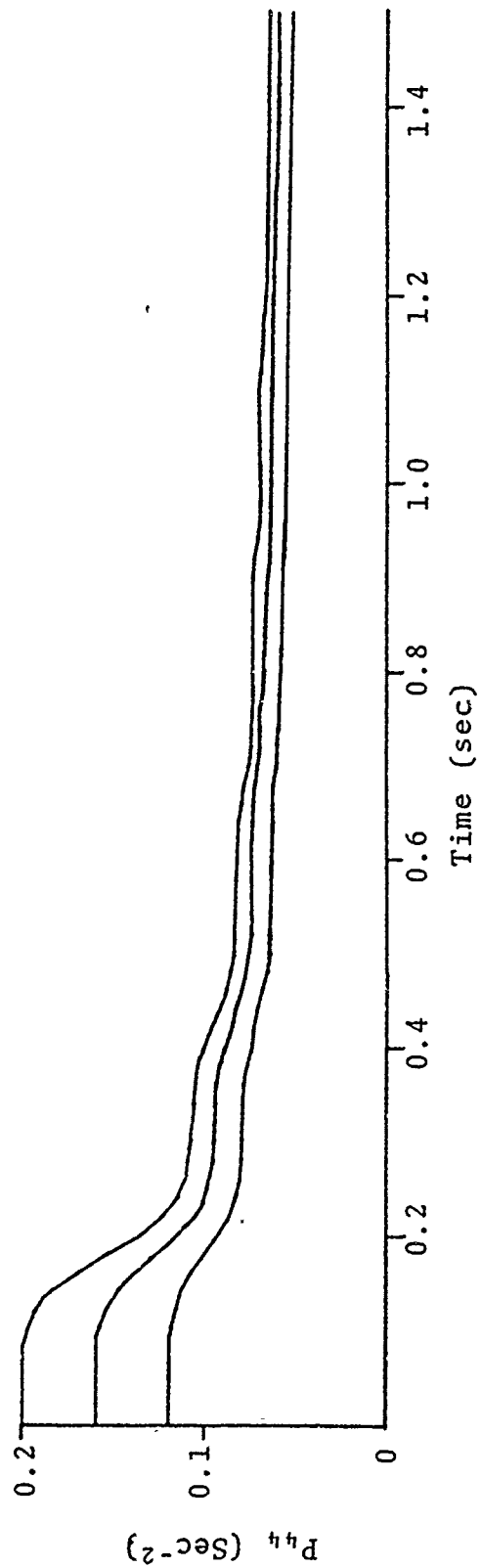


Figure 32. Variation of \hat{X}_4 Variance with Time for Three Initial Estimates
(Linear System, Measurement Errors and System Noise)

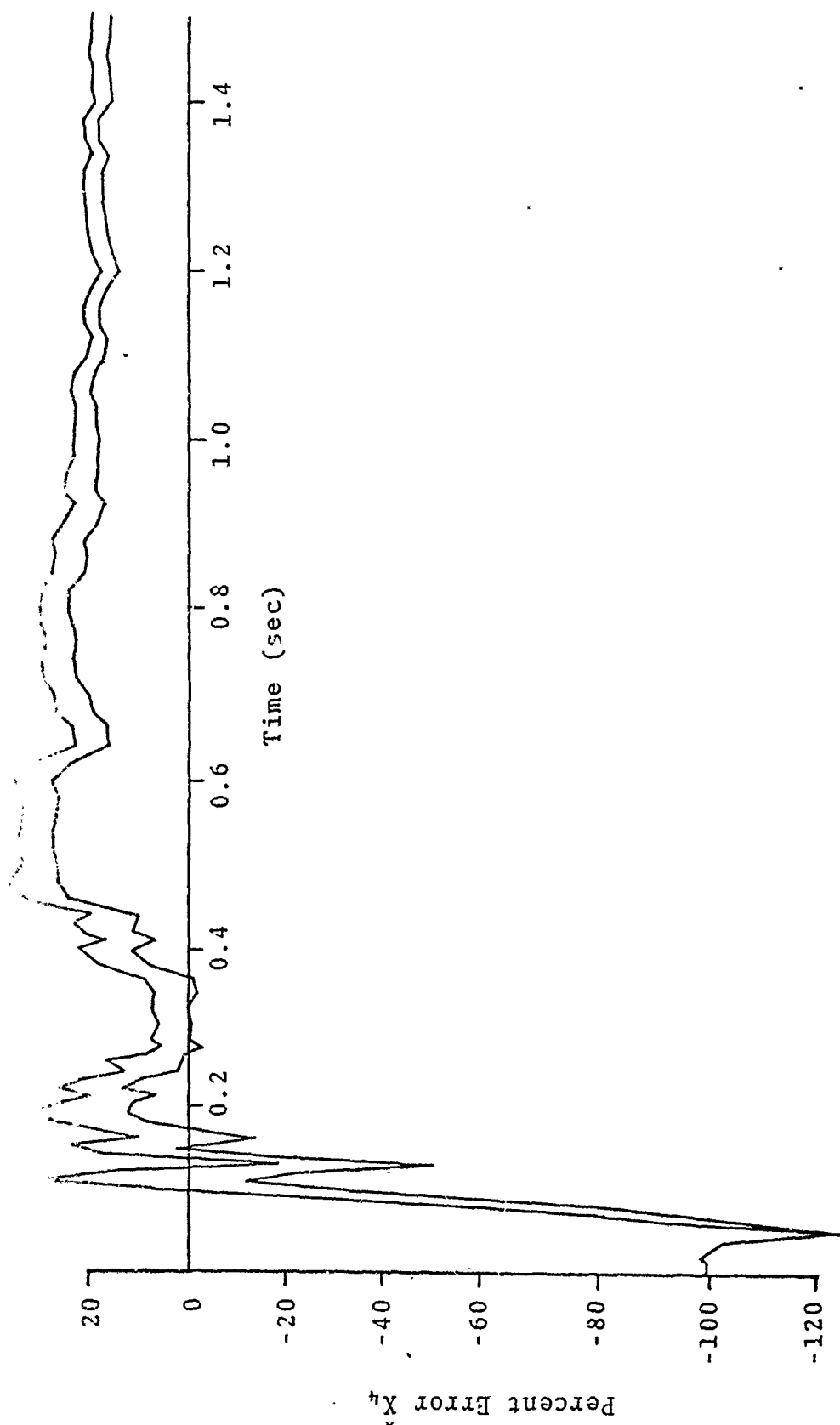


Figure 33. Variation of Percent Error in \hat{X}_4 with Time for Large Initial Parameter Estimate Error and Two Initial Parameter Variance Estimates (Linear System, Measurement Errors and System Noise)

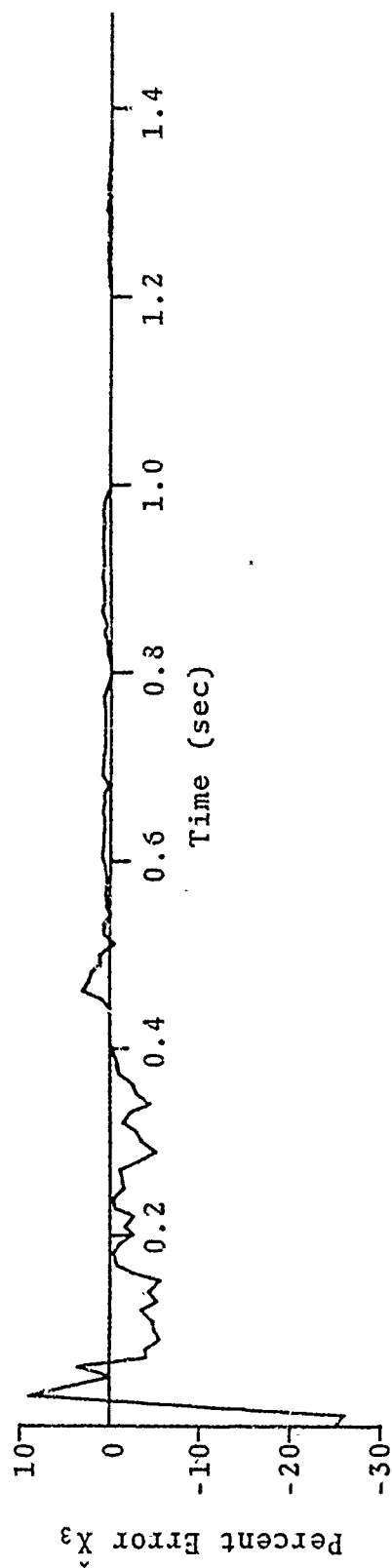


Figure 34. Variation of Percent Error in \hat{X}_3 with Time (Nonlinear System, Measurement Errors Only)

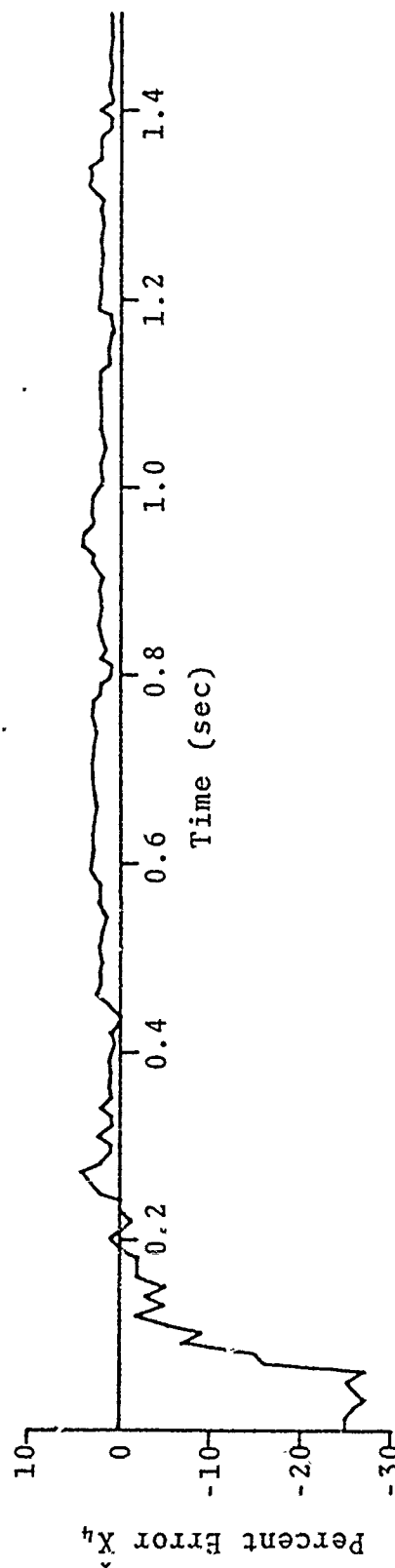


Figure 35. Variation of Percent Error in \hat{X}_4 with Time (Nonlinear System, Measurement Errors Only)

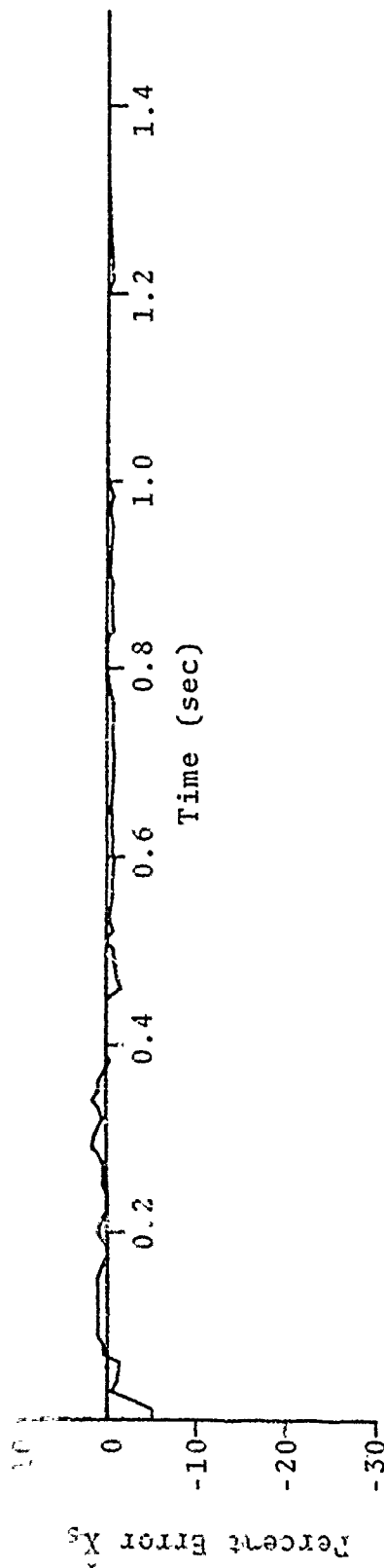


Figure 36. Variation of Percent Error in \hat{X}_5 with Time (Nonlinear System, Measurement Errors Only)

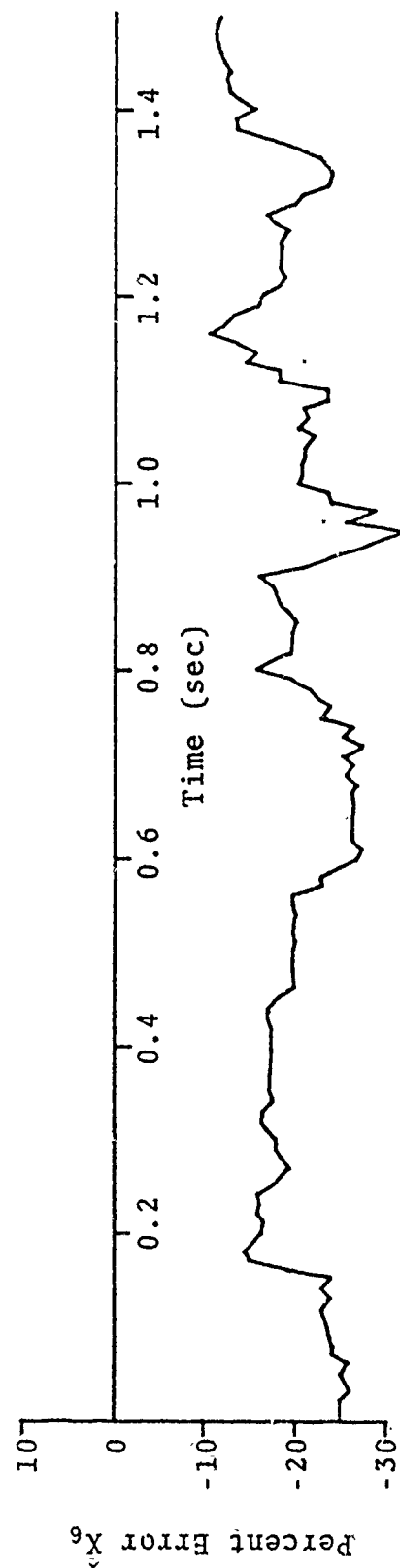


Figure 37. Variation of Percent Error in \hat{X}_6 with Time (Nonlinear System, Measurement Errors Only)

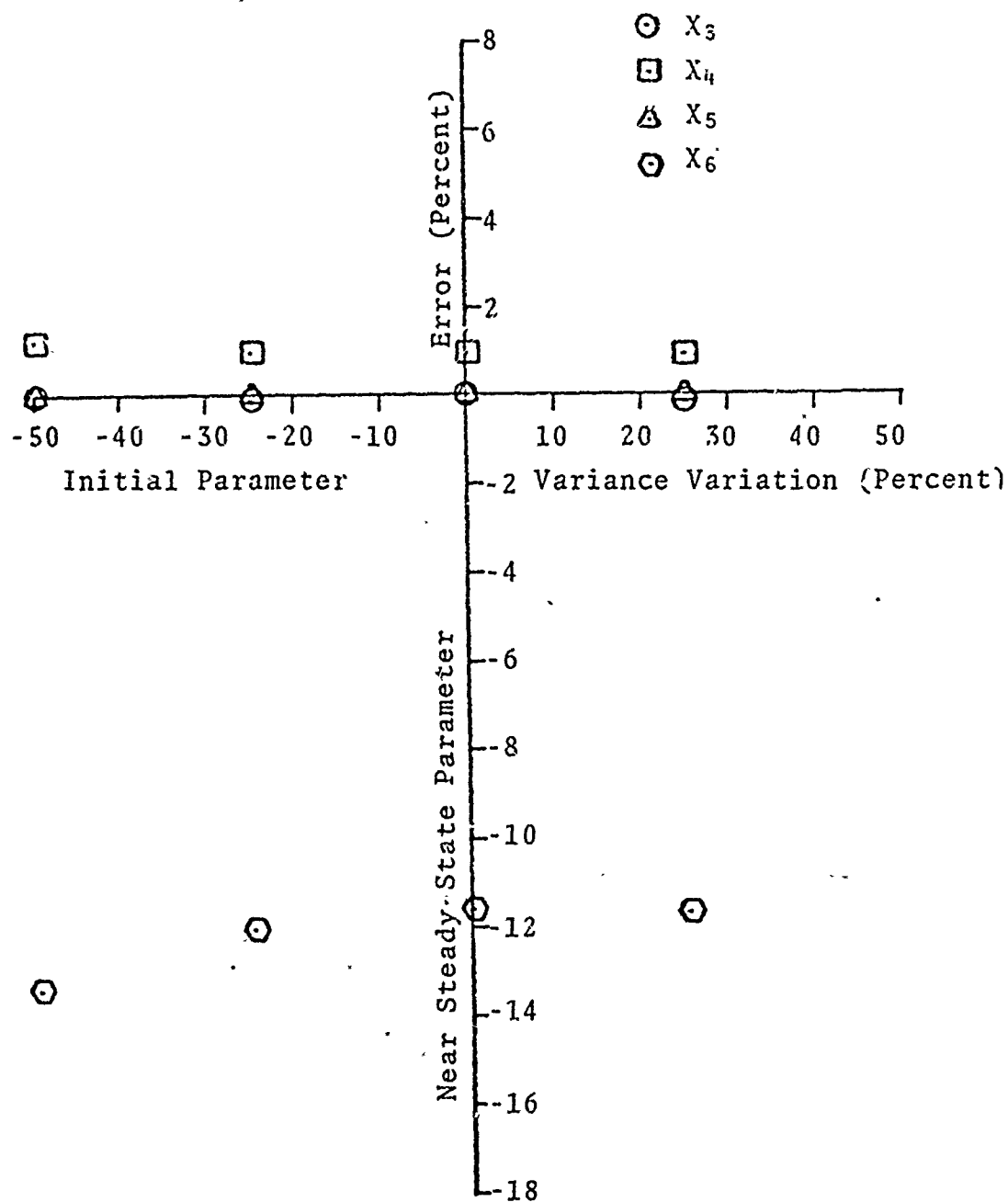


Figure 38. Variation of Near Steady-State Parameter Error with Initial Parameter Variance (Nonlinear System, Measurement Errors Only)

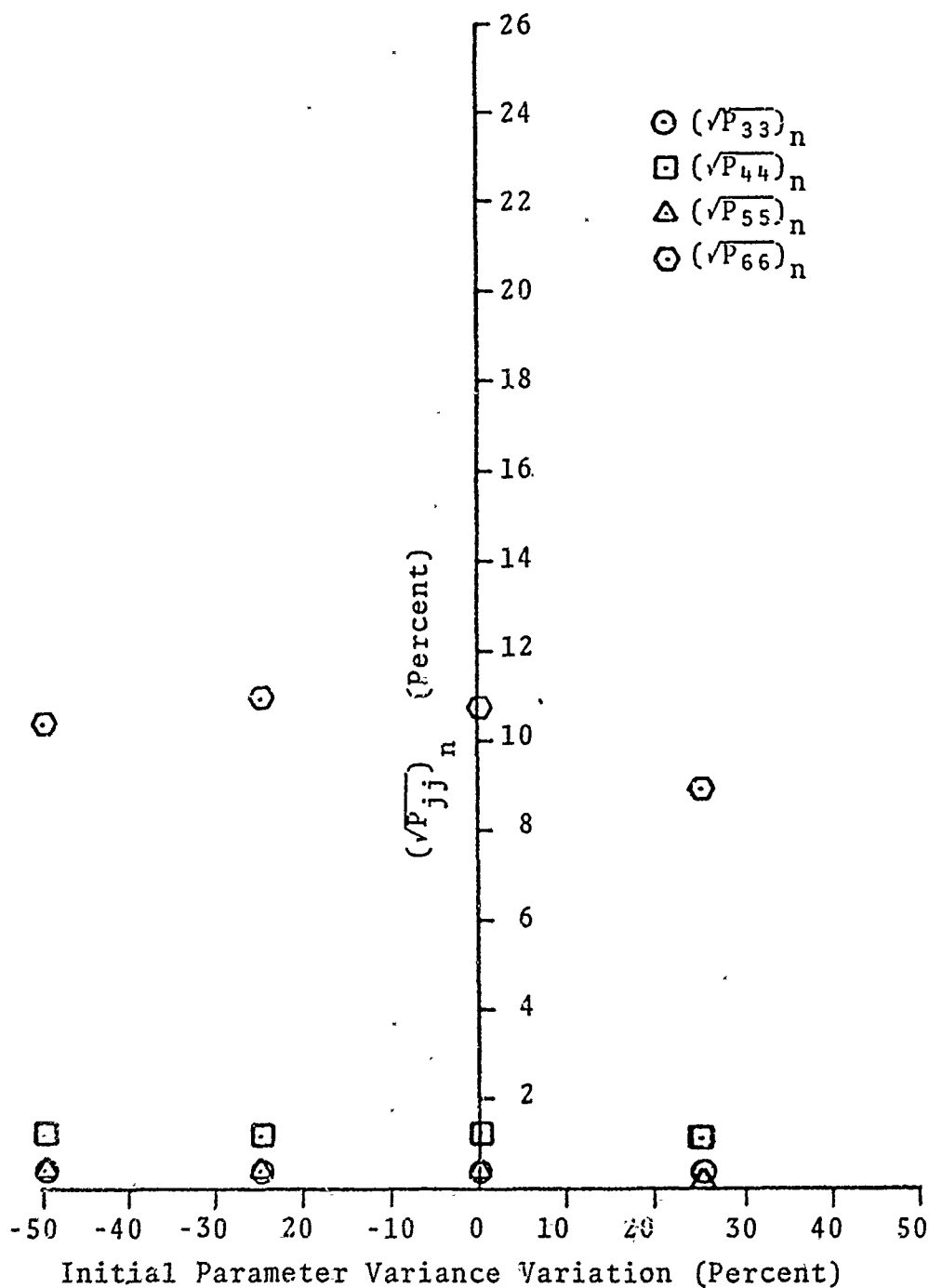


Figure 39. Variation of Normalized Near Steady-State Parameter Uncertainty with Initial Parameter Variance (Nonlinear System, Measurement Errors Only)

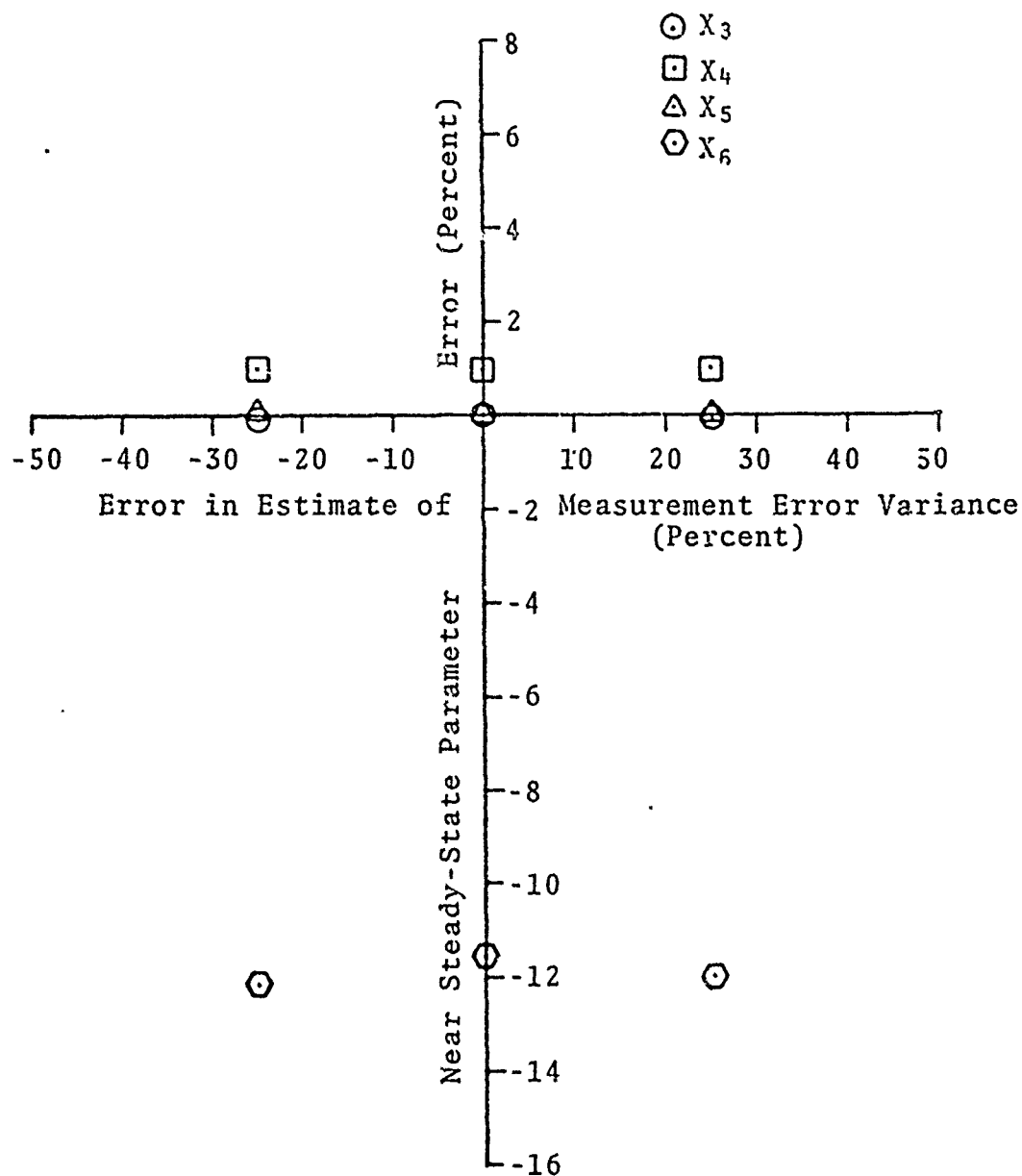


Figure 40. Variation of Near Steady-State Parameter Error with Error in the Estimate of the Measurement Error Variance (Nonlinear System, Measurement Errors Only)

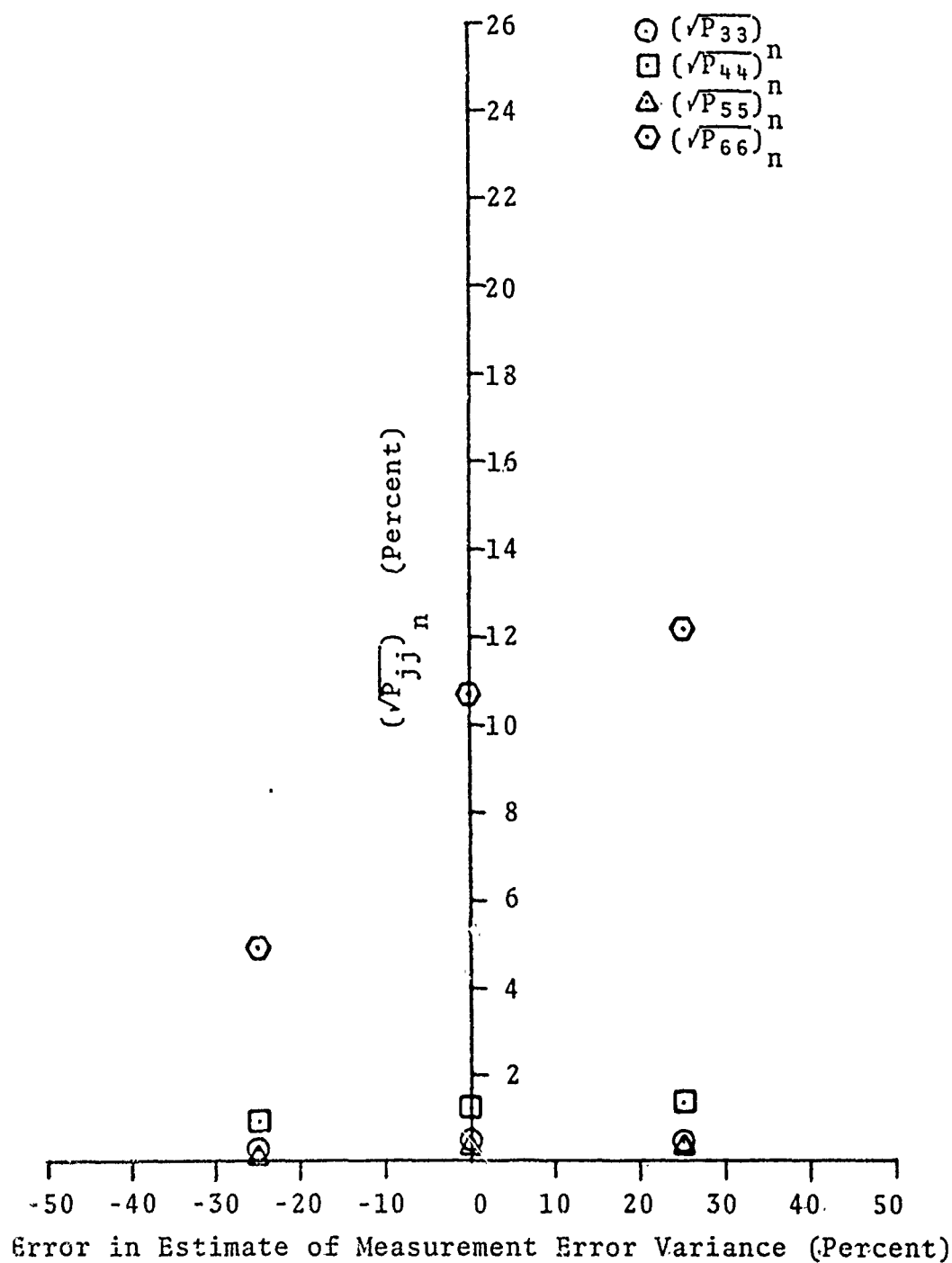


Figure 41. Variation of Normalized Near Steady-State Parameter Uncertainty with Error in the Estimate of Measurement Error Variance (Nonlinear System, Measurement Errors Only)

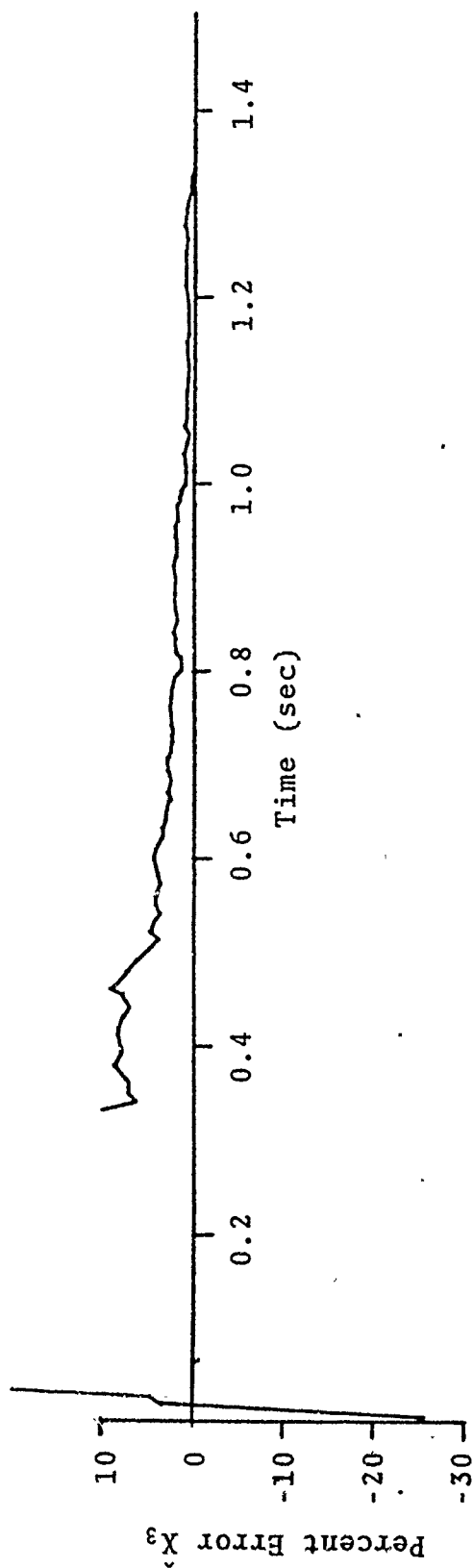


Figure 42. Variation of Percent Error in \hat{X}_3 with Time for Large Error in Initial Estimate of X_5 (Nonlinear System, Measurement Errors Only)

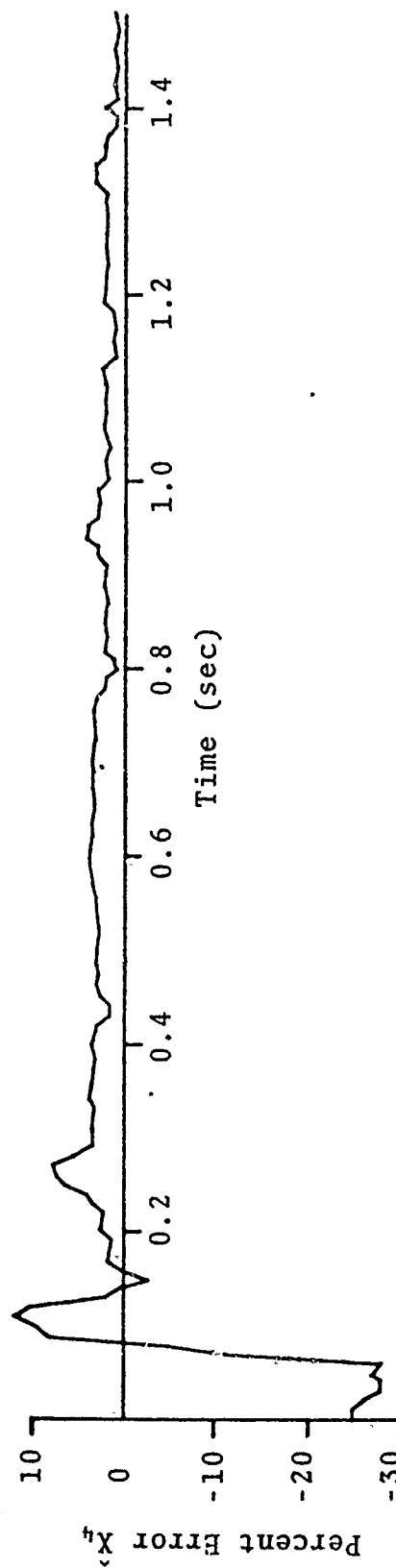


Figure 43. Variation of Percent Error in \hat{X}_4 with Time for Large Error in Initial Estimate of X_5 (Nonlinear System, Measurement Errors Only)

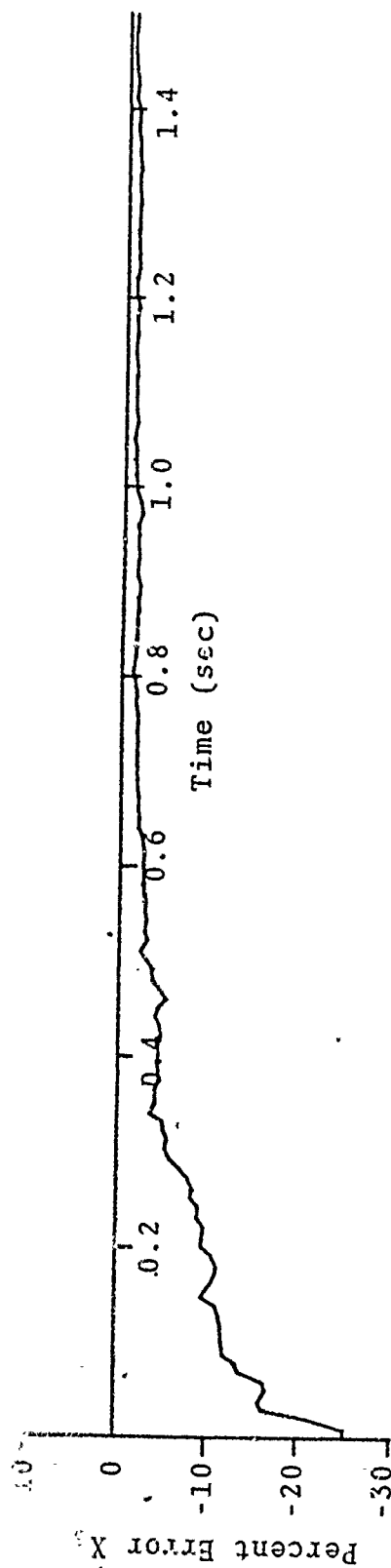


Figure 44. Variation of Percent Error in \hat{X}_5 with Time for Large Error in Initial Estimate of X_5 (Nonlinear System, Measurement Errors Only)

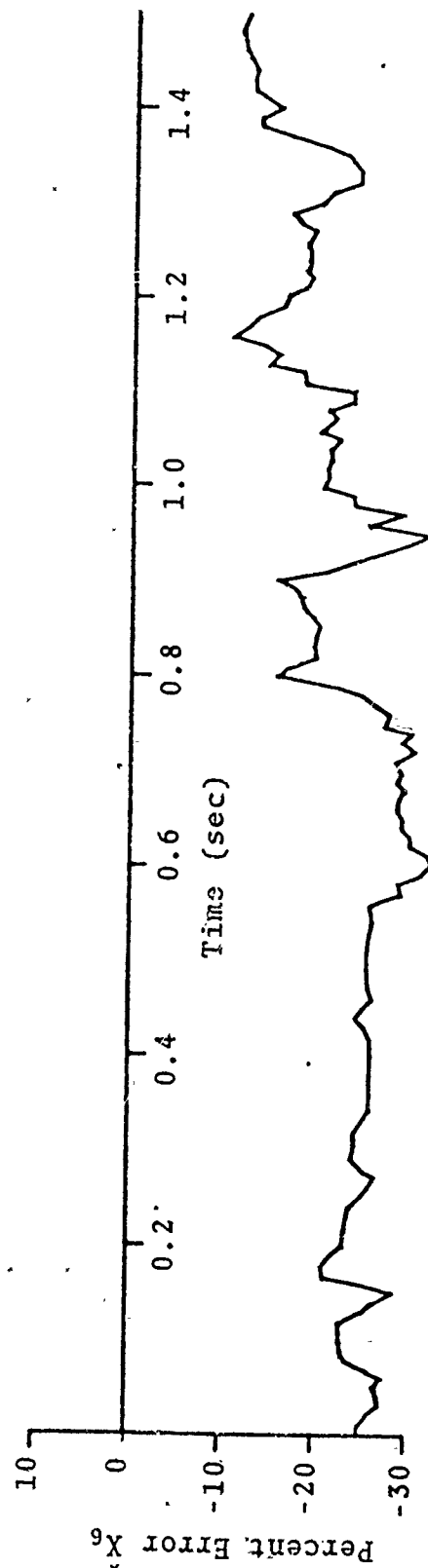


Figure 45. Variation of Percent Error in \hat{X}_6 with Time for Large Error in Initial Estimate of X_5 (Nonlinear System, Measurement Errors Only)

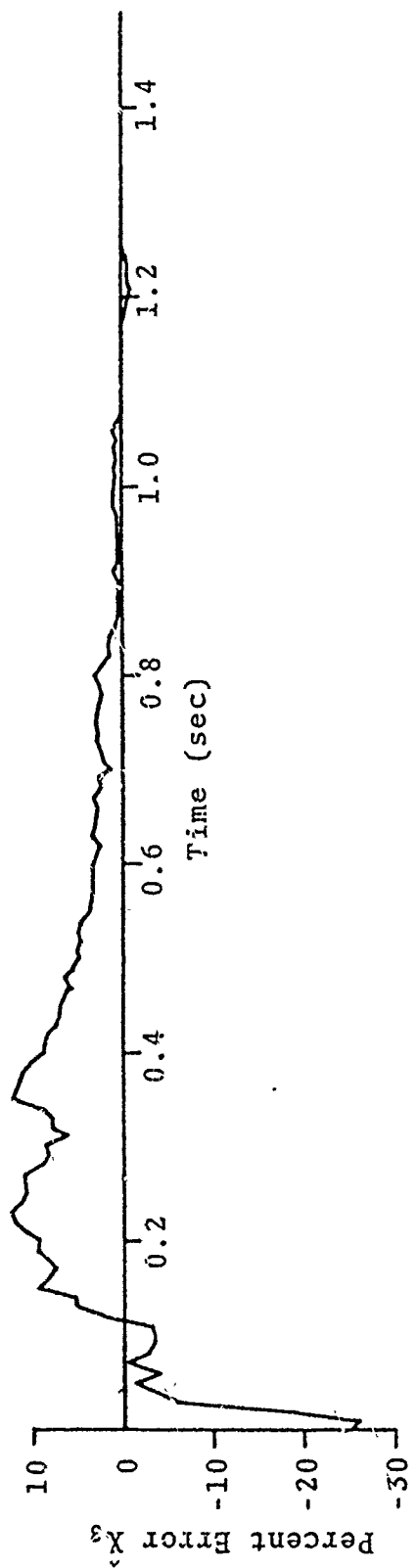


Figure 46. Variation of Percent Error in \hat{X}_3 with Time (Nonlinear System, Measurement Errors and System Noise)

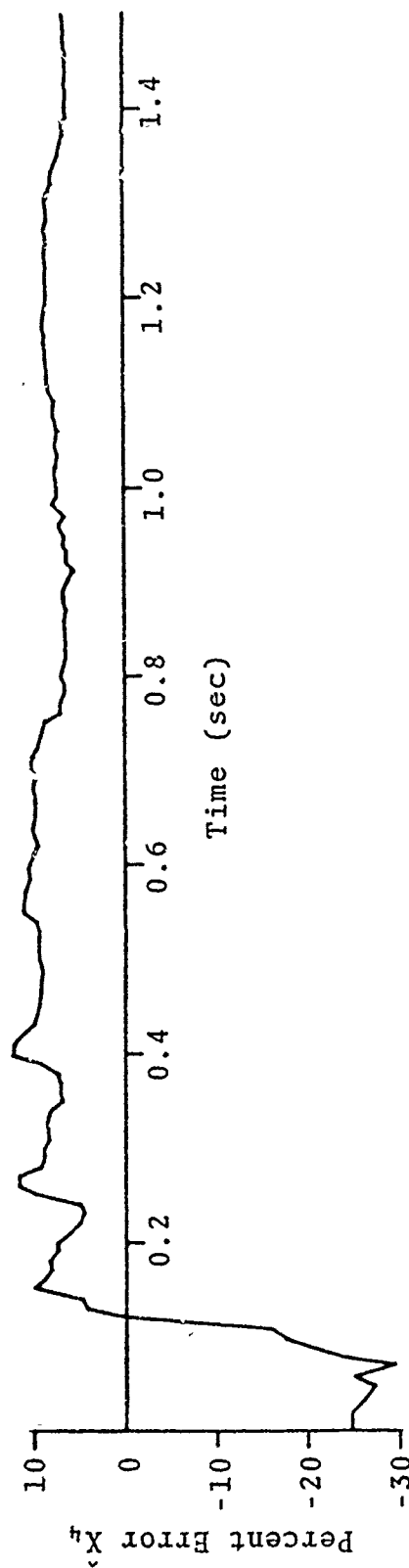


Figure 47. Variation of Percent Error in \hat{X}_4 with Time (Nonlinear System, Measurement Errors and System Noise)

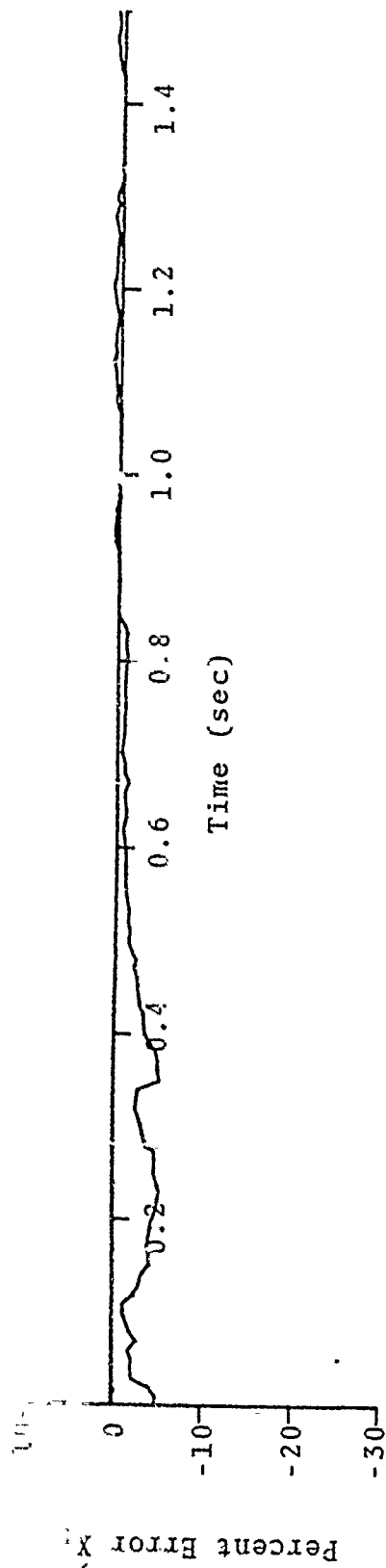


Figure 48. Variation of Percent Error in \hat{X}_5 with Time (Nonlinear System, Measurement Errors and System Noise)

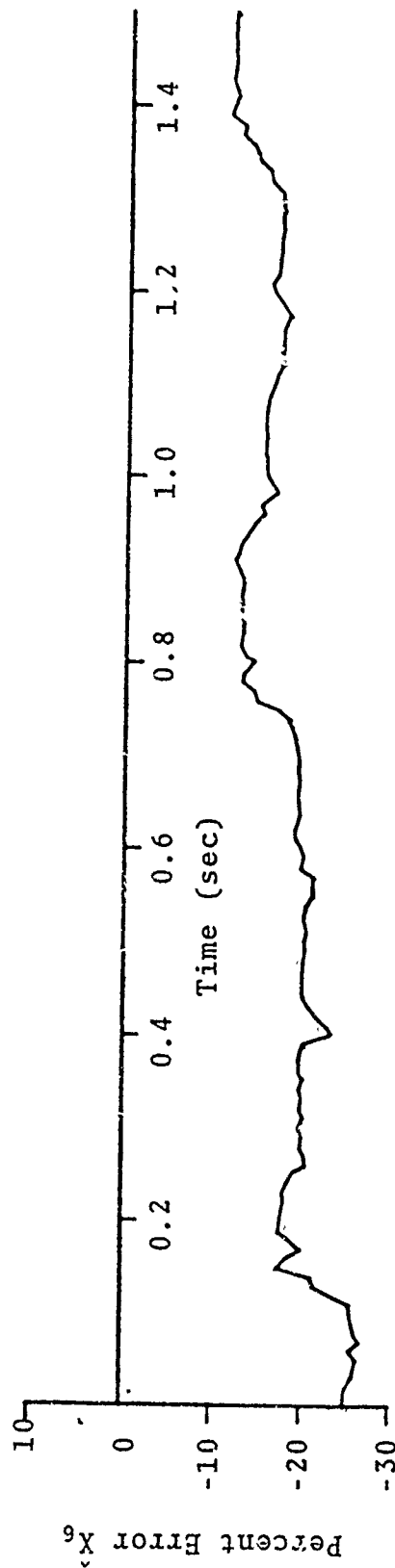


Figure 49. Variation of Percent Error in \hat{X}_6 with Time (Nonlinear System, Measurement Errors and System Noise)

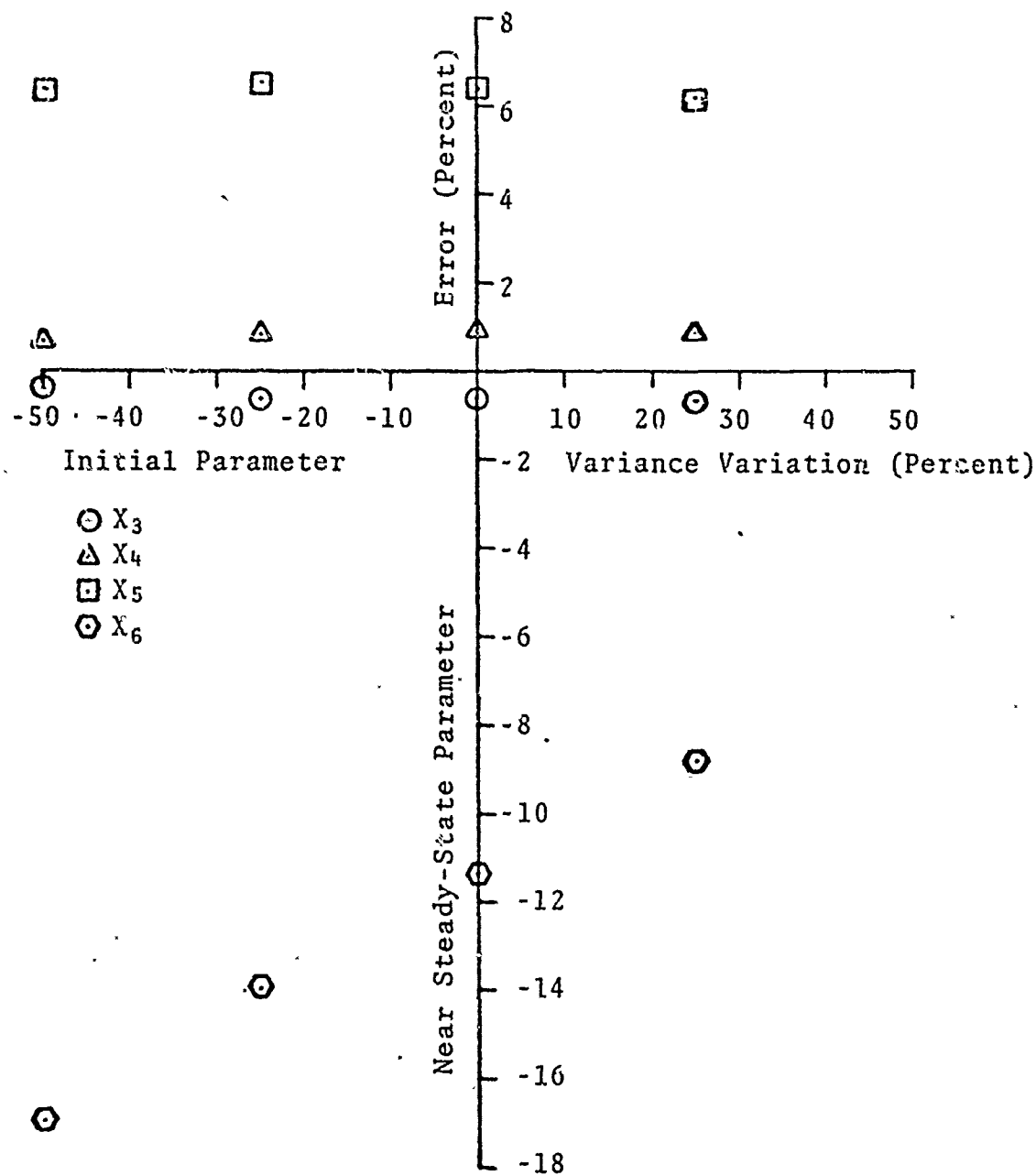


Figure 50. Variation of Near Steady-State Parameter Error with Initial Parameter Variance (Nonlinear System, Measurement Errors and System Noise)

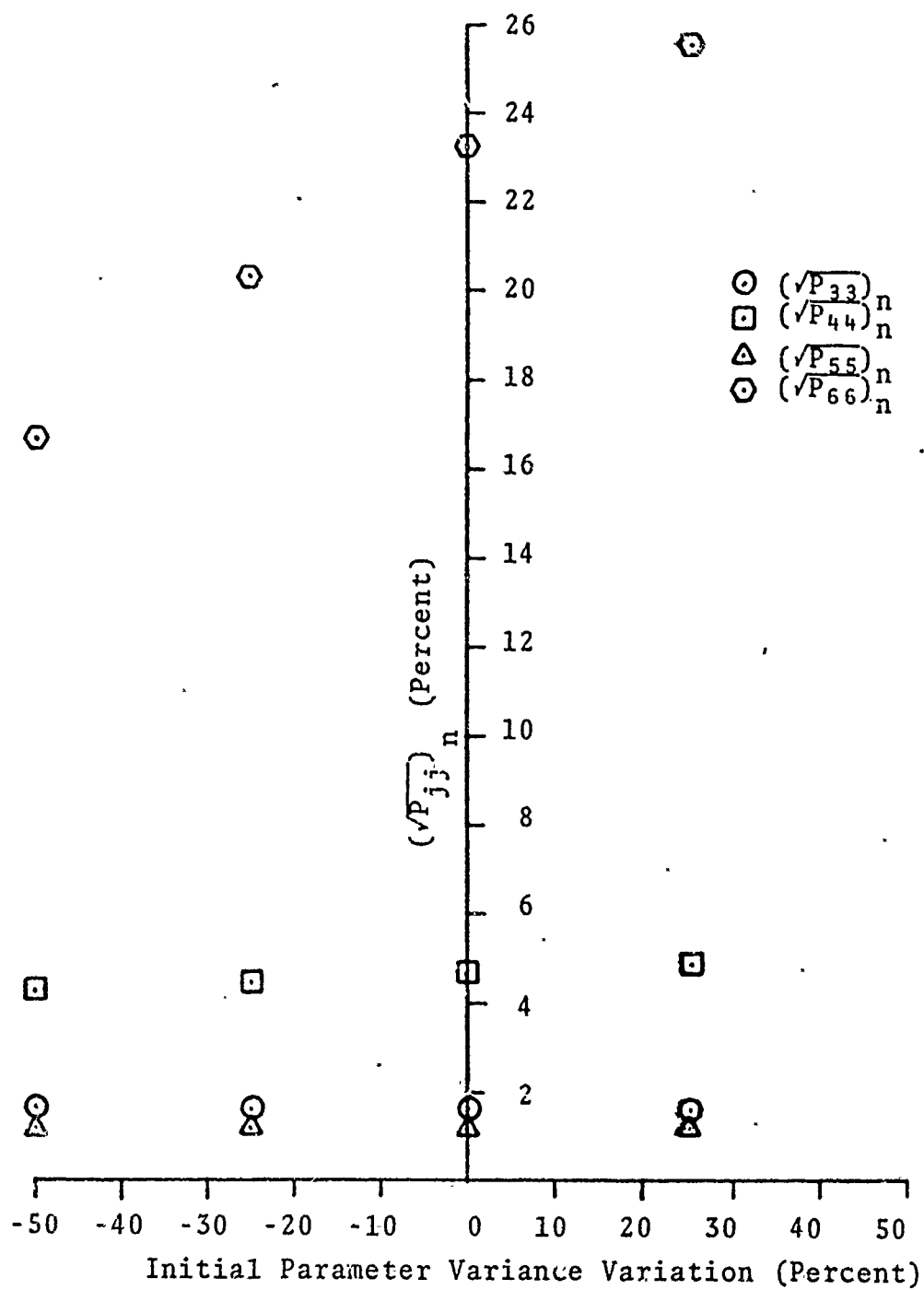


Figure 51. Variation of Normalized Near Steady-State Parameter Uncertainty with Initial Parameter Variance (Nonlinear System, Measurement Errors and System Noise)

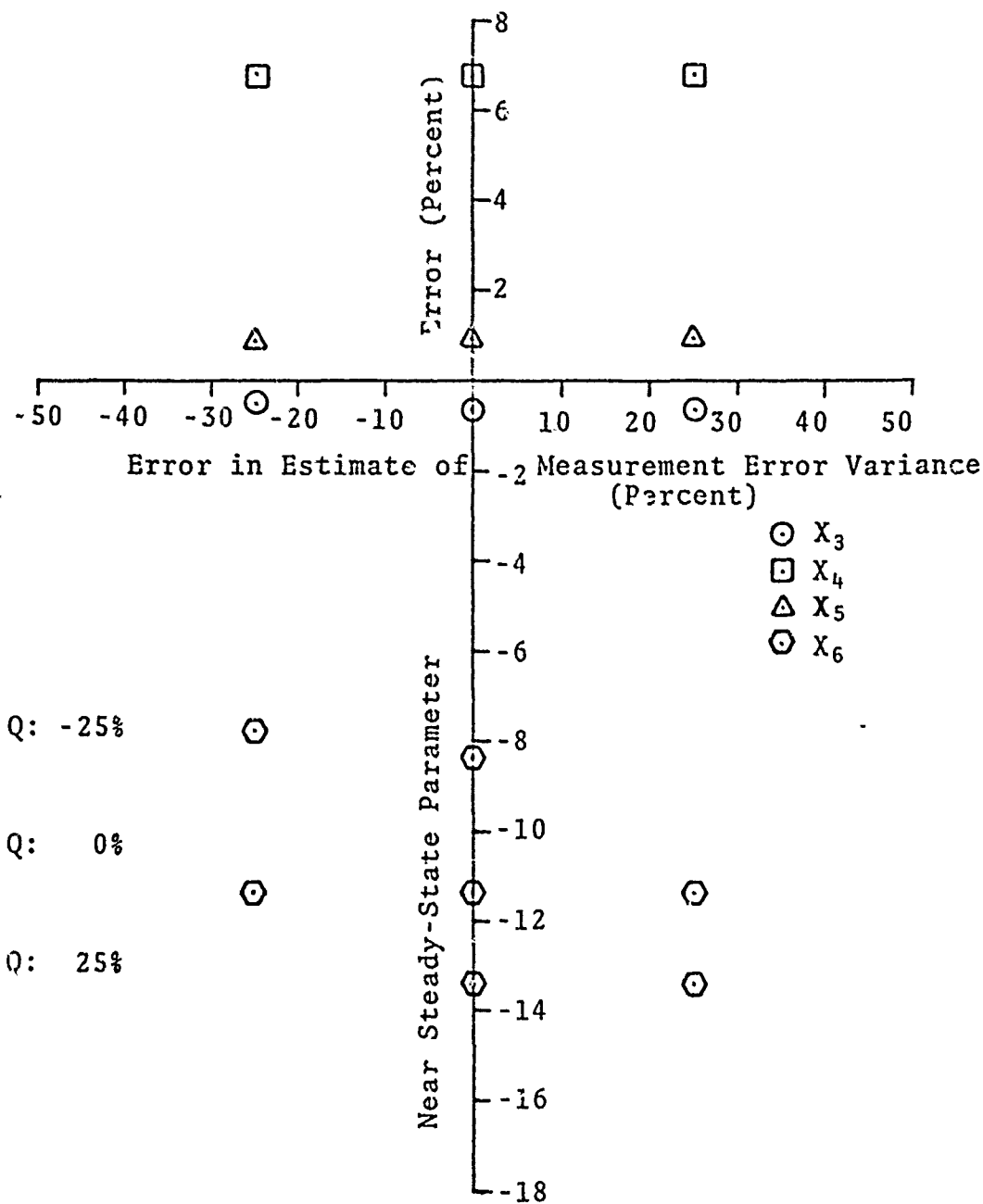


Figure 52. Variation of Near Steady-State Parameter Error with Error in the Estimate of the Measurement Error Variance (Nonlinear System, Measurement Errors and System Noise)

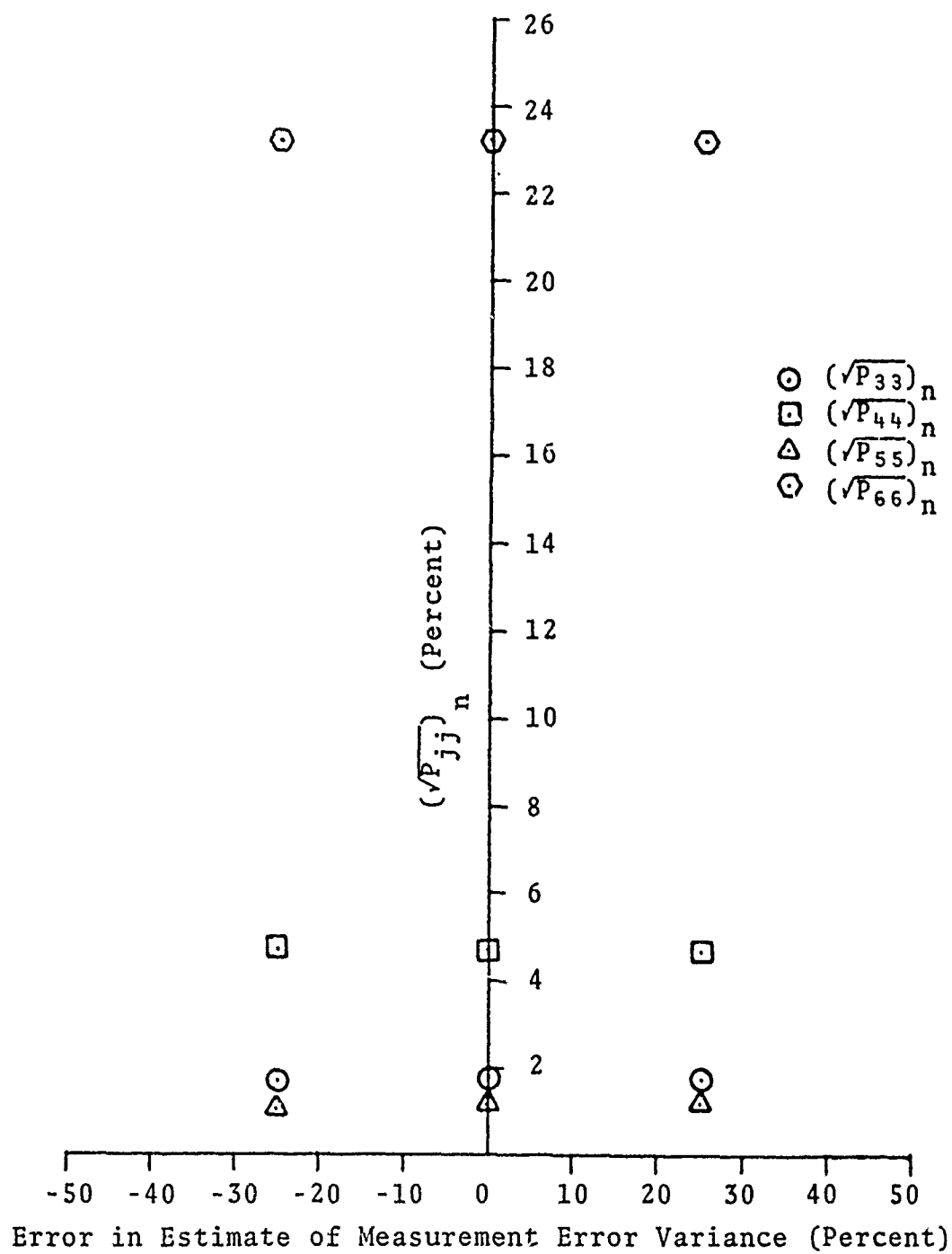


Figure 53. Variation of Normalized Near Steady-State Parameter Uncertainty with Error in the Estimate of Measurement Error Variance (Nonlinear System, Measurement Errors and System Noise)

SECTION VI

SUMMARY

1. General Comments

An analysis of methods for extracting aerodynamic coefficients from dynamic test data has been conducted. The primary concern of this analysis has been to determine the accuracy with which these coefficients can be determined from dynamic data containing system noise and measurement errors. Both linear and nonlinear systems have been considered with emphasis on the latter.

Two methods of extracting the coefficients of interest and estimating their uncertainties have been analyzed. The first of these is a deterministic technique due to Chapman and Kirk.¹⁰ The second technique considered, based on the extended Kalman filter^{20,31} with augmented state vector, is stochastic in nature and has been developed and applied in this report.

Both techniques considered will extract aerodynamic coefficients from noisy dynamic data. The degree to which they are successful has been presented in detail in the preceding chapters.

2. The Chapman-Kirk Technique

The presence of noise in the dynamic data being used by the Chapman-Kirk technique affects the accuracy with which this technique can be used to determine the coefficients of interest. When measurement errors only are present in the data, these effects have been found to be negligible on linear and nonlinear static pitching moment coefficient derivatives; the effects on the linear and nonlinear pitch damping coefficients are more pronounced, especially for the higher measurement errors. The error in the linear pitch damping coefficient is as high as 25 percent in one case. The greatest error in the extracted nonlinear pitch damping coefficient is 227 percent. There is an obvious inconsistency in the extracted values of this nonlinear coefficient since another data set, which also contained measurement errors with similar statistical properties, yielded a value for this term which was in error by only 5 percent.

The estimation of the uncertainties of the extracted coefficients when the data contained only measurement errors is quite adequate. In every case considered, the true value

of the coefficient of interest was found to be within three estimated standard deviations of the extracted value. The estimated standard deviations of a given coefficient are also consistent for different data sets with the same measurement error statistics.

Use of data containing measurement errors and system noise has a more adverse effect on the accuracy of coefficients determined with the Chapman-Kirk technique. As is the case when analyzing data containing only measurement errors, the pitch damping coefficients are the ones most affected. The errors in these coefficients range from 1 percent to 43 percent for the linear term and from 45 percent to 600 percent for the nonlinear term over the range of system noise considered.

Perhaps of even greater consequence than the errors in the damping parameters is the fact that their estimated uncertainties are too small to indicate the true errors in the coefficients when the data contain both system noise and measurement errors. This is a direct consequence of using the root-mean-square-error between the final solution to the equation of motion and the data to calculate the parameter uncertainties. This error term is an adequate representation of noise in the data only if all errors are randomly superimposed on the pitch angle output after the equation of motion has been integrated. The integrated effects of the random accelerations representing system noise, however, are such that the final errors appearing in the pitch angle due to the system noise are not necessarily random or independent from one discrete time to another.

3. Extended Kalman Filter

The feasibility of using an extended Kalman filter with parameter augmented state vector for determining the values of aerodynamic coefficients and their uncertainties from dynamic test data has been demonstrated for a one-degree-of-freedom system.

The specific filter used here generally reaches near steady-state estimates of the parameters in less than one second for the system model and error combinations considered.

For linear systems with measurement errors only in the data, the extended filter yielded estimates of both the static pitching moment coefficient derivative and the pitch damping coefficient to within 1 percent of their respective true values. Slightly less accurate determinations of the same two parameters are obtained using the filter when both measurement errors and system noise are present in the data. Variations in the initial parameter variances of 25 percent seem to affect essentially only the pitch damping term and then only when both measurement errors and system noise are present in the data.

For nonlinear systems with measurement errors only in the data the estimates obtained with the extended filter for the linear and nonlinear static pitching moment coefficient derivatives as well as the linear pitch damping coefficient are within approximately 1 percent or less of their respective true values for a variety of initial parameter variances and noise variance estimates. The error in the estimate of the nonlinear damping coefficient varies between 11 percent and 14 percent. None of the coefficients are particularly sensitive to variations in the initial parameter variances or noise variances considered. As is the case for the linear system, the extended filter is less accurate in its parameter identification when both measurement errors and system noise of the magnitudes considered here are present in the data. The linear damping coefficient is most affected by having the addition of system noise in the data. The error in this coefficient increases from approximately 1 percent to approximately 6 percent for the cases considered. Errors in the noise variance estimates of ± 25 percent have little or no appreciable effect on the linear and nonlinear static pitching moment coefficient derivatives or the linear pitch damping coefficient, and only a slight effect on the nonlinear pitch damping coefficient.

For the range of initial parameter errors, initial parameter variances, and noise variances considered, the extended Kalman filter produces an excellent estimate of the parameter uncertainties after integrating the Ricatti equations for approximately 1.5 to 2.0 seconds. For every case considered where the initial parameter estimates are within 25 percent of their true values, the near steady-state parameter value is within two standard deviations of its true value, and in the majority of cases within one standard deviation, when the standard deviation is taken to be the square root of the near steady-state parameter variance.

4. Comparison of the Two Techniques

There is a natural tendency when analyzing two techniques to compare the results achieved when solving similar problems. Both techniques investigated have strengths and weaknesses.

From the preceding discussion in this report, it is obvious that the Chapman-Kirk technique requires less information initially than the extended Kalman filter since it is only necessary to make initial estimates of the parameters in the former whereas the parameters, their variances, and the noise variances must all be estimated initially when using the extended filter.

There is no provision to weigh the initial parameter estimates when using the Chapman-Kirk technique whereas

this is done through the initial variances when using the extended filter. This feature is particularly important if certain parameters are known with significantly higher accuracy before the extraction process begins. The Chapman-Kirk technique will sometimes actually downgrade good initial estimates when analyzing noisy data whereas the extended filter will tend to improve the estimate or at least not downgrade it for the initial parameter variances considered.

From an economic point of view, both techniques are essentially equal concerning computational costs. The extended filter must numerically integrate more equations but only has to do this once; the Chapman-Kirk technique, being an iterative scheme, must integrate fewer equations more times.

The programming requirements of the extended filter are less than those of the Chapman-Kirk technique since the former requires no matrix inversions, simultaneous equation solutions, or logic for including or excluding certain parameters from the extraction process.

5. Areas for Additional Work

This research is part of a larger and continuing program whose goal is a critical study of methods for extracting aerodynamic coefficients from dynamic data. Since the research is of a continuing nature, it seems appropriate to mention some areas where additional work should be contemplated.

The results presented here are entirely for a one-degree-of-freedom system. The effect of system noise and measurement errors on coefficients extracted from coupled, multiple-degree-of-freedom systems may prove to be extremely interesting and is the next logical area to be investigated.

The response of the extended filter to highly erroneous initial parameter estimates seems quite good for the limited number of cases investigated. It seems possible to take advantage of this feature by constructing an extraction algorithm which employs both an extended filter and a Chapman-Kirk or similar technique in tandem with the extended filter providing the required accurate initial estimates for the differential corrections, iterative technique.

Additional work in relating the effects of inadequate or erroneous modeling to the noise problem should also prove to be of interest.

Finally, when considering new techniques such as the use of the extended Kalman filter, it must be concluded that additional numerical investigations will add to the confidence or more clearly define the limitations of this method of coefficient extraction.

APPENDIX I

CHAPMAN-KIRK COEFFICIENT EXTRACTION NOMENCLATURE LIST AND PROGRAM LISTING

Nomenclature List

<u>PROGRAM VARIABLE</u>	<u>MATH SYMBOL</u>	<u>DEFINITION</u>
H	Δt	Numerical integration step size (sec)
ITO		Frequency of numerical integration output
TMAX	t_{\max}	Cutoff time for numeri- cal integration (sec)
TZERO	t_0	Initial time for numerical integration (sec)
N	N	Number of first order differential equations to be integrated
XZ(1)	α_0	Initial condition of α (rad)
XZ(2)	$\dot{\alpha}_0$	Initial condition of $\dot{\alpha}$ (rad/sec)
XZ(3)	$\left. \frac{\partial \alpha}{\partial c_1} \right _{t=0}$	Initial condition for parametric differential equation
XZ(4)	$\left. \frac{d}{dt} \frac{\partial \alpha}{\partial c_1} \right _{t=0}$	Initial condition for parametric differntial equation
XZ(5)	$\left. \frac{\partial \alpha}{\partial c_2} \right _{t=0}$	Initial condition for parametric differential equation

<u>PROGRAM VARIABLE</u>	<u>MATH SYMBOL</u>	<u>DEFINITION</u>
XZ (6)	$\left. \frac{d}{dt} \frac{\partial \alpha}{\partial c_2} \right _{t=0}$	Initial condition for parametric differential equation
XZ (7)	$\left. \frac{\partial \alpha}{\partial c_3} \right _{t=0}$	Initial condition for parametric differential equation
XZ (8)	$\left. \frac{d}{dt} \frac{\partial \alpha}{\partial c_3} \right _{t=0}$	Initial condition for parametric differential equation
XZ (9)	$\left. \frac{\partial \alpha}{\partial c_4} \right _{t=0}$	Initial condition for parametric differential equation
XZ (10)	$\left. \frac{d}{dt} \frac{\partial \alpha}{\partial c_4} \right _{t=0}$	Initial condition for parametric differential equation
XZ (11)	$\left. \frac{\partial \alpha}{\partial c_5} \right _{t=0}$	Initial condition for parametric differential equation
XZ (12)	$\left. \frac{d}{dt} \frac{\partial \alpha}{\partial c_5} \right _{t=0}$	Initial condition for parametric differential equation
XZ (13)	$\left. \frac{\partial \alpha}{\partial c_6} \right _{t=0}$	Initial condition for parametric differential equation
XZ (14)	$\left. \frac{d}{dt} \frac{\partial \alpha}{\partial c_6} \right _{t=0}$	Initial condition for parametric differential equation
XZ (15)	$\left. \frac{\partial \alpha}{\partial c_7} \right _{t=0}$	Initial condition for parametric differential equation
XZ (16)	$\left. \frac{d}{dt} \frac{\partial \alpha}{\partial c_7} \right _{t=0}$	Initial condition for parametric differential equation

<u>PROGRAM VARIABLE</u>	<u>MATH SYMBOL</u>	<u>DEFINITION</u>
XZ (17)	$\left. \frac{\partial \alpha}{\partial C_8} \right _{t=0}$	Initial condition for parametric differential equation
XZ (18)	$\left. \frac{d}{dt} \frac{\partial \alpha}{\partial C_8} \right _{t=0}$	Initial condition for parametric differential equation
Q	q	Dynamic pressure (lb/ft ²)
V	v	Freestream velocity (ft/sec)
AMI	I	Moment of inertia about an axis through the vehicle C.G. and normal to the pitch plane (slug . ft ²)
A	A	Vehicle reference area (ft ²)
D	d	Vehicle reference length (ft)
CMAO	C _{mα0}	Static pitching moment coefficient derivative (rad ⁻¹)
CMA2	C _{mα2}	Static pitching moment coefficient derivative (rad ⁻³)
CMA4	C _{mα4}	Static pitching moment coefficient derivative (rad ⁻⁵)
CMQO	C _{mq0}	Pitch damping coeffi- cient
CMQ2	C _{mq2}	Pitch damping coeffi- cient (rad ⁻²)
CMDA	C _{mδα}	Static pitching moment coefficient at α = 0

<u>PROGRAM VARIABLE</u>	<u>MATH SYMBOL</u>	<u>DEFINITION</u>
M	M	Number of experimental data points
JJJMAX		Maximum number of iterations allowed before program termination
EP	ϵ	Convergence criteria for change in RMS ERROR (rad)
AE(I)	$\alpha_{\text{exp}}(t_i)$	Experimental values of α (rad)
C1	C_1	α_0
C2	C_2	$\dot{\alpha}_0$
C3	C_3	$-\frac{C_{m\alpha_0} q Ad}{I}$
C4	C_4	$-\frac{C_{mq_0} q Ad^2}{2VI}$
C5	C_5	$-\frac{C_{m\alpha_2} q Ad}{I}$
C6	C_6	$-\frac{C_{mq_2} q Ad^2}{2VI}$
C7	C_7	$-\frac{C_{m\alpha_4} q Ad}{I}$
C8	C_8	$-\frac{C_{m\delta\alpha} q Ad}{I}$

Categories of Input Data

Numerical integration constants.--These constants include the numerical integration step size, frequency of output, time at which the numerical integration is to stop, initial time, and the number of first order equations to be integrated.

Initial conditions.--These data are the first estimates of the initial conditions for the equation of motion as well as the initial conditions for the parametric differential equations.

Aerodynamic and physical constants.--These constants are the dynamic pressure and freestream velocity that existed when the experimental pitch angle data were recorded, the moment of inertia about an axis normal to the pitch plane and through the center of gravity of the vehicle of interest, a vehicle reference area, and a vehicle reference length.

Aerodynamic coefficient estimates.--These data are the first estimates of the aerodynamic coefficients to be extracted.

Convergency criteria.--These constants are the number of experimental data points to be read, the maximum number of iterations to be allowed, and the tolerance against which the root-mean-square error difference is tested for convergence.

Experimental data.--These data are the experimental pitch angle values from which the aerodynamic coefficients are to be extracted. The program assumes the time between each data point is constant, and this time must be equal to the product of the numerical integration step size and frequency of numerical integration output that were given on the first data card.

```

DIMENSION XZ(18),AE(500),X(120),PC(8,500),AC(500),
IAJK(4,4),B(8),DELC(8),L1(4),L2(4),DIF(500),STDEV(8),
2ACDOT(500)
COMMON N,T,X
COMMON C1,C2,C3,C4,C5,C6,C7,C8,KKK,M,PC,AC,ACDOT
EXTERNAL XDOT,OUT
JJJ=0
READ(5,1) H,ITO,TMAX,TZERO,N
1 FORMAT (F5.3,I3,2F5.3,I3)
READ(5,2) (XZ(I),I=1,N)
2 FORMAT(2F10.4)
READ(5,3) Q,V,AMI,A,D,
3 FORMAT(5F8.4)
READ(5,4) CMAO,CMA2,CMA4,CMQO,CMQ2,CMDA
4 FORMAT(6F10.3)
IF(N.EQ.2) M=TMAX/H
IF(N.EQ.2) GO TO 55
READ(5,41) M,JJMAX,EP
41 FORMAT(2I3,E10.3)
READ(5,5) (AE(I),I=1,M)
5 FORMAT(10F8.5)
55 CONTINUE
CON1=(-Q*A*D**2/(2.0*V*AMI))
CON2=(-Q*A*D/AMI)
C1=XZ(1)
C2=XZ(2)
C3=CON2*CMAO
C4=CON1*CMQO
C5=CON2*CMA2
C6=CON1*CMQ2
C7=CON2*CMA4
C8=CON2*CMQ4
WRITE(6,52)
59 FORMAT('0',5X,'NUMERICAL INTEGRATION CONSTANTS')

```

```

WRITE(6,60) H,ITO,TMAX,TZERO,N
60 FORMAT('0',5X,'H(SEC)=' ,F5.3,5X,'ITO=' ,I3,5X,'TMAX(SEC=' ,
1F5.3,5X,'TZERO(SEC)=' ,F5.3,5X,'N=' ,I3)
WRITE(6,61)
61 FORMAT('0',5X,'AERODYNAMIC INPUT DATA')
WRITE(6,6) Q,V,AMI,A,D
6 FORMAT('0',5X,'Q(LB/FT**2)=' ,F8.4,5X,'V(FT/SEC)=' ,F8.4,5X,
1'AMI(SLG*FT**2)=' ,F8.4,5X,'A(FT**2)=' ,F8.4,5X,'D(FT)=' ,
2F8.4)
WRITE(6,62)
62 FORMAT('0',5X,'AERODYNAMIC COEFFICIENT ESTIMATES')
WRITE(6,7) CMAO,CMA2,CMA4,CMQO,CMQ2,CMDA,M
7 FORMAT('0',5X,'CMAO=' ,F9.3,5X,'CMA2=' ,F9.3,5X,'CMA4=' ,F9.3,
15X,'CMQO=' ,F9.3,5X,'CMQ2=' ,F9.3,5X,'CMDA=' ,F9.3,5X,'M=' ,I3)
IF(N.EQ.2) GO TO 88
WRITE(6,63)
63 FORMAT('0',5X,'EXPERIMENTAL ANGLE OF ATTACK DATA')
WRITE(6,8) (I,AE(I),I=1,M)
8 FORMAT('0',5X,'AE(' ,I3,' )=' ,F8.5,5X,'AE(' ,I3,' )=' ,F8.5,5X,
1'AE(' ,I3,' )=' ,F8.5,5X,'AE(' ,I3,' )=' ,F8.5,5X,'AF(' ,I3,' )=' ,
2F8.5)
88 CONTINUE
KKK=0
CALL ADDUM(H,ITO,TZERO,TMAX,XZ,XDOT,OUT)
IF(N.EQ.2) GO TO 148
IF(JJT.GT.O) GO TO 880
WRITE(6,879)
879 FORMAT('1',5X,'CURRENT PARAMETER VALUES')
GO TO 881
880 WRITE(6,881)
881 FORMAT('0',5X,'CURRENT PARAMETER VALUES')
881 WRITE(6,882) JJJ,C1,C2,C3,C4,C5,C6,C7,C8
882 FORMAT('0',5X,'JJJ=' ,I2,5X,'C1=' ,E12.5,5X,'C2=' ,E12.5,5X,
1'C3=' ,E12.5,5X,'C4=' ,E12.5,5X,'C5=' ,E12.5,5X,'C6=' ,

```



```

ZE12.5 5X, 'C8=' ,E12.5,5X, 'C8=' ,E12.5)
NN=(N-2)/2
DO 890 J=1,8
  DELC(J)=0.
  STDEV(J)=0.
890 CONTINUE
  DO 89 J=1,NN
    B(J)=0.
    DO 89 K=1,NN
      AJK(J,K)=0.
89 CONTINUE
    DO 9 L=1,M
      DO 9 J=1,NN
        DIF(L)=AE(L)-AC(L)
        B(J)=B(J)+PC(J,L)*DIF(L)
        DO 9 K=1,J
          AJK(J,K)=AJK(J,K)+PC(J,L)*PC(K,L)
          IF(J.EQ.K) GO TO 9
          AJK(K,J)=AJK(J,K)
9 CONTINUE
        WRITE(6,900) ((AJK(J,K),K=1,NN),J=1,NN)
        FORMAT(6E15.5)
        CALL MINV(AJK,NN,R,L1,L2)
        WRITE(6,902) ((AJK(J,K),K=1,NN),J=1,NN)
        FORMAT(6E15.5)
        WRITE(6,901) (B(J),J=1,NN)
        FORMAT(6E15.5)
        R=0.
        DO 883 L=1,M
          R=R+DIF(L)*DIF(L)
883 CONTINUE
          IF(JJ.GT.0) GO TO 91
          RMSE=0.
91 RMSEP=RMSE

```

```

RMSR=SQRT(R/M)
RMSE=SQRT(R/(M-NN))
DIFF=ABS(RMSE-RMSEP)
DO 910 J=1,NN
STDEV(J)=RMSE*(SQRT(AJK(J,J)))
910 CONTINUE
WRITE(6,92) RMSR
92 FORMAT('0',5X,'RMS RESIDUAL=',E16.8)
WRITE(6,93) RMSE
93 FORMAT('0',5X,'RMS ERROR=',E16.8)
WRITE(6,94)
94 FORMAT('0',5X,
1'ESTIMATED STANDARD DEVIATIONS OF THE PARAMETERS')
WRITE(6,95) STDEV
95 FORMAT('0',16X,'SDC1=',E12.5,5X,'SDC2=',E12.5,5X,'SDC3=',
1E12.5,5X,'SDC4=',E12.5,5X,'SDC5=',E12.5,5X,'SDC6=',
2E12.5,5X,'SDC7=',E12.5,5X,'SDC8=',E12.5)
IF(DIFF.LT.EP) GO TO 13
IF(JJJ.GE.JJJMAX) GO TO 121
JJJ=JJJ*1
DO 10 I=1,NN
DO 10 J=1,NN
10 DELC(I)=DELC(I)+AJK(I,J)*B(J)
WRITE(6,101)
101 FORMAT('0',5X,'CURRENT PARAMETER CORRECTIONS')
WRITE(6,102) JJJ,DELC
102 FORMAT('0',5X,'JJJ=',I2,5X,'DELC1=',E12.5,5X,'DELC2=',
1E12.5,5X,'DELC3=',E12.5,5X,'DELC4=',E12.5,5X,'DELC5=',
2E12.5,5X,'DELC6=',E12.5,5X,'DELC7=',E12.5,5X,'DELC8=',
3E12.5)
GO TO 1222
121 WRITE(6,122)
122 FORMAT('0
1'ITERATIONS EXCEEDED')
CONVERGENCE FAILED - MAXIMUM NUMBER OF '

```

```

1222 C1=C1+DELC(1)
      C2=C2+DELC(2)
      C3=C3+DELC(3)
      C4=C4+DELC(4)
      C5=C5+DELC(5)
      C6=C6+DELC(6)
      C7=C7+DELC(7)
      C8=C8+DELC(8)
      XZ(1)=C1
      XZ(2)=C2
      GO TO 88
13 CONTINUE
      CMAO=C3/CON2
      CMQO=C4/CON1
      CMA2=C5/CON2
      CMQ2=C6/CON1
      CMA4=C7/CON2
      CMDA=C8/CON2
      SDCMAO=STDEV(3)/(-CON2)
      SDCMQO=STDEV(4)/(-CON1)
      SDCMA2=STDEV(5)/(-CON2)
      SDCMQ2=STDEV(6)/(-CON1)
      SDCMA4=STDEV(7)/(-CON2)
      SDCMDA=STDEV(8)/(-CON2)
      WRITE(6,139)
139 FORMAT('1 EXTRACTED INITIAL CONDITIONS AND ',
1 ' AERODYNAMIC COEFFICIENTS')
      WRITE(6,14) C1,C2
14 FORMAT('0',5X,'ALPHAO=',F6.4,5X,'ALDOTO=',F6.4)
      WRITE(6,140) STDEV(1),STDEV(2)
140 FORMAT('0',5X,'SDA=',F9.4,5X,'SDADOT=',F6.4)
      WRITE(6,15) CMAO,CMA2,CMA4,CMQO,CMQ2,CMDA
15 FORMAT('0',5X,'CMAO=',F9.3,5X,'CMA2=',F9.3,5X,'CMA4=',

```

```

146 1F9.3,5X,'CMQO=' ,F9.3,5X,'CMQ2=' ,F9.3,5X,'CMDA=' ,F9.3)
    WRITE(6,147) SDCMAO,SDCMA2,SDCMA4,SDCMQO,SDCMQ2,SDCMDA
147 FORMAT('0',5X,'SDCMAO=' ,F7.3,5X,'SDCMA2=' ,F7.3,5X,
1' SDCMA4=' ,F7.3,5X,'SDCMQO=' ,F7.3,5X,'SDCMQ2=' ,F7.3,5X,
2' SDCMDA=' ,F7.3)
148 CONTINUE
    IF(N.GT.2) GO TO 1491
    WRITE(7,1490) (AC(I,I=1,M)
1490 FORMAT(10F8.5)
1491 CONTINUE
    WRITE(6,149)
149 FORMAT('0',5X,'CALCULATED ANGLE OF ATTACK OUTPUT')
    WRITE(6,150) (I,AC(I),I=1,M)
150 FORMAT('0',5X'AC(' ,I3,')=' ,F8.5,5X,'AC(' ,I3,')=' ,F8.5,5X
1' AC(' ,I3,')=' ,F8.5,5X,AC(' ,I3,')=' ,F8.5,5X,'AC(' ,I3,')=' ,
2F8.5)
16 CONTINUE
    STOP
    END

```

SUBROUTINE ADDUM(AITCH, IEO, TZERO, TMAX, XZ, F, OUT)
 DIMENSION XZ(20), C(20,6), X(20,6)

COMMON N, T, X,
 H = ABS(AITCH)

IT = ITO-1

D=TZERO-TMAX

IF (D) 2,1,1

1 H=-H

2 HH=0.5*H

D=H/24.0

T=TZERO

ISET=0

DO 3 I=1,N

X(I,1)=XZ(I)

3 X(I,6)=XZ(I)

CALL F(X(1,5),1)

CALL F(X(1,2),0)

IF(ITO) 5,5,4

4 CALL OUT(1)

5 DO 6 K=1,N

C(K,1) = X(K,5)*HH

6 X(K,1) = X(K,1) + C(K,1)

T=T+HH

CALL F(C(1,2),3)

DO 7 K=1,N

C(K,2) = HH*C(K,2)

7 X(K,1) = X(K,6) + C(K,2)

CALL F(C(1,3),3)

DO 8 K=1,N

C(K,3) = H*C(K,3)

8 X(K,1) = X(K,6) + C(K,3)

T=T+HH

CALL F(C(1,4),1)

DO 9 K=1,N

ADDU 100

ADDU 110

ADDU 120

ADDU 130

ADDU 140

ADDU 150

ADDU 160

ADDU 170

ADDU 180

ADDU 190

ADDU 200

ADDU 210

ADDU 220

ADDU 230

ADDU 240

ADDU 250

ADDU 260

ADDU 270

ADDU 280

ADDU 290

ADDU 300

ADDU 310

ADDU 320

ADDU 330

ADDU 340

ADDU 350

ADDU 360

ADDU 370

ADDU 380

ADDU 390

ADDU 400

ADDU 410

ADDU 420

ADDU 430

C(K,4) = C(K,4)*H
X(K,1)=X(K,6)+(C(K,1)+2.0*C(K,2)+C(K,3)+0.5*C(K,4))
*0.33333333

1

9 X(K,6) = X(K,1)

ISET = ISET+1

GO TO (10,12,17), ISET

10 CALL F(X(1,5),0)

DO 11 K = 1,N

11 X(K,3)=X(K,5)

GO TO 18

12 CALL F(X(1,5),0)

DO 13 K=1,N

13 X(K,4)=X(K,5)

GO TO 18

14 T=T+H

DO 15 K=1,N

X(K,1)=X(K,6)+D*(55.0*X(K,5)-59.0*X(K,4)+37.0*X(K,3)

1 -9.0*X(K,2))

X(K,2)=X(K,3)

X(K,3)=X(K,4)

15 X(K,4)=X(K,5)

CALL F(X(1,5),1)

DO 16 K = 1,N

X(K,6)=X(K,6)+D*(9.0*X(K,5)+19.0*X(K,4)-5.0*X(K,3)

+X(K,2))

1 X(K,1)=X(K,6)

ISET = 3

17 CALL F(X(1,5),0)

18 IF (IT)20,19,20

19 IT = ITO

CALL OUT(0)

20 IT = IT - 1

IF (ABS (TMAX-T) -ABS (HH)) 21,21,22

21 RETURN

ADDU 450

ADDU 460

ADDU 470

ADDU 480

ADDU 490

ADDU 500

ADDU 510

ADDU 520

ADDU 530

ADDU 540

ADDU 550

ADDU 560

ADDU 570

AD U 5 0

ADDU 600

ADDU 610

ADDU 620

ADDU 630

++ 4

ADDU 670

4

ADDU 690

ADDU 700

ADDU 710

ADDU 720

ADDU 730

22 GO TO (5,5 1.), SGT
END

ADU 750

```

SUBROUTINE XDOT(A,K)
  DIMENSION A(18),X(120),PC(8,500),AC(500),ACDOT(500)
  COMMON N,T,X
  COMMON C1,C2,C3,C4,C5,C6,C7,C8,KKK,M,PC,AC,ACDOT
  EQUIVALENCE (X(1),ALPHA),(X(2),ALDOT),(X(3),P1),(X(4),
1P1DOT),(X(5),P2),(X(6),P2DOT),(X(7),P3),(X(8),P3DOT),(X(9),
2P4),(X(10),P4DOT),(X(11),P5),(X(12),P5DOT),(X(13),P6),
3(X(14),P6DOT),(X(15),P7),(X(16),P7DOT),(X(17),P8),(X(18),
4P8DOT)
  A(1)=ALDOT
  A1=-(C4 + C6*ALPHA**2)*ALDOT
  A2=-(C3+C5*ALPHA**2+C7*ALPHA**4)*ALPHA
  A(2)=A1+A2-C8
  IF(N.EQ.2) GO TO 1
  A3=C4+C6*ALPHA**2
  A4=2*C6*ALPHA*ALDOT+C3+3*C5*ALPHA**2+5*C7*ALPHA**4
  A(3)=PIDOT
  A(4)=-A3*P1DOT-A4*P1
  A(5)=P2DOT
  A(6)=-A3*P2DOT-A4*P2
  A(7)=P3DOT
  A(8)=-A3*P3DOT-A4*P3-ALPHA
  IF(N.EQ.8) GO TO 1
  A(9)=P4DOT
  A(10)=-A3*P4DOT-A4*P4-ALDOT
  IF(N.EQ.10) GO TO 1
  A(11)=P5DOT
  A(12)=-A3*P5DOT-A4*P5-ALPHA**3
  IF(N.EQ.12) GO TO 1
  A(13)=P6DOT
  A(14)=-A3*P6DOT-A4*P6-(ALPHA**2)*ALDOT
  IF(N.EQ.14) GO TO 1
  A(15)=P7DOT
  A(16)=-A3*P7DOT-A4*P7-ALPHA**5

```



```
IF(N.P2.16) GO TO 1
A(17)=P8DOT
A(18)=-A3*P8DOT-A4*P8-1.0
1 CONTINUE
RETURN
END
```

```

SUBROUTINE OUT(K)
  DIMENSION X(120),PC(8,500),AC(500),ACDOT(500)
  COMMON N,T,X
  COMMON C1,C2,C3,C4,C5,C6,C7,C8,KKK,M,PC,AC,ACDOT
  KKK=KKK+1
  IF(N.EQ.2) GO TO 1
  PC(1,KKK)=X(3)
  PC(2,KKK)=X(5)
  PC(3,KKK)=X(7)
  IF(N.EQ.8) GO TO 1
  PC(4,KKK)=X(9)
  IF(N.EQ.10) GO TO 1
  PC(5,KKK)=X(11)
  IF(N.EQ.12) GO TO 1
  PC(6,KKK)=X(13)
  IF(N.EQ.14) GO TO 1
  PC(7,KKK)=X(15)
  IF(N.EQ.16) GO TO 1
  PC(8,KKK)=X(17)
  1 AC(KKK)=X(1)
  ACDOT(KKK)=X(2)
  RETURN
  END

```

.....	MINV 10
	MINV 20
	MINV 30
	MINV 40
	MINV 50
	MINV 60
	MINV 70
	MINV 80
	MINV 90
	MINV 100
	MINV 110
	MINV 120
	MINV 130
	MINV 140
	MINV 150
	MINV 160
	MINV 170
	MINV 180
	MINV 190
	MINV 200
	MINV 210
	MINV 220
	MINV 230
	MINV 240
	MINV 250
	MINV 260
	MINV 270
	MINV 280
	MINV 290
	MINV 300
	MINV 310
	MINV 320
	MINV 330
	MINV 340


```

SUBROUTINE MINV
  PURPOSE
    INVERT A MATRIX
  USAGE
    CALL MINV(A,N,D,L,)
  DESCRIPTION OF PARAMETERS
    A - INPUT MATRIX, DESTROYED IN COMPUTATION AND REPLACED BY
        RESULTANT INVERSE.
    N - ORDER OF MATRIX A
    D - RESULTANT DETERMINANT
    L - WORK VECTOR OF LENGTH N
    M - WORK VECTOR OF LENGTH N
  REMARKS
    MATRIX A MUST BE A GENERAL MATRIX
  SUBROUTINES AND FUNCTION SUBPROGRAMS REQUIRED
    NONE
  METHOD
    THE STANDARD GAUSS-JORDAN METHOD IS USED. THE DETERMINANT
    IS ALSO CALCULATED. A DETERMINANT OF ZERO INDICATES THAT
    THE MATRIX IS SINGULAR.
  SUBROUTINE MINV(A,N,D,L,M)
    DIMENSION A(1), L(1), M(1)

```

C	MINV	350
C	MINV	360
C	MINV	370
C	MINV	380
C	MINV	390
C	MINV	400
C	MINV	420
C	MINV	410
C	MINV	430
C	MINV	440
C	MINV	450
C	MINV	460
C	MINV	470
C	MINV	480
C	MINV	490
C	MINV	500
C	MINV	510
C	MINV	520
C	MINV	530
C	MINV	540
C	MINV	550
C	MINV	560
C	MINV	570
C	MINV	580
C	MINV	590
C	MINV	600
C	MINV	610
C	MINV	620
C	MINV	630
C	MINV	640
C	MINV	650
C	MINV	660
C	MINV	670
C	MINV	680


```

.....
IF A DOUBLE PRECISION VERSION OF THIS ROUTINE IS DESIRED, THE
C IN COLUMN 1 SHOULD BE REMOVED FROM THE DOUBLE PRECISION
STATEMENT WHICH FOLLOWS.
DOUBLE PRECISION A,D,BIGA,HOLD

THE C MUST ALSO BE REMOVED FROM DOUBLE PRECISION STATEMENTS
APPEARING IN OTHER ROUTINES USED IN CONJUNCTION WITH THIS
ROUTINE.

THE DOUBLE PRECISION VERSION OF THIS SUBROUTINE MUST ALSO
CONTAIN DOUBLE PRECISION FORTRAN FUNCTIONS. ABS IN STATEMENT
10 MUST BE CHANGED TO DABS.
.....
SEARCH FOR LARGEST ELEMENT

D=1.0
NK=-N
DO 80 K=1,N
NK=NK+N
L(K)=K
M(K)=K
KK=NK+K
BIGA=A(KK)
DO 20 J=K,N
IZ=N*(J-1)
DO 20 I=K,N
IJ=IZ+I
10 IF( ABS(BIGA) - ABS(A(IJ))) 15,20,20

```

MINV 690
MINV 700
MINV 710
MINV 720
MINV 730
MINV 740
MINV 750
MINV 760
MINV 770
MINV 780
MINV 790
MINV 800
MINV 810
MINV 820
MINV 830
MINV 840
MINV 850
MINV 860
MINV 870
MINV 880
MINV 890
MINV 900
MINV 910
MINV 920
MINV 930
MINV 940
MINV 950
MINV 960
MINV 970
MINV 980
MINV 990
MINV1000
MINV1010
MINV1020

15 BIGA=A(IJ)

L(K)=I

M(K)=J

20 CONTINUE

C
C
C

INTERCHANGE ROWS

J=L(K)

IF (J-K) 35,35,25

25 KI=K-N

DO 30 I=1,N

KI=KI+N

HOLD=-A(KI)

JI=KI-K+J

A(KI)=A(JI)

30 A(JI)=HOLD

C
C
C

INTERCHANGE COLUMNS

35 I=M(K)

IF (I-K) 45,45,38

38 JP=N*(I-1)

DO 40 J=1,N

JK=NK+J

JI=JP+J

HOLD=-A(JK)

A(JK)=A(JI)

40 A(JI)=HOLD

C
C
C
C

DIVIDE COLUMN BY MINUS PIVOT (VALUE OF PIVOT ELEMENT IS
CONTAINED IN BIGA)

45 IF (BIGA) 48,46,48

46 D=0.0

MINV1030
MINV1040
MINV1050
MINV1060
MINV1070
MINV1080
MINV1090
MINV1100
MINV1110
MINV1120
MINV1130
MINV1140
MINV1150
MINV1160
MINV1170
MINV1180
MINV1190
MINV1200
MINV1210
MINV1220
MINV1230
MINV1240
MINV1250
MINV1260
MINV1270
MINV1280
MINV1290
MINV1300
MINV1310
MINV1320
MINV1330
MINV1340
MINV1350
MINV1360

```

RETURN
48 DO 55 I=1,N
   IF(I-K) 50,55,50
50 IK=NK+I
   A(IK)=A* IK)/(-BIGA)
55 CONTINUE

      REDUCE MATRIX
C
C
C
DO 65 I=1,N
IK=NK+I
HOLD=A(IK)
IJ=I-N
DO 65 J=1,N
IJ=IJ+N
IF(I-K) 60,65,60
60 IF(J-K) 62,65,62
62 KJ=IJ-I+K
   A(IJ)=HOLD*A(KJ)+A(IJ)
65 CONTINUE

      DIVIDE ROW BY PIVOT
C
C
C
KJ=K-N
DO 75 J=I,N
KJ=KJ+N
IF(J-K) 70,75,70
70 A(KJ)=A(KJ)/BIGA
75 CONTINUE

      PRODUCT OF PIVOTS
C
C
C
D=D*BIGA
C

```

MINV1370
 MINV1380
 MINV1390
 MINV1400
 MINV1410
 MINV1420
 MINV1430
 MINV1440
 MINV1450
 MINV1460
 MINV1470
 MINV1480
 MINV1490
 MINV1500
 MINV1510
 MINV1520
 MINV1530
 MINV1540
 MINV1550
 MINV1560
 MINV1570
 MINV1580
 MINV1590
 MINV1600
 MINV1610
 MINV1620
 MINV1630
 MINV1640
 MINV1650
 MINV1660
 MINV1670
 MINV1680

C REPLACE PIVOT BY RECIPROCAL
 C A(KK)=1.0/PIGA
 80 CONTINUE
 C FINAL ROW AND COLUMN INTERCHANGE
 C K=N
 C 100 K=(K-1)
 IF(K) 150,150,105
 105 I=L(K)
 IF(I-K) 120,120,108
 108 JQ=N*(K-1)
 JR=N*(I-1)
 DO 110 J=1,N
 JK=JQ+J
 HOLD=A(JK)
 JI=JR+J
 A(JK)=-A(JI)
 110 A(JI) =HOLD
 120 J=M(K)
 IF(J-K) 100,100,125
 125 KI=K-N
 DO 130 I=1,N
 KI=KI+N
 HOLD=A(KI)
 JI=KI-K+J
 A(KI)=-A(JI)
 130 A(JI) +HOLD
 GO TO 100
 150 RETURN
 END

APPENDIX II

EXTENDED KALMAN FILTER NOMENCLATURE LIST AND PROGRAM LISTING

Nomenclature List

<u>PROGRAM VARIABLE</u>	<u>MATH SYMBOL</u>	<u>DEFINITION</u>
H	Δt	Numerical integration step size (sec)
TZERO	t_0	Initial time for num- erical integration (sec)
TMAX	t_{\max}	Cutoff time for numerical integration (sec)
ITO		Frequency of numerical integration output
N	N	Number of first order equations to be integrated
M	M	Number of data points
XZ(I)		Initial conditions for state variable and variance estimates
Q	q	System noise variance (rad ² /sec ⁴)
R	r	Measurement error variance (rad ²)
Z(I)	z	Experimental data values of α
X(1)	\hat{x}_1	Estimated value of α (rad)

<u>PROGRAM VARIABLE</u>	<u>MATH SYMBOL</u>	<u>DEFINITION</u>
X(2)	\hat{X}_2	Estimated value of \dot{a} (rad/sec)
X(3)	\hat{X}_3	Estimated value of C_3
X(4)	\hat{X}_4	Estimated value of C_4
X(5)	\hat{X}_5	Estimated value of C_5
X(6)	\hat{X}_6	Estimated value of C_6
X(7)	P_{11}	Variance of \hat{X}_1
X(8)	P_{12}	Covariance of \hat{X}_1 and \hat{X}_2
X(9)	P_{13}	Covariance of \hat{X}_1 and \hat{X}_3
X(10)	P_{14}	Covariance of \hat{X}_1 and \hat{X}_4
X(11)	P_{15}	Covariance of \hat{X}_1 and \hat{X}_5
X(12)	P_{16}	Covariance of \hat{X}_1 and \hat{X}_6
X(13)	P_{22}	Variance of \hat{X}_2
X(14)	P_{23}	Covariance of \hat{X}_2 and \hat{X}_3
X(15)	P_{24}	Covariance of \hat{X}_2 and \hat{X}_4
X(16)	P_{25}	Covariance of \hat{X}_2 and \hat{X}_5
X(17)	P_{26}	Covariance of \hat{X}_2 and \hat{X}_6
X(18)	P_{33}	Variance of \hat{X}_3
X(19)	P_{34}	Covariance of \hat{X}_3 and \hat{X}_4
X(20)	P_{35}	Covariance of \hat{X}_3 and \hat{X}_5

<u>PROGRAM VARIABLE</u>	<u>MATH SYMBOL</u>	<u>DEFINITION</u>
X(21)	P_{36}	Covariance of \hat{X}_3 and \hat{X}_6
X(22)	P_{44}	Variance of \hat{X}_4
X(23)	P_{45}	Covariance of \hat{X}_4 and \hat{X}_5
X(24)	P_{46}	Covariance of \hat{X}_4 and \hat{X}_6
X(25)	P_{55}	Variance of \hat{X}_5
X(26)	P_{56}	Covariance of \hat{X}_5 and \hat{X}_6
X(27)	P_{66}	Variance of \hat{X}_6
A(I)	\dot{X}	Derivative of X with respect to time

Categories of Input Data

Numerical integration constants.--These constants are the numerical integration step size, frequency of output, time at which the integration is to stop, initial time, the number of first order equations being integrated, and the number of data points to be read.

Initial conditions.--These data are the initial estimates of the state variable components, state variable variances, and covariances.

Noise variance estimates.--These two entries are the estimates of the measurement error variance and system noise variance.

Pitch angle data.--These data are the experimental values of the pitch angle from which the filter estimates the states and parameters. The time between data points is not input but is assumed to be the same as the numerical integration step size input on the first data card.

```

C EXTENDED KALMAN FILTER WITH AUGMENTED STATE VECTOR NON LINEAR
C
      DIMENSION X(162),XZ(27),Z(500),AC(500)
      COMMON N,T,X
      COMMON Q,Z,J,JJITO,KI,R,ITD,JJJ,AC
      EXTERNAL XDOT,OUT
501 READ(5,501,END=999) H,TZERO,TMAX,ITO,N,M,ITD,NP
501 FORMAT(3F10.4,5I10)
502 READ(5,502) (XZ(I),I=1,N)
502 FORMAT(5E15.5)
503 READ(5,503) Q,R
503 FORMAT(2E15.5)
504 READ(5,504) (Z(I),I=1,M)
504 FORMAT(10F8.5)
505 WRITE(6,601) H,TZERO,TMAX,ITO,N,M,ITD
6010 FORMAT('1',5X,'INCREMENT=' ,F5.3,'SEC',5X,'TZERO=' ,F5.3,5X
1'TMAX=' ,F5.3,5X,'ITO=' ,I3,5X,'N=' ,I3,5X,'M-' ,I3,5X,'ITD=' ,I3)
505 WRITE(6,602) (I,XZ(I),I=1,N)
6020 FORMAT(1H0,5X,'XZ(' ,I2,' )=' ,E15.5,10X,'XZ(' ,I2,' )=' ,E15.5)
505 WRITE(6,603) Q,R
603 FORMAT(1H0,5X,'Q=' ,E15.5,5X,'R=' ,E15.5)
505 WRITE(6,604) (I,Z(I),I=1,M)
6040 FORMAT(1H0,5X,'Z(' ,I3,' )=' ,F8.5,5X,'Z(' ,I3,' )=' ,F8.5,5X,'Z(' ,I3,
1')=' ,F8.5,5X,'Z(' ,I3,' )=' ,F8.5,5X,'Z(' ,I3,' )=' ,F8.5)
      JJJ=0
      J=0
      JJITO=ITO
      KI=-1
      CALL ADDUM (H,ITO,TZERO,TMAX,XZ,XDOT,OUT)
      IF(NP.EQ.0) GO TO 990
      WRITE(7,900) (AC(I),I=1,M)
      900 FORMAT(10F8.5)
      990 GO TO 001
      999 STOP
      END

```

```

SUBROUTINE ADDUM(AITCH,ITO,TZERO,TMAX,XZ,E,OUT)
DIMENSION XZ(27),C(27,6),X(27,6)
COMMON N,T,X,
H = ABS(AITCH)
IT -ITO-1
D=TZERO-TMAX
IF (D)2,1,1
1 H=-H
2 HH=0.5*H
D=H/24.0
T=TZERO
ISET=0
DO 3 I=1,N
X(I,1)=XZ(I)
3 X(I,6)=XZ(I)
CALL F(X(1,5),1)
CALL F(X(1,2),0)
IF(ITO)5,5,4
4 CALL OUT(1)
5 DO 6 K=1,N
C(K,1) = X(K,5)*HH
X(K,1) = X(K,1) + C(K,1)
T=T+HH
CALL F(C(1,2),3)
DO 7 K=1,N
C(K,2) = HH*C(K,2)
X(K,1) = X(K,6) + C(K,2)
7 CALL F(C(1,3),3)
DO 8 K=1,N
C(K,3) = H*C(K,3)
X(K,1) = X(K,6) + C(K,3)
T=T+HH
CALL F(C(1,4),1)
DO 9 K=1,N

```

```

ADDU 110
ADDU 120
ADDU 130
ADDU 140
ADDU 150
ADDU 160
ADDU 170
ADDU 180
ADDU 190
ADDU 200
ADDU 210
ADDU 220
ADDU 230
ADDU 240
ADDU 250
ADDU 260
ADDU 270
ADDU 280
ADDU 290
ADDU 300
ADDU 310
ADDU 320
ADDU 330
ADDU 340
ADDU 350
ADDU 360
ADDU 370
ADDU 380
ADDU 390
ADDU 400
ADDU 410
ADDU 420

```

```

      C(K,4) = C(K,4)*H
      X(K,1)=X(K,6)+(C(K,1)+2.0*C(K,2)+C(K,3)+0.5*C(K,4))*0.333333333
      X(K,6) = X(K,1)
9      ISET = ISET+1
      GO TO (10,12,17), ISET
10     CALL F(X(1,5),0)
      DO 11 K= 1,N
11     X(K,3)=X(K,5)
      GO TO 18
12     CALL F(X(1,5),01)
      DO 13 K=1,N
13     X(K,4)=X(K,5)
      GO TO 18
14     T=T+H
      DO 15 K=1,N
      X(K,1)=X(K,6)+D*(55.0*X(K,5)-59.0*X(K,4)+37.0*X(K,3)-9.0*X(K,2))
      X(K,2)=X(K,3)
      X(K,3)=X(K,4)
      X(K,4)=X(K,5)
15     CALL F(X(1,5),1)
      DO 16 K = 1,N
      X(K,6)=X(K,6)+D*(9.0*X(K,5)+19.0*X(K,4)-5.0*X(K,3)+X(K,2))
      X(K,1)=X(K,6)
16     ISET = 3
17     CALL F(X(1,5),0)
18     IF (IT(20,19,20)
19         IT = ITO
          CALL OUT())
20     IT = IT - 1
      IF (ABS(TMAX-T)-ABS(HH)) 21,21,22
21     RETURN
22     GO TO (5,5,14), ISET
      END

```

```

ADDU 430
ADDU 440
ADDU 450
ADDU 460
ADDU 470
ADDU 480
ADDU 490
ADDU 500
ADDU 510
ADDU 520
ADDU 530
ADDU 540
ADDU 550
ADDU 560
ADDU 570
ADDU 580
ADDU 590
ADDU 600
ADDU 610
ADDU 620
ADDU 640
ADDU 670
ADDU 690
ADDU 700
ADDU 710
ADDU 730
ADDU 750

```

```

SUBROUTINE ADIF (X, Y, Z, KI)
  DIMENSION X(162), A(27), Z(500)
  COMMON N,T,X
  COMMON Q,Z,J,JJITO,KI,R,ITD,JJJ,AC
  II=0
  IF (J.GT.ITD) II=1
  IF (K.EQ.3) GO TO 81
  KI=KI+K
  IF (KI.NE.0) GO TO 70
  IF (K.EQ.0) GO TO 70
  J=J+1
  70 ADIF= (Z(J) - (Z(J) - Z(J+1)) * (KI/JJITO)) - X(1)
  IF (KI+1) .LT. JJITO) GO TO 81
  IF (K.EQ.1) GO TO 81
  KI=-1
  81 E=2*X(1)*X(2)*X(6)+X(3)+3*X(1)**2*X(5)
  B=X(4)+X(6)*X(1)**2
  C=X(1)**3
  D=X(1)**2*X(2)
  A(1)=X(2)+(X(7)/R)*ADIF
  A(2)=- (X(4)+X(6)*X(1)**2)*X(2) - (X(3)+X(5)*X(1)**2)*
  X(1) + (X(8)/R)*ADIF
  A(3)=(X(9)/R)*ADIF
  A(4)=(X(10)/R)*ADIF
  A(5)=(X(11)/R)*ADIF
  A(6)=(X(12)/R)*ADIF
  A(7)=2*X(8)-X(7)**2/R
  A(8)=X(13)-E*X(7)-B*X(8)-X(1)*X(9)-X(2)*X(10)-C*X(11)-
  D*X(12)-X(7)*X(8)/R
  A(9)=X(14)-X(7)*X(9)/R
  A(10)=X(15)-X(7)*X(10)/R
  A(11)=X(16)-X(7)*X(11)/R
  A(12)=X(17)-X(7)*X(12)/R
  A(13)=Q*II-2*(E*X(8)+B*X(13)+X(1)*X(14)+X(2)*X(15)+C*X(16)+

```

```

D*X(17))-X(8)**2/R
A(14)=-E*X(9)-B*X(14)-X(1)*X(18)-X(2)*X(19)-C*X(20)-
D*X(21)-X(8)*X(9)/R
A(15)=-E*X(10)-B*X(15)-X(1)*X(19)-X(2)*X(22)-C*X(23)-
D*X(24)-X(8)*X(10)/R
A(16)=-E*X(11)-B*X(16)-X(1)*X(20)-X(2)*X(23)-C*X(25)-
D*X(26)-X(8)*X(11)/R
A(17)=-E*X(12)-B*X(17)-X(1)*X(21)-X(2)*X(24)-C*X(26)-
D*X(27)-X(8)*X(12)/R
A(18)=-X(9)**2/R
A(19)=-X(9)*X(10)/R
A(20)=-X(9)*X(11)/R
A(21)=-X(9)*X(12)/R
A(22)=-X(10)**2/R
A(23)=-X(10)*X(11)/R
A(24)=-X(10)*X(12)/R
A(25)=-X(11)**2/R
A(26)=-X(11)*X(12)/R
A(27)=-X(12)**2/R
RETURN
END

```



```

SUBROUTINE OUT (F)
  DIMENSION X(162), Z(500), AC(500), XPE(4)
  COMMON N,T,X
  COMMON Q,Z,J,JJITO,KI,R,ITD,JJJ,AC
  WRITE (6,608) T,(X(I),I=1,N)
6080FORMAT (1H0,5X,'T=',F6.3,11X,'X1=',E11.4,5X,'X2=',
1E11.4,5X,'X3=',E11.4,5X,'X4=',E11.4,/6X,'X5=',E11.4,
25X,'X6=',E11.4,5X,'X7=',E11.4,4X,'X8=',E11.4,4X,
3'P13=',E11.4,/6X,'P14=',E11.4,4X,'P15=',E11.4,4X,
4'P16=',E11.4,4X,'P22=',E11.4,4X,'P23=',E11.4,/6X,'P24=',
5E11.4,4X,'P25=',E11.4,4X,'P26=',E11.4,4X,'P33=',E11.4,
64X,'P34=',E11.4,4X,'P45=',E11.4,4X,'P46=',E11.4,4X,
7'P44=',E11.4,4X,'P56=',E11.4,4X,'P66=',E11.4,/6X,
8'P55=',E11.4,4X,'P56=',E11.4,4X,'P66=',E11.4)
  XPE(1) = (X(3)-160.)/1.6
  XPE(2) = (X(4)-1.6)/0.016
  XPE(3) = (X(5)-1960.)/19.6
  XPE(4) = (X(6)-4.35)/0.0435
  WRITE(6,500) (XPE(I),I=1,4)
500 FORMAT(6X,'PEX3=',E11.4,3X,'PEX4=',E11.4,3X,'PEX5=',E11.4,3X,'PEX6
1=',E11.4)
  JJJ=JJJ+1
  AC(JJJ)=X(1)
  RETURN
  END

```

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13. ABSTRACT An analysis of numerical methods for extracting aerodynamic coefficients from dynamic test data has been conducted. The emphasis of the analysis is on the effects that random measurement errors in the data and random disturbances in the system have on the accuracy with which the coefficients for linear and nonlinear systems can be determined. Both deterministic and stochastic methods for extracting the coefficients and determining their uncertainties are considered. The deterministic technique considered, due to Chapman and Kirk, provides excellent estimates of both linear and nonlinear static pitching moment coefficients for the range of measurement efforts and system noise considered. Somewhat less accurate estimates of pitch damping coefficients are obtained. The stochastic approach considered demonstrates the feasibility of using an extended Kalman filter, with a parameter augmented state vector, for determining the values of the aerodynamic coefficients and their uncertainties from noisy dynamic test data. Parameter estimates obtained from the extended filter compare favorably with previously obtained results using deterministic techniques. Estimates of the parameter uncertainties provided by the filter are generally superior to those obtained with deterministic techniques particularly when system noise has corrupted the data.		

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